

PATTERNS IN RANDOM PERMUTATIONS AVOIDING SOME SETS OF MULTIPLE PATTERNS

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ABSTRACT. We consider a random permutation drawn from the set of permutations of length n that avoid some given set of patterns of length 3. We show that the number of occurrences of another pattern σ has a limit distribution, after suitable scaling. In several cases, the number is asymptotically normal; this contrasts to the cases of permutations avoiding a single pattern of length 3 studied in earlier papers.

1. Introduction

Let \mathfrak{S}_n be the set of permutations of $[n] := \{1, \dots, n\}$, and $\mathfrak{S}_* := \bigcup_{n \geq 1} \mathfrak{S}_n$. If $\sigma = \sigma_1 \cdots \sigma_m \in \mathfrak{S}_m$ and $\pi = \pi_1 \cdots \pi_n \in \mathfrak{S}_n$, then an *occurrence* of σ in π is a subsequence $\pi_{i_1} \cdots \pi_{i_m}$, with $1 \leq i_1 < \cdots < i_m \leq n$, that has the same order as σ , i.e., $\pi_{i_j} < \pi_{i_k} \iff \sigma_j < \sigma_k$ for all $j, k \in [m]$. We let $n_\sigma(\pi)$ be the number of occurrences of σ in π , and note that

$$\sum_{\sigma \in \mathfrak{S}_m} n_\sigma(\pi) = \binom{n}{m}, \quad (1.1)$$

for every $\pi \in \mathfrak{S}_n$. For example, an inversion is an occurrence of 21, and thus $n_{21}(\pi)$ is the number of inversions in π .

We say that π *avoids* another permutation τ if $n_\tau(\pi) = 0$; otherwise, π *contains* τ . Let

$$\mathfrak{S}_n(\tau) := \{\pi \in \mathfrak{S}_n : n_\tau(\pi) = 0\}, \quad (1.2)$$

the set of permutations of length n that avoid τ . More generally, for any set $T = \{\tau_1, \dots, \tau_k\}$ of permutations, let

$$\mathfrak{S}_n(T) = \mathfrak{S}_n(\tau_1, \dots, \tau_k) := \bigcap_{i=1}^k \mathfrak{S}_n(\tau_i), \quad (1.3)$$

the set of permutations of length n that avoid all $\tau_i \in T$. We also let $\mathfrak{S}_*(T) := \bigcup_{n=1}^{\infty} \mathfrak{S}_n(T)$ be the set of T -avoiding permutations of arbitrary length.

The classes $\mathfrak{S}_*(\tau)$ and, more generally, $\mathfrak{S}_*(T)$ have been studied for a long time, see e.g. Knuth [16, Exercise 2.2.1-5], Simion and Schmidt [20], Bóna [3]. In particular, one classical problem is to enumerate the sets $\mathfrak{S}_n(\tau)$,

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either exactly or asymptotically, see Bóna [3, Chapters 4–5]. We note the fact that for any τ with length $|\tau| = 3$, $\mathfrak{S}_n(\tau)$ has the same size $|\mathfrak{S}_n(\tau)| = C_n := \frac{1}{n+1} \binom{2n}{n}$, the n -th Catalan number, see e.g. [16, Exercises 2.2.1-4,5], [20], [21, Exercise 6.19ee,ff], [3, Corollary 4.7]; furthermore, the cases when T consists of several permutations of length 3 are all treated by Simion and Schmidt [20]. (The situation for $|\tau| \geq 4$ is more complicated.)

The general problem that concerns us is to take a fixed set T of one or several permutations and let $\pi_{T;n}$ be a uniformly random T -avoiding permutation, i.e., a uniformly random element of $\mathfrak{S}_n(T)$, and then study the distribution of the random variable $n_\sigma(\pi_{T;n})$ for some other fixed permutation σ . (Only σ that are themselves T -avoiding are interesting, since otherwise $n_\sigma(\pi_{T;n}) = 0$.) One instance of this problem was studied already by Robertson, Wilf and Zeilberger [19], who gave a generating function for $n_{123}(\pi_{132;n})$. The exact distribution of $n_\sigma(\pi_{\tau;n})$ for a given n was studied numerically in [15], where higher moments and mixed moments are calculated for small n . We are mainly interested in asymptotics of the distribution of $n_\sigma(\pi_{T;n})$, and of its moments, as $n \rightarrow \infty$, for some fixed T and σ .

In the present paper we study the cases when T is a set of two or more permutations of length 3. The cases when $T = \{\tau\}$ for a single permutation τ of length $|\tau| = 3$ were studied in [12; 13] (by symmetries, see Section 2.2, only two such cases have to be considered), and the cases when T contains a permutation of length ≤ 2 are trivial (there is then at most one permutation in $\mathfrak{S}_n(T)$ for any n); hence the present paper completes the study of forbidding one or several permutations of length ≤ 3 . The case of forbidding one or several permutations of length ≥ 4 seems much more complicated, but there are recent impressive results in some cases by Bassino, Bouvel, Féray, Gerin and Pierrot [2] and Bassino, Bouvel, Féray, Gerin, Maazoun and Pierrot [1].

The expectation $\mathbb{E} n_\sigma(\pi_{T;n})$, or equivalently, the total number of occurrences of σ in all T -avoiding permutations, has previously been treated in a number of papers for various cases, beginning with Bóna [5; 7] (with $\tau = 132$). In particular, Zhao [22] has given exact formulas when $|\sigma| = 3$ for the (non-trivial) cases treated in the present paper, where T consist of two or more permutations of length 3.

Remark 1.1. For the non-restricted case of uniformly random permutations in \mathfrak{S}_n , it is well-known that if π_n is a uniformly random permutation in \mathfrak{S}_n , then $n_\sigma(\pi_n)$ has an asymptotic normal distribution as $n \rightarrow \infty$ for every fixed permutation σ ; more precisely, if $|\sigma| = m$ then, as $n \rightarrow \infty$,

$$\frac{n_\sigma(\pi_n) - \frac{1}{m!} \binom{n}{m}}{n^{m-1/2}} \xrightarrow{d} N(0, \gamma^2) \quad (1.4)$$

for some $\gamma^2 > 0$ depending on σ ; see Bóna [4; 6] and Janson, Nakamura and Zeilberger [15, Theorem 4.1].

We obtain below similar asymptotic normal results in several cases (Sections 4, 5, 6, 8); note that the asymptotic normality in particular implies

concentration in these cases, in the sense

$$\frac{n_\sigma(\boldsymbol{\pi}_{T;n})}{\mathbb{E} n_\sigma(\boldsymbol{\pi}_{T;n})} \xrightarrow{p} 1. \tag{1.5}$$

On the other hand, in other cases (Sections 3, 7, 9, 10, 11) we find a different type of limits, where $n_\sigma(\boldsymbol{\pi}_{T;n})/\mathbb{E} n_\sigma(\boldsymbol{\pi}_{T;n})$ converges to some non-trivial positive random variable. The same holds in the case $T = \{2413, 3142\}$ studied by Bassino, Bouvel, Féray, Gerin and Pierrot [2].

We see no obvious pattern in the occurrence of these two types of limits in the cases below; nor do we know whether these are the only possibilities for a general set T of forbidden permutations.

Remark 1.2. In the present paper we consider for simplicity often only univariate limits; corresponding multivariate results for several $\sigma_1, \dots, \sigma_k$ follow by the same methods. In particular, (1.4) and all instances of normal limit laws below extend to multivariate normal limits, with covariance matrices that can be computed explicitly.

Remark 1.3. In the present paper we study only the numbers n_σ of occurrences of some pattern in $\boldsymbol{\pi}_{\tau;n}$. There is also a number of papers by various authors that study other properties of random τ -avoiding permutations, see e.g. the references in [13]; such results will not be considered here.

2. Preliminaries

2.1. Notation. Let $\iota = \iota_n$ be the identity permutation of length n . Let $\bar{\iota}_n = n \cdots 21$ be its reversal.

Let $\pi = \pi_1 \cdots \pi_n$ be a permutation. We say that a value π_i is a *maximum* if $\pi_i > \pi_j$ for every $j < i$, and a *minimum* if $\pi_i < \pi_j$ for every $j < i$. (These are sometimes called *LR maximum* and *LR minimum*.) Note that π_1 always is both a maximum and a minimum.

2.2. Symmetries. There are many cases treated in the present paper, but the number is considerably reduced by three natural symmetries (used by many previous authors). For any permutation $\pi = \pi_1 \cdots \pi_n$, define its *inverse* π^{-1} in the usual way, and its *reversal* and *complement* by

$$\pi^r := \pi \circ \bar{\iota} = \pi_n \cdots \pi_1, \tag{2.1}$$

$$\pi^c := \bar{\iota} \circ \pi = (n + 1 - \pi_1) \cdots (n + 1 - \pi_n). \tag{2.2}$$

These three operations are all involutions, and they generate a group \mathfrak{G} of 8 symmetries (isomorphic to the dihedral group D_4). It is easy to see that, for any permutations σ and π ,

$$n_{\sigma^{-1}}(\pi^{-1}) = n_{\sigma^r}(\pi^r) = n_{\sigma^c}(\pi^c) = n_\sigma(\pi), \tag{2.3}$$

and consequently, for any symmetry $s \in \mathfrak{G}$,

$$n_{\sigma^s}(\pi^s) = n_\sigma(\pi). \tag{2.4}$$

For a set T of permutations we define $T^s := \{\tau^s : \tau \in T\}$. It follows from (2.4) that

$$\mathfrak{S}_n(T^s) = \{\pi^s : \pi \in \mathfrak{S}_n(T)\}, \quad (2.5)$$

and, furthermore, that for any permutation σ ,

$$n_{\sigma^s}(\pi_{T^s;n}) \stackrel{d}{=} n_\sigma(\pi_{T;n}). \quad (2.6)$$

We say that the sets of forbidden permutations T and T^s are *equivalent*, and note that (2.6) implies that it suffices to consider one set T in each equivalence class $\{T^s : s \in \mathfrak{S}\}$. We do this in the sequel without further comment. (We choose representatives T that we find convenient. One guide is that we choose T such that the identity permutation ι_n avoids T .)

2.3. Compositions and decompositions of permutations. If $\sigma \in \mathfrak{S}_m$ and $\tau \in \mathfrak{S}_n$, their *composition* $\sigma * \tau \in \mathfrak{S}_{m+n}$ is defined by letting τ act on $[m+1, m+n]$ in the natural way; more formally, $\sigma * \tau = \pi \in \mathfrak{S}_{m+n}$ where $\pi_i = \sigma_i$ for $1 \leq i \leq m$, and $\pi_{j+m} = \tau_j + m$ for $1 \leq j \leq n$. It is easily seen that $*$ is an associative operation that makes \mathfrak{S}_* into a semigroup (without unit, since we only consider permutations of length ≥ 1). We say that a permutation $\pi \in \mathfrak{S}_*$ is *decomposable* if $\pi = \sigma * \tau$ for some $\sigma, \tau \in \mathfrak{S}_*$, and *indecomposable* otherwise; we also call an indecomposable permutation a *block*. Equivalently, $\pi \in \mathfrak{S}_n$ is decomposable if and only if $\pi : [m] \rightarrow [m]$ for some $1 \leq m < n$. See e.g. [8, Exercise VI.14].

It is easy to see that any permutation $\pi \in \mathfrak{S}_*$ has a unique decomposition $\pi = \pi_1 * \cdots * \pi_\ell$ into indecomposable permutations (blocks) π_1, \dots, π_ℓ (for some, unique, $\ell \geq 1$); we call these the *blocks of π* .

We shall see that some (but not all) of the classes considered below can be characterized in terms of their blocks. (See [2] for another, more complicated, example.)

2.4. U -statistics. An (asymmetric) *U -statistic* is a random variable of the form

$$U_n = \sum_{1 \leq i_1 < \cdots < i_d \leq n} f(X_{i_1}, \dots, X_{i_d}), \quad n \geq 0, \quad (2.7)$$

where X_1, X_2, \dots is an i.i.d. sequence of random variables and f is a given function of $d \geq 1$ variables. These were (in the symmetric case) introduced by Hoeffding [9]; see further e.g. [14] and the references there. We say that d is the *order* of the U -statistic.

We shall use the central limit theorem for U -statistics, originally due to Hoeffding [9], in the asymmetric version given in [10, Theorem 11.20] and [14, Corollary 3.5 and (moment convergence) Theorem 3.15]. Let, with X denoting a generic X_i ,

$$\mu := \mathbb{E} f(X_1, \dots, X_d), \quad (2.8)$$

$$f_i(x) := \mathbb{E}(f(X_1, \dots, X_d) \mid X_i = x), \quad (2.9)$$

$$\beta_{ij} := \text{Cov}(f_i(X), f_j(X)), \quad (2.10)$$

$$\beta^2 := \sum_{i,j=1}^d \frac{(i+j-2)!(2d-i-j)!}{(i-1)!(j-1)!(d-i)!(d-j)!(2d-1)!} \beta_{ij}. \quad (2.11)$$

Note that $f_i(x)$ in [14] is $f_i(x) - \mu$ in the present notation.

Proposition 2.1 ([10; 14]). *Suppose that $f(X_1, \dots, X_d) \in L^2$. Then, with the notation in (2.8)–(2.11), as $n \rightarrow \infty$,*

$$\frac{U_n - \binom{n}{d} \mu}{n^{d-1/2}} \xrightarrow{d} N(0, \beta^2). \quad (2.12)$$

Furthermore, $\beta^2 > 0$ unless $f_i(X) = \mu$ a.s. for $i = 1, \dots, d$.

Moreover, if $f(X_1, \dots, X_d) \in L^p$ for some $p \geq 2$, the (2.12) holds with convergence of all moments of order $\leq p$. \square

Example 2.2. A uniformly random permutation π_n of length n (without other restrictions) can be constructed as the relative order of X_1, \dots, X_n , where X_i are i.i.d. with, for example, a uniform distribution $U(0, 1)$. For any given permutation $\sigma \in \mathfrak{S}_m$, we can then write $n_\sigma(\pi_n)$ as a U -statistic (2.7) for a suitable indicator function f . Then Proposition 2.1 yields a limit theorem showing that $n_\sigma(\pi_n)$ is asymptotically normal. See [15] for details.

We shall also use a renewal theory version of Proposition 2.1. With the notations above, assume (for simplicity) that $X_i \geq 0$. Define $S_n := \sum_{i=1}^n X_i$, and let for each $x > 0$

$$N_-(x) := \sup\{n : S_n < x\}, \quad (2.13)$$

$$N_+(x) := \inf\{n : S_n \geq x\} = N_-(x) + 1. \quad (2.14)$$

Remark 2.3. The definitions (2.13)–(2.14) differ slightly from the ones in [14], where instead $S_n \leq x$ and $S_n > x$ are used. This does not affect the asymptotic results used here. Note that the event $\{S_k = n \text{ for some } k \geq 0\}$ equals $\{S_{N_+(n)} = n\}$ in the present notation.

The following results are special cases of [14, Theorems 3.11, 3.13(iii) and 3.18] (with somewhat different notation). $N_\pm(x)$ means either $N_-(x)$ or $N_+(x)$; the results holds for both.

Proposition 2.4 ([14]). *Suppose that $f(X_1, \dots, X_d) \in L^2$, $X \in L^2$, $X \geq 0$ a.s., and $\nu := \mathbb{E} X > 0$. Then, with notations as above, as $x \rightarrow \infty$,*

$$\frac{U_{N_\pm(x)} - \nu^{-d} \mu d!^{-1} x^d}{x^{d-1/2}} \xrightarrow{d} N(0, \gamma^2), \quad (2.15)$$

where

$$\gamma^2 := \nu^{1-2d} \beta^2 - 2 \frac{\nu^{-2d} \mu}{(d-1)! d!} \sum_{i=1}^d \text{Cov}(f_i(X), X) + \frac{\nu^{-2d-1} \mu^2}{(d-1)!^2} \text{Var}(X). \quad (2.16)$$

Moreover, $\gamma^2 > 0$ unless $f_i(X) = \frac{\mu}{\nu} X$ a.s. for $i = 1, \dots, d$. \square

Proposition 2.5 ([14]). *Suppose in addition to the hypotheses in Proposition 2.4 that X is integer-valued. Then (2.15) holds also conditioned on $S_{N_+(x)} = x$ (cf. Remark 2.3) for integers $x \rightarrow \infty$. \square*

Proposition 2.6 ([14]). *Suppose in addition to the hypotheses in Proposition 2.4 or 2.5 that $f(X_1, \dots, X_d) \in L^p$ and $X \in L^p$ for every $p < \infty$. Then the conclusion (2.15) holds with convergence of all moments. \square*

2.5. Trivial cases. We consider in the present paper sets $T \subseteq \mathfrak{S}_3$. Then, see [20], the following cases are trivial in the sense that for all $n \geq 5$, $|\mathfrak{S}_n(T)| = 0, 1$ or 2 .

- (i) $T = \{123, 321\}$,
- (ii) $|T| = 3$ and $T \supset \{123, 321\}$,
- (iii) $|T| \geq 4$.

We ignore these cases in the sequel. This leaves 6 cases with $|T| = 1$ (Section 3), 14 cases with $|T| = 2$ (Sections 4–7), and 16 cases with $|T| = 3$ (Sections 8–11). Symmetries reduce these to the $2 + 4 + 4 = 10$ non-equivalent cases discussed below.

3. Avoiding a single permutation of length 3

There are 6 cases where a single permutation of length 3 is avoided, but by the symmetries in Subsection 2.2 these reduce to 2 non-equivalent cases, for example 132 (equivalent to 231, 213, 312) and 321 (equivalent to 123). These cases are treated in detail in [12] and [13], respectively. Both analyses are based on bijections with binary trees and Dyck paths, and the well-known convergence in distribution of random Dyck paths to a Brownian excursion, but the details are very different, and so are in general the resulting limit distributions.

For comparison with the results in later sections, we quote the main results of [12] and [13], referring to these papers for further details and proofs. Recall that the standard Brownian excursion $\mathbf{e}(x)$ is a random non-negative function on $[0, 1]$.

First, for 132, let

$$\lambda(\sigma) := |\sigma| + D(\sigma) \tag{3.1}$$

where $D(\sigma)$ is the number of *descents* in σ , i.e., indices i such that $\sigma_i > \sigma_{i+1}$ or (as a convenient convention) $i = |\sigma|$. Note that $1 \leq D(\sigma) \leq |\sigma|$, and thus

$$|\sigma| + 1 \leq \lambda(\sigma) \leq 2|\sigma|, \tag{3.2}$$

with the extreme values $\lambda(\sigma) = |\sigma| + 1$ if and only if $\sigma = 1 \cdots k$, and $\lambda(\sigma) = 2|\sigma|$ if and only if $\sigma = k \cdots 1$, for some $k = |\sigma|$.

Theorem 3.1 ([12]). *There exist strictly positive random variables Λ_σ such that*

$$n_\sigma(\boldsymbol{\pi}_{132;n})/n^{\lambda(\sigma)/2} \xrightarrow{d} \Lambda_\sigma, \tag{3.3}$$

as $n \rightarrow \infty$, jointly for all $\sigma \in \mathfrak{S}_*(132)$. Moreover, this holds with convergence of all moments.

For a monotone decreasing permutation $k \cdots 1$, $\Lambda_{k \cdots 1} = 1/k!$ is deterministic, but not for any other σ . \square

The limit variables Λ_σ in Theorem 3.1 can be expressed as functionals of a Brownian excursion $\mathbf{e}(x)$, see [12]; the description is, in general, rather complicated, but some cases are simple.

Example 3.2. In the special case $\sigma = 12$, $\Lambda_{12} = \sqrt{2} \int_0^1 \mathbf{e}(x) dx$, see [12, Example 7.6]; this is (apart from the factor $\sqrt{2}$) the well-known *Brownian excursion area*, see e.g. [11] and the references there.

For the number n_{21} of inversions, we thus have

$$\frac{\binom{n}{2} - n_{21}(\boldsymbol{\pi}_{132;n})}{n^{3/2}} = \frac{n_{12}(\boldsymbol{\pi}_{132;n})}{n^{3/2}} \xrightarrow{d} \Lambda_{12} = \sqrt{2} \int_0^1 \mathbf{e}(x) dx. \quad (3.4)$$

By Subsection 2.2, the left-hand side can also be seen as the number of inversions $n_{21}(\boldsymbol{\pi}_{231;n})$ or $n_{21}(\boldsymbol{\pi}_{312;n})$, normalized by $n^{3/2}$, where we instead avoid 231 or 312.

Theorem 3.3 ([13]). *Let $\sigma \in \mathfrak{S}_*(321)$. Let $m := |\sigma|$, and suppose that σ has ℓ blocks of lengths m_1, \dots, m_ℓ . Then, as $n \rightarrow \infty$,*

$$n_\sigma(\boldsymbol{\pi}_{321;n})/n^{(m+\ell)/2} \xrightarrow{d} W_\sigma \quad (3.5)$$

for a positive random variable W_σ that can be represented as

$$W_\sigma = w_\sigma \int_{0 < t_1 < \dots < t_\ell < 1} \mathbf{e}(t_1)^{m_1-1} \dots \mathbf{e}(t_\ell)^{m_\ell-1} dt_1 \dots dt_\ell, \quad (3.6)$$

where w_σ is positive constant.

Moreover, the convergence (3.5) holds jointly for any set of $\sigma \in \mathfrak{S}_*(321)$, and with convergence of all moments.

Example 3.4. Let $\sigma = 21$. Then $w_{21} = 2^{-1/2}$, see [13], and thus (3.5)–(3.6), with $\ell = 1$ and $m_1 = m = 2$, yield for the number of inversions,

$$\frac{n_{21}(\boldsymbol{\pi}_{321;n})}{n^{3/2}} \xrightarrow{d} 2^{-1/2} \int_0^1 \mathbf{e}(x) dx. \quad (3.7)$$

Note that the limit in (3.7) differs from the one in (3.4) by a factor 2.

4. Avoiding $\{132, 312\}$

In this section we avoid $T = \{132, 312\}$. Equivalent sets are $\{132, 231\}$, $\{213, 231\}$, $\{213, 312\}$.

It was shown by Simion and Schmidt [20] that $|\mathfrak{S}_n(132, 312)| = 2^{n-1}$, together with the following characterization (in an equivalent formulation).

Proposition 4.1 ([20, Proposition 12]). *A permutation π belongs to the class $\mathfrak{S}_*(132, 312)$ if and only if every entry π_i is either a maximum or a minimum.* \square

We encode a permutation $\pi \in \mathfrak{S}_n(132, 312)$ by a sequence $\xi_2, \dots, \xi_n \in \{\pm 1\}^{n-1}$, where $\xi_j = 1$ if π_j is a maximum in π , and $\xi_j = -1$ if π_j is a minimum. This is by Proposition 4.1 a bijection, and hence the code for a uniformly random $\pi_{132,312;n}$ has ξ_2, \dots, ξ_n i.i.d. with the symmetric Bernoulli distribution $\mathbb{P}(\xi_j = 1) = \mathbb{P}(\xi_j = -1) = \frac{1}{2}$. We let ξ_1 have the same distribution and be independent of ξ_2, \dots, ξ_n .

Let $\sigma \in \mathfrak{S}_m(132, 312)$ have the code η_2, \dots, η_m . Then $\pi_{i_1} \cdots \pi_{i_m}$ is an occurrence of σ in π if and only if $\xi_{i_j} = \eta_j$ for $2 \leq j \leq m$. Consequently, cf. Subsection 2.4, $n_\sigma(\pi_{132,312;n})$ is a U -statistic

$$n_\sigma(\pi_{132,312;n}) = \sum_{i_1 < \cdots < i_m} f(\xi_{i_1}, \dots, \xi_{i_m}), \quad (4.1)$$

where

$$f(\xi_1, \dots, \xi_m) := \prod_{j=2}^m \mathbf{1}\{\xi_j = \eta_j\}. \quad (4.2)$$

Note that f does not depend on the first argument. It follows that, with the notation (2.8)–(2.11),

$$\mu = \mathbb{E} f(\xi_1, \dots, \xi_m) = 2^{-(m-1)}, \quad (4.3)$$

$$f_i(\xi) = \begin{cases} 0, & i = 1, \\ 2^{-(m-2)} \mathbf{1}\{\xi = \eta_i\}, & 2 \leq i \leq m, \end{cases} \quad (4.4)$$

$$\beta_{ij} = \text{Cov}(f_i(\xi), f_j(\xi)) = 2^{2-2m} \eta_i \eta_j, \quad i, j \geq 2, \quad (4.5)$$

$$\beta^2 = 2^{2-2m} \sum_{i,j=2}^m \frac{(i+j-2)!(2m-i-j)!}{(i-1)!(j-1)!(m-i)!(m-j)!(2m-1)!} \eta_i \eta_j. \quad (4.6)$$

Proposition 2.1 yields:

Theorem 4.2. *For any $m \geq 1$ and $\sigma \in \mathfrak{S}_m(132, 312)$, as $n \rightarrow \infty$,*

$$\frac{n_\sigma(\pi_{132,312;n}) - 2^{1-m} n^m / m!}{n^{m-1/2}} \xrightarrow{d} N(0, \beta^2), \quad (4.7)$$

with $\beta^2 > 0$ given by (4.6).

Moreover, (4.7) holds with convergence of all moments.

Example 4.3. For the number of inversions, we have $\sigma = 21$ and $m = 2$, $\eta_2 = -1$. Thus, (4.4) yields $f_1(\xi) = 0$ and $f_2(\xi) = \mathbf{1}\{\xi = -1\}$. We find, from (4.3)–(4.6), $\mu = \frac{1}{2}$, $\beta_{22} = \frac{1}{4}$ and $\beta^2 = \frac{1}{12}$, and thus Theorem 4.2 yields

$$\frac{n_{21}(\pi_{132,312;n}) - n^2/4}{n^{3/2}} \xrightarrow{d} N(0, \frac{1}{12}), \quad (4.8)$$

Remark 4.4. It is easily seen from (4.1)–(4.2) that the expected number of occurrences $\mathbb{E} n_\sigma(\pi_{132,312;n}) = 2^{1-m} \binom{n}{m}$, for every $\sigma \in \mathfrak{S}_m(132, 312)$; hence the expectation depends only on the length $m = |\sigma|$.

The variance depends not only on $|\sigma|$, not even asymptotically, by (4.6).

5. Avoiding $\{231, 312\}$

In this section we consider $T = \{231, 312\}$. The only equivalent set is $\{132, 213\}$.

It was shown by Simion and Schmidt [20] that $|\mathfrak{S}_n(231, 312)| = 2^{n-1}$, together with the following characterization (in an equivalent form).

Proposition 5.1 ([20, Proposition 12]). *A permutation π belongs to the class $\mathfrak{S}_*(231, 312)$ if and only if every block in π is decreasing, i.e., of the type $\ell(\ell - 1) \cdots 21$ for some ℓ . \square*

Hence there exists exactly one block of each length $\ell \geq 1$, and a permutation $\pi \in \mathfrak{S}_*(231, 312)$ is uniquely determined by the block lengths. In this section, let $\pi_{\ell_1, \dots, \ell_b}$ denote the permutation in $\mathfrak{S}_*(231, 312)$ with block lengths ℓ_1, \dots, ℓ_b , i.e.,

$$\pi_{\ell_1, \dots, \ell_b} := \bar{\ell}_1 * \cdots * \bar{\ell}_b. \quad (5.1)$$

If $\sigma, \pi \in \mathfrak{S}_*(231, 312)$, then in an occurrence of σ in π , each block in σ has to be mapped into a block in π , and distinct blocks have to be mapped into distinct blocks. Conversely, any such increasing map $[m] \rightarrow [n]$ defines an occurrence of σ . It follows that if $\sigma = \pi_{\ell_1, \dots, \ell_b}$, then

$$n_\sigma(\pi_{L_1, \dots, L_B}) = \sum_{1 \leq i_1 < \dots < i_b \leq B} \prod_{j=1}^b \binom{L_{i_j}}{\ell_j}. \quad (5.2)$$

This is similar to a U -statistic (2.7), but note that if we write $\pi_{231, 312; n}$ as π_{L_1, \dots, L_B} , then the block lengths L_1, \dots, L_B are not independent (since their sum is fixed $= n$), and the number of blocks B is random. However, we can analyze this variable using the renewal theory in Subsection 2.4 as follows.

First, mark each endpoint of the blocks in $\pi \in \mathfrak{S}_n(231, 312)$ by 1, and mark all other indices in $[n]$ by 0. Thus π defines a string $\xi_1, \dots, \xi_n \in \{0, 1\}^n$, where necessarily $\xi_n = 1$ but ξ_1, \dots, ξ_{n-1} are arbitrary. This yields a bijection between $\mathfrak{S}_n(231, 312)$ and the 2^{n-1} such strings; hence, we obtain a uniformly random $\pi_{231, 312; n}$ by letting ξ_1, \dots, ξ_{n-1} be i.i.d. $\text{Be}(\frac{1}{2})$, i.e., with $\mathbb{P}(\xi_i = 0) = \mathbb{P}(\xi_i = 1) = \frac{1}{2}$.

We change notation a little, to avoid problems at the endpoint, and define ξ'_1, ξ'_2, \dots as an infinite i.i.d. sequence with $\xi'_i \sim \text{Be}(\frac{1}{2})$. Regard each i with $\xi'_i = 1$ as the end of a block, and let X_1, X_2, \dots , be the successive lengths of these (infinitely many) blocks. Then X_i are i.i.d. with

$$X_i \sim \text{Ge}(\frac{1}{2}). \quad (5.3)$$

Given n , we then may let $\xi_i := \xi'_i$ for $1 \leq i < n$, and $\xi_n := 1$; this determines ξ_1, \dots, ξ_n as above, and thus a uniformly random $\pi_{231, 312; n}$. With this construction, the number of blocks in $\pi_{231, 312; n}$ is, recalling (2.13)–(2.14), $B = N_+(n)$, and the block lengths are

$$L_i = \begin{cases} X_i, & i < N_+(n) \\ n - \sum_{i < N_+(n)} X_i \leq X_{N_+(n)}, & i = N_+(n). \end{cases} \quad (5.4)$$

Consequently, if $\sigma = \pi_{\ell_1, \dots, \ell_b}$ and we define

$$f(x_1, \dots, x_b) := \prod_{j=1}^b \binom{x_j}{\ell_j}, \quad (5.5)$$

then (5.2) and (2.7) show that

$$U_{N_-(n)} \leq n_\sigma(\boldsymbol{\pi}_{231,312;n}) \leq U_{N_+(n)}. \quad (5.6)$$

Consequently, the asymptotic result in (2.15), which holds for both $U_{N_-(n)}$ and $U_{N_+(n)}$, holds also for $n_\sigma(\boldsymbol{\pi}_{231,312;n})$.

Remark 5.2. Alternatively, we can obtain (ξ_i) from (ξ'_i) by conditioning on $\xi'_n = 1$, and note that this holds when $S_{N_+(n)} = n$ (see Remark 2.3), and then $n_\sigma(\boldsymbol{\pi}_{231,312;n}) = U_{N_+(n)}$. The result then follows from Proposition 2.5.

To calculate the parameters, note that, by (5.3), X has the probability generating function

$$g(z) := \mathbb{E} z^X = \sum_{k=0}^{\infty} 2^{-k} z^k = \frac{z}{2-z} = \frac{2}{2-z} - 1 \quad (5.7)$$

and it follows that for any integers $k, l \geq 0$ with $(k, l) \neq (0, 0)$,

$$\mathbb{E} \left(\binom{X}{k} \binom{X}{l} \right) = [z^k w^l] \mathbb{E}((1+z)^X (1+w)^X) \quad (5.8)$$

$$= [z^k w^l] g((1+z)(1+w)) \quad (5.9)$$

$$= [z^k w^l] \frac{2}{2 - (1+z)(1+w)} \quad (5.10)$$

$$= [z^k w^l] \frac{2}{1 - z - w - zw} \quad (5.11)$$

$$= 2D(k, l) = 2 \sum_{i=0}^{k \wedge l} \frac{(k+l-i)!}{(k-i)! (\ell-i)! i!} \quad (5.12)$$

where $D(k, l)$ denotes the Delannoy numbers. ($D(k, l)$ is, e.g., the number of lattice paths from $(0, 0)$ to (k, l) with steps $(1, 0)$, $(0, 1)$ or $(1, 1)$; see [21, Example 6.3.8] and [18, A008288 and A001850] and the references there.) Simple calculations then yield

$$\nu = \mathbb{E} X = 2, \quad (5.13)$$

$$\text{Var}(X) = 2. \quad (5.14)$$

$$\mu = \mathbb{E} \prod_{j=1}^b \binom{X}{\ell_j} = \prod_{j=1}^b \mathbb{E} \binom{X}{\ell_j} = 2^b, \quad (5.15)$$

$$f_i(X) = 2^{b-1} \binom{X}{\ell_i}, \quad (5.16)$$

$$\beta_{ij} = \text{Cov}(f_i(X), f_j(X)) = 2^{2b-1} D(\ell_i, \ell_j) - 2^{2b}, \quad (5.17)$$

$$\text{Cov}(f_i(X), X) = 2^b D(\ell_i, 1) - 2^{b+1} = (2\ell_i - 1)2^b. \quad (5.18)$$

Consequently, we obtain by Propositions 2.4 and 2.6 asymptotic normality in the following form.

Theorem 5.3. *Let $\sigma \in \mathfrak{S}_m(231, 312)$ have block lengths ℓ_1, \dots, ℓ_b . Then, as $n \rightarrow \infty$,*

$$\frac{n_\sigma(\pi_{231,312;n}) - n^b/b!}{n^{b-1/2}} \xrightarrow{d} N(0, \gamma^2), \quad (5.19)$$

where γ^2 can be calculated by (2.16) and (5.13)–(5.18).

Moreover, (5.19) holds with convergence of all moments. \square

Example 5.4. For the number of inversions, we have $\sigma = 21$ and $b = 1$, $\ell_1 = 2$. A calculation yields $\gamma^2 = 6$, and Theorem 5.3 yields

$$\frac{n_{21}(\pi_{231,312;n}) - n}{n^{1/2}} \xrightarrow{d} N(0, 6). \quad (5.20)$$

Remark 5.5. Theorem 5.3 shows that the typical order of $n_\sigma(\pi_{231,312;n})$ depends only on the number of blocks b in σ (but not on the length $|\sigma|$); more precisely, the asymptotic mean depends only on b . (Cf. the different situation when avoiding $\{132, 312\}$ in Section 4, see Remark 4.4.) Calculations (assisted by Maple) show, however, that the asymptotic variance γ^2 depends not only on m and b ; for example $\sigma = 2143 = \bar{i}_2 * \bar{i}_2$ has $\gamma^2 = 6$ while $\sigma = 3214 = \bar{i}_3 * \bar{i}_1$ has $\gamma^2 = 52/3$.

Remark 5.6. The asymptotic variance $\gamma^2 = 0$ when $\sigma = \iota_m = 1 \cdots m$, in which case $b = m$ and all blocks have length 1. This can be seen directly, since all other patterns occur only $O_p(n^{m-1})$ times (by Theorem 5.3), and thus ι_m occurs $\binom{n}{m} - O_p(n^{m-1})$ times. This argument also shows that the asymptotic variance of $n_{1 \cdots m}(\pi_{231,312;n})$ is of the order n^{2m-3} .

It follows from Proposition 2.4 that $\gamma^2 > 0$ for any other $\sigma \in \mathfrak{S}_*(231, 312)$.

6. Avoiding $\{231, 321\}$

In this section we consider $T = \{231, 321\}$. Equivalent sets are $\{123, 132\}$, $\{123, 213\}$, $\{312, 321\}$.

It was shown by Simion and Schmidt [20] that $|\mathfrak{S}_n(231, 321)| = 2^{n-1}$, together with the following characterization (in an equivalent form).

Proposition 6.1 ([20, Proposition 12]). *A permutation π belongs to the class $\mathfrak{S}_*(231, 321)$ if and only if every block in π is of the type $\ell 12 \cdots (\ell - 1)$ for some ℓ . \square*

Thus, as in Section 5, a permutation in $\mathfrak{S}_*(231, 321)$ is determined by its block lengths, and these can be arbitrary. In this section, let $\pi_{\ell_1, \dots, \ell_b}$ denote the permutation in $\mathfrak{S}_*(231, 321)$ with block lengths ℓ_1, \dots, ℓ_b .

Again, in an occurrence of σ in π , each block in σ has to be mapped into a block in π . However, this time, several consecutive blocks in σ may be mapped to the same block in π , provided they have length 1. Moreover, if a

block of length $\ell \geq 2$ in σ is mapped to a block in π , then the first element has to be mapped to the first element. Hence, we obtain instead of (5.2), if $\sigma = \pi_{\ell_1, \dots, \ell_b}$,

$$n_\sigma(\pi_{L_1, \dots, L_B}) = \sum_{1 \leq i_1 < \dots < i_b \leq B} \prod_{j=1}^b h_{\ell_i}(L_{i_j}) + R, \quad (6.1)$$

where

$$h_\ell(x) := \begin{cases} x, & \ell = 1, \\ \binom{x-1}{\ell-1}, & \ell \geq 2, \end{cases} \quad (6.2)$$

and R counts the occurrences where less than b different blocks in π_{L_1, \dots, L_B} are used. We represent the block lengths as in Section 5, in particular (5.3)–(5.4), again using an infinite i.i.d. sequence $X_i \sim \text{Ge}(\frac{1}{2})$. Then, the main term in (6.1) is sandwiched between U -statistics as in (5.6), and we can apply Proposition 2.4 to it. (Alternatively, we can use Proposition 2.5 as in Remark 5.2.)

By (5.7), $\mathbb{E} z^{X-1} = (2-z)^{-1}$, and calculations similar to (5.8) yield

$$\mathbb{E} \left(\binom{X-1}{k} \binom{X-1}{\ell} \right) = D(k, \ell), \quad k, \ell \geq 0. \quad (6.3)$$

Hence

$$\mathbb{E} h_\ell(X) = \begin{cases} D(\ell-1, 0) = 1, & \ell \geq 2, \\ 2, & \ell = 1. \end{cases} \quad (6.4)$$

Simple calculations then yield, in addition to (5.13)–(5.14), letting b_1 be the number of blocks of length 1,

$$\mu = \prod_{j=1}^b \mathbb{E} h_{\ell_j}(X) = 2^{b_1}, \quad (6.5)$$

$$f_i(X) = \begin{cases} 2^{b_1} \binom{X-1}{\ell_i-1}, & \ell_i \geq 2, \\ 2^{b_1-1} X, & \ell_i = 1, \end{cases} \quad (6.6)$$

$$\beta_{ij} = \text{Cov}(f_i(X), f_j(X)) = \begin{cases} 2^{2b_1} D(\ell_i-1, \ell_j-1) - 2^{2b_1}, & \ell_i, \ell_j \geq 2, \\ 2^{2b_1} (\ell_i-1), & \ell_i \geq 2 > \ell_j = 1, \\ 2^{2b_1-1}, & \ell_i = \ell_j = 1 \end{cases} \quad (6.7)$$

$$\text{Cov}(f_i(X), X) = \begin{cases} 2^{b_1+1} (\ell_i-1), & \ell_i \geq 2, \\ 2^{b_1}, & \ell_i = 1. \end{cases} \quad (6.8)$$

Consequently, we obtain by Propositions 2.4 and 2.6 asymptotic normality in the following form.

Theorem 6.2. *Let $\sigma \in \mathfrak{S}_m(231, 321)$ have block lengths ℓ_1, \dots, ℓ_b . Then, as $n \rightarrow \infty$,*

$$\frac{n_\sigma(\boldsymbol{\pi}_{231,321;n}) - 2^{b_1-b} n^b / b!}{n^{b-1/2}} \xrightarrow{d} N(0, \gamma^2), \quad (6.9)$$

where γ^2 can be calculated by (2.16) and (6.5)–(6.8).

Moreover, (6.9) holds with convergence of all moments.

Proof. The argument above yields the stated limit for the first (main) term on the right-hand side of (6.1). We show that the remainder term R is negligible.

The term R can be split up as a sum $\sum_{d=1}^{b-1} R_d$, where R_d counts the occurrences that use d blocks in $\pi = \pi_{L_1, \dots, L_B}$. Each R_d may be written as a sum over d -tuples of blocks, and thus bounded as in (5.6) by some U -statistics $U_{N_+(n)}^{(d)}$ of order d . Applying Proposition 2.4 (or Proposition 2.1, together with $N_+(n) \leq n$) to the latter, we find $R_d = O_p(n^d) = O_p(n^{b-1})$, and thus $R_d/n^{b-1/2} \xrightarrow{p} 0$. For moments, we similarly have by Proposition 2.6 or 2.1 $\mathbb{E}|R_d|^p = O(n^{pd}) = O(n^{p(b-1)}) = o(n^{p(b-1/2)})$. Hence, each R_d is negligible in the limit (6.9), and the result follows. \square

Example 6.3. For the number of inversions, we have $\sigma = 21$ and $b = 1$, $\ell_1 = 2$, $b_1 = 0$. A calculation yields $\gamma^2 = 1/4$, and Theorem 6.2 yields

$$\frac{n_{21}(\boldsymbol{\pi}_{231,321;n}) - n/2}{n^{1/2}} \xrightarrow{d} N(0, \frac{1}{4}). \quad (6.10)$$

In fact, we have the exact distribution

$$n_{21}(\boldsymbol{\pi}_{231,321;n}) \sim \text{Bi}(n-1, \frac{1}{2}). \quad (6.11)$$

To see this, note that, by Proposition 6.1, if we define ξ_2, \dots, ξ_n by

$$\xi_i := \mathbf{1}\{\text{no block begins at position } i\}, \quad (6.12)$$

then every sequence $\xi_2, \dots, \xi_n \in \{0, 1\}^{n-1}$ occurs for exactly one permutation in $\mathfrak{S}_n(231, 321)$, and thus ξ_2, \dots, ξ_n are i.i.d. $\text{Be}(\frac{1}{2})$. (This is a minor variation of the similar argument in Section 5.) Furthermore, for each $j \geq 2$, the number of inversions ij with $i < j$ equals ξ_j , so the total number is $\sum_2^n \xi_i \sim \text{Bi}(n-1, \frac{1}{2})$.

Remark 6.4. Unlike in Section 5, here the asymptotic mean depends not only on the number of blocks in σ , but also on their lengths.

Remark 6.5. As in Section 5, the asymptotic variance $\gamma^2 = 0$ when $\sigma = \iota_m = 1 \cdots m$, in which case $b = m$ and all blocks have length 1, but $\gamma^2 > 0$ for any other $\sigma \in \mathfrak{S}_*(231, 321)$.

7. Avoiding $\{132, 321\}$

In this section we consider $T = \{132, 321\}$. Equivalent sets are $\{123, 231\}$, $\{123, 312\}$, $\{213, 321\}$.

It was shown by Simion and Schmidt [20] that $|\mathfrak{S}_n(132, 321)| = \binom{n}{2} + 1$. (The case $\mathfrak{S}_n(132, 321)$ is thus more degenerate than the cases considered above, in the sense that the allowed set of permutations is much smaller; $|\mathfrak{S}_n(132, 321)|$ grows polynomially as (roughly) n^2 , compared to 2^{n-1} in the previous cases forbidding two permutations of length 3.) [20] gave also the following characterization. Given $k, \ell \geq 1$ and $m \geq 0$, let, in this section,

$$\pi_{k,\ell,m} := (\ell + 1, \dots, \ell + k, 1, \dots, \ell, k + \ell + 1, \dots, k + \ell + m) \in \mathfrak{S}_{k+\ell+m}. \quad (7.1)$$

Thus $\pi_{k,\ell,m}$ consists of three increasing runs of lengths k, ℓ, m (where the third run is empty when $m = 0$).

Proposition 7.1 ([20, Proposition 13]).

$$\mathfrak{S}_n(132, 321) = \{\pi_{k,\ell,n-k-\ell} : k, \ell \geq 1, k + \ell \leq n\} \cup \{\iota_n\}. \quad (7.2)$$

□

For asymptotic results, we may ignore the case when $\pi_{132,321;n} = \iota_n$, which has probability $1/(\binom{n}{2} + 1) = o(1)$. Conditioning on $\pi_{132,321;n} \neq \iota_n$, we see by Proposition 7.1 that $\pi_{132,321;n} = \pi_{K,L,n-K-L}$, where K and L are random with (K, L) uniformly distributed over the set $\{K, L \geq 1 : K + L \leq n\}$. As $n \rightarrow \infty$, we thus have $(K/n, L/n) \xrightarrow{d} (X, Y)$ with (X, Y) uniformly distributed on the triangle $\{(X, Y) \in \mathbb{R}_+^2 : X + Y \leq 1\}$. Equivalently, letting $Z := 1 - X - Y$,

$$\left(\frac{K}{n}, \frac{L}{n}, \frac{n - K - L}{n}\right) \xrightarrow{d} (X, Y, Z) \sim \text{Dir}(1, 1, 1), \quad (7.3)$$

where we recall that the Dirichlet distribution $\text{Dir}(1, 1, 1)$ is the uniform distribution on the simplex $\{(x, y, z) \in \mathbb{R}_+^3 : x + y + z = 1\}$.

If $\sigma = \pi_{i,j,p}$ for some i, j, p , then it is easily seen that an occurrence of σ in $\pi_{k,\ell,m}$ is obtained by selecting i, j and p elements from the three runs of $\pi_{k,\ell,m}$, and thus

$$n_\sigma(\pi_{k,\ell,m}) = \binom{k}{i} \binom{\ell}{j} \binom{m}{p}. \quad (7.4)$$

Similarly, if $\sigma = \iota_i$, then an occurrence of σ in $\pi_{k,\ell,m}$ is obtained by selecting i elements from either the union of the first and last run, or from the union of the two last. Hence, by inclusion-exclusion,

$$n_\sigma(\iota_i) = \binom{k+m}{i} + \binom{\ell+m}{i} - \binom{m}{i}. \quad (7.5)$$

These exact formulas together with the description of $\pi_{132,321;n}$ above and (7.3) yield the following asymptotic result.

Theorem 7.2. *Let $\sigma \in \mathfrak{S}_*(132, 321)$. Then the following hold as $n \rightarrow \infty$.*

(i) *If $\sigma = \pi_{i,j,p}$ for some i, j, p , then*

$$n^{-(i+j+p)} n_\sigma(\pi_{132,321;n}) \xrightarrow{d} W_{i,j,p} := \frac{1}{i!j!p!} X^i Y^j Z^p, \quad (7.6)$$

where $(X, Y, Z) \sim \text{Dir}(1, 1, 1)$.

(ii) If $\sigma = \iota_i$, then

$$n^{-i} n_\sigma(\boldsymbol{\pi}_{132,321;n}) \xrightarrow{d} W_i := \frac{1}{i!} ((X+Z)^i + (Y+Z)^i - Z^i), \quad (7.7)$$

with $(X, Y, Z) \sim \text{Dir}(1, 1, 1)$ as in (i).

Moreover, these hold jointly for any set of such σ , and with convergence of all moments. In particular, in case (i),

$$n^{-(i+j+p)} \mathbb{E} n_\sigma(\boldsymbol{\pi}_{132,321;n}) \xrightarrow{d} \mathbb{E} W_{i,j,p} = \frac{2}{(i+j+p+2)!} \quad (7.8)$$

and in case (ii),

$$n^{-i} \mathbb{E} n_\sigma(\boldsymbol{\pi}_{132,321;n}) \xrightarrow{d} \mathbb{E} W_i = \frac{4i+2}{(i+2)!} \quad (7.9)$$

Proof. The limits in distribution (7.6) and (7.7) hold (with joint convergence) by the discussion before the theorem. Moment convergence holds because the normalized variables in (7.6) and (7.7) are bounded (by 1). Finally, the expectation in (7.8) is easily computed using the multidimensional extension of the beta integral [17, (5.14.2)], which implies

$$\mathbb{E} X^a Y^b Z^c = \frac{2\Gamma(a+1)\Gamma(b+1)\Gamma(c+1)}{\Gamma(a+b+c+3)}, \quad a, b, c > -1. \quad (7.10)$$

For the expectation in (7.9), we note also that $X+Z \stackrel{d}{=} Y+Z \sim B(2, 1)$; the result follows by a short calculation. \square

Higher moments of $W_{i,j,p}$ follow also from (7.10).

Corollary 7.3. *The number of inversions has the asymptotic distribution*

$$n^{-2} n_{21}(\boldsymbol{\pi}_{132,321;n}) \xrightarrow{d} W := XY, \quad (7.11)$$

with (X, Y) as above; the limit variable W has density function

$$2 \log(1 + \sqrt{1-4x}) - 2 \log(1 - \sqrt{1-4x}), \quad 0 < x < 1/4, \quad (7.12)$$

and moments

$$\mathbb{E} W^r = 2 \frac{r!^2}{(2r+2)!}, \quad r > 0. \quad (7.13)$$

Proof. We have $21 = \pi_{1,1,0}$, and thus (7.6) yields (7.11). The formula (11.8) for the moments $\mathbb{E} W^r = \mathbb{E} X^r Y^r$ follow by (7.10). Finally, for $0 < t < 1/4$, $\mathbb{P}(W > t) = \mathbb{P}(XY > t)$ equals 2 times the area of the set $\{(x, y) \in \mathbb{R}_+^2 : x + y \leq 1, xy > t\}$. A differentiation and a simple calculation yield (7.12). \square

Example 7.4. For the four allowed patterns of length 3, we find

$$n^{-3} \mathbb{E} n_{123}(\boldsymbol{\pi}_{132,321;n}) \rightarrow \mathbb{E} W_3 = \frac{7}{60}, \quad (7.14)$$

$$n^{-3} \mathbb{E} n_{213}(\boldsymbol{\pi}_{132,321;n}) \rightarrow \mathbb{E} W_{1,1,1} = \frac{1}{60}, \quad (7.15)$$

$$n^{-3} \mathbb{E} n_{231}(\boldsymbol{\pi}_{132,321;n}) \rightarrow \mathbb{E} W_{2,1,0} = \frac{1}{60}, \quad (7.16)$$

$$\mathbb{E} n_{312}(\boldsymbol{\pi}_{132,321;n}) = \mathbb{E} W_{1,2,0} = \frac{1}{60}. \quad (7.17)$$

(See Zhao [22] for exact formulas for finite n .) Note that by (7.8), all $W_{i,j,q}$ with the same $i+j+q$ have the same expectation; their distributions differ, however, in general, as is shown by higher moments. For example, in the present example, by (7.10), $\mathbb{E} W_{1,1,1}^2 = 2/7!$ and $\mathbb{E} W_{2,1,0}^2 = 3/7!$.

The expected number of occurrences of σ can also easily be found exactly for finite n , as follows. As noted above, (7.8) shows that all σ in (i) of the same length occur in $\boldsymbol{\pi}_{132,321;n}$ with asymptotically equal frequencies. In fact, this holds also exactly, for any n . (Note also that (7.8) is an immediate consequence of (7.18).)

Theorem 7.5. *Let $\sigma = \pi_{i,j,p}$, with $i, j \geq 1$ and $p \geq 0$. Then, for any n ,*

$$\mathbb{E} n_{\sigma}(\boldsymbol{\pi}_{132,321;n}) = \frac{\binom{n+2}{i+j+p+2}}{\binom{n}{2} + 1}. \quad (7.18)$$

Proof. By (7.4) and the discussion before it, for any given k, ℓ, m , the number of occurrences of σ in $\pi_{k,\ell,m}$ equals the number of sequences $q_1, \dots, q_i, q'_1, \dots, q'_j, q''_1, \dots, q''_m$ such that

$$1 \leq q_1 < \dots < q_i \leq k < q'_1 < \dots < q'_j \leq \ell < q''_1 < \dots < q''_m \leq n. \quad (7.19)$$

Since σ does not occur in ι_n , the total number of occurrences of σ in all elements of $\mathfrak{S}_n(132, 321)$ is thus, recalling (7.2), equal to the number of all sequences $(q_1, \dots, q_i, k, q'_1, \dots, q'_j, \ell, q''_1, \dots, q''_m)$ of integers satisfying (7.19). By increasing k and all q'_r by 1, and ℓ and all q''_s by 2, we obtain a bijection with the collection of all subsets of $i+j+q+2$ elements of $\{1, \dots, n+2\}$. Hence, the total number of occurrences is $\binom{n+2}{i+j+p+2}$, and (7.18) follows. \square

8. Avoiding $\{231, 312, 321\}$

We proceed to avoiding sets of three permutations. In this section we avoid $T = \{231, 312, 321\}$. An equivalent set is $\{123, 132, 213\}$.

It was shown by Simion and Schmidt [20] that $|\mathfrak{S}_n(231, 312, 321)| = F_{n+1}$, the $(n+1)$ th Fibonacci number (with the initial conditions $F_0 = 0, F_1 = 1$); they also gave the following characterization (in an equivalent form).

Proposition 8.1 ([20, Proposition 15*]). *A permutation π belongs to the class $\mathfrak{S}_*(231, 312, 321)$ if and only if every block in π is decreasing and has length ≤ 2 , i.e., every block is 1 or 21. \square*

Cf. Proposition 5.1; we have here added the restriction that block lengths are 1 or 2. With this restriction in mind, we use again the notation (5.1) and note that (5.2) holds. A permutation $\pi \in \mathfrak{S}_n(231, 312, 321)$ is thus of the form π_{L_1, \dots, L_B} for some sequence L_1, \dots, L_B of $\{1, 2\}$ with sum n ; furthermore, this yields a bijection with all such sequences.

Define p to be the golden ratio:

$$p := \frac{\sqrt{5} - 1}{2}, \quad (8.1)$$

so that $p + p^2 = 1$. Let X be a random variable with the distribution

$$\mathbb{P}(X = 1) = p, \quad \mathbb{P}(X = 2) = p^2. \quad (8.2)$$

Consider an i.i.d. sequence X_1, X_2, \dots of copies of X , and let $S_n := \sum_{i=1}^n X_i$. Then for any sequence ℓ_1, \dots, ℓ_b with $b \geq 1$, $\ell_i \in \{1, 2\}$ and $\sum_1^b \ell_i = n$,

$$\mathbb{P}(X_i = \ell_i, i = 1, \dots, b) = \prod_{i=1}^b p^{\ell_i} = p^n. \quad (8.3)$$

This probability is thus the same for all such sequences, which means that, conditioned on the event that $S_b = n$ for some (unspecified) $b \geq 1$, the sequence (X_1, \dots, X_b) is equidistributed over all $\{1, 2\}$ -sequences with sum n ; we have seen above that this equals the distribution of the sequence of block lengths (L_1, \dots, L_B) of a random permutation $\pi_{231,312,321;n}$ in $\mathfrak{S}_n(231, 312, 321)$. Consequently, recalling (2.14) and Remark 2.3,

$$(L_1, \dots, L_B) \stackrel{d}{=} ((X_1, \dots, X_{N_+(n)}) \mid S_{N_+(n)} = n). \quad (8.4)$$

It follows from this and (5.2) that if $\sigma = \pi_{\ell_1, \dots, \ell_b} \in \mathfrak{S}_*(231, 312, 321)$, and f is defined by (5.5), then $n_\sigma(\pi_{231,312,321;n})$ has the same distribution as $U_{N_+(n)}$ conditioned on $S_{N_+(n)} = n$. Consequently, Proposition 2.5 applies and yields asymptotic normality of $n_\sigma(\pi_{231,312,321;n})$, and Proposition 2.6 adds moment convergence.

To find the parameters, let σ have b_1 blocks of length 1 and b_2 blocks of length 2 (so $b_1 + b_2 = b$ and $b_1 + 2b_2 = |\sigma|$). Then, noting $\binom{X}{2} = X - 1$,

$$\nu = \mathbb{E} X = p + 2p^2 = 2 - p = \frac{5 - \sqrt{5}}{2}, \quad (8.5)$$

$$\text{Var } X = p^3 = 2p - 1 = \sqrt{5} - 2, \quad (8.6)$$

$$\mathbb{E} \binom{X}{2} = \mathbb{P}(X = 2) = p^2 = 1 - p, \quad (8.7)$$

$$\mu = (2 - p)^{b_1} (1 - p)^{b_2} = \left(\frac{5 - \sqrt{5}}{2}\right)^{b_1} \left(\frac{3 - \sqrt{5}}{2}\right)^{b_2}, \quad (8.8)$$

$$f_i(X) = \begin{cases} (2 - p)^{b_1 - 1} (1 - p)^{b_2} X, & \ell_i = 1, \\ (2 - p)^{b_1} (1 - p)^{b_2 - 1} (X - 1), & \ell_i = 2, \end{cases} \quad (8.9)$$

$$\beta_{ij} = \begin{cases} (2 - p)^{2b_1 - 2} (1 - p)^{2b_2} (2p - 1), & \ell_i = \ell_j = 1, \\ (2 - p)^{2b_1 - 1} (1 - p)^{2b_2 - 1} (2p - 1), & \ell_i = 1 < \ell_j = 2, \\ (2 - p)^{2b_1} (1 - p)^{2b_2 - 2} (2p - 1), & \ell_i = \ell_j = 2. \end{cases} \quad (8.10)$$

$$\text{Cov}(f_i(X), X) = \begin{cases} (2 - p)^{b_1 - 1} (1 - p)^{b_2} (2p - 1), & \ell_i = 1, \\ (2 - p)^{b_1} (1 - p)^{b_2 - 1} (2p - 1), & \ell_i = 2. \end{cases} \quad (8.11)$$

We summarize.

Theorem 8.2. *Let $\sigma \in \mathfrak{S}_m(231, 312, 321)$ have block lengths ℓ_1, \dots, ℓ_b . Then, as $n \rightarrow \infty$,*

$$\frac{n_\sigma(\pi_{231,312,321;n}) - \mu n^b/b!}{n^{b-1/2}} \xrightarrow{d} N(0, \gamma^2), \quad (8.12)$$

where μ is given by (8.8) and γ^2 can be calculated by (2.16) and (8.5)–(8.11).

Moreover, (8.12) holds with convergence of all moments. \square

Example 8.3. For the number of inversions, $\sigma = 21$, $b = 1 = b_2$ and $b_1 = 0$. Hence, $\mu = 1 - p = (3 - \sqrt{5})/2$ and, by a calculation, $\gamma^2 = (2 - p)^{-3} \text{Var } X = 5^{-3/2}$. Consequently,

$$\frac{n_{21}(\pi_{231,312,321;n}) - \frac{3-\sqrt{5}}{2}n}{n^{1/2}} \xrightarrow{d} N(0, 5^{-3/2}). \quad (8.13)$$

Remark 8.4. Again, $\gamma^2 > 0$ unless $\sigma = \iota_m$.

9. Avoiding $\{132, 231, 312\}$

In this section we avoid $\{132, 231, 312\}$. Equivalent sets are $\{132, 213, 231\}$, $\{132, 213, 312\}$, $\{213, 231, 312\}$.

It was shown by Simion and Schmidt [20] that $|\mathfrak{S}_n(132, 231, 312)| = n$, together with the following characterization (in an equivalent form). In this section, let

$$\pi_{k,\ell} := \bar{\iota}_k * \iota_\ell = (k, \dots, 1, k+1, \dots, k+\ell) \in \mathfrak{S}_{k+\ell}, \quad k \geq 1, \ell \geq 0. \quad (9.1)$$

Note that $\pi_{1,\ell} = \iota_{1+\ell}$.

Proposition 9.1 ([20, Proposition 16*]).

$$\mathfrak{S}_n(132, 231, 312) = \{\pi_{k,n-k} : 1 \leq k \leq n\}. \quad \square$$

Cf. Propositions 4.1 and 5.1, which characterize supersets. (Equivalently, $\pi \in \mathfrak{S}_*(132, 231, 312)$ if the first block is decreasing and all other blocks have length 1.)

Hence, the random $\pi_{132,231,312;n} = \pi_{K,n-K}$, where $K \in [n]$ is uniformly random. Obviously, as $n \rightarrow \infty$,

$$K/n \xrightarrow{d} U \sim \text{U}(0, 1). \quad (9.2)$$

Furthermore, if $\sigma = \pi_{k,\ell}$, then it is easy to see that

$$n_\sigma(\pi_{K,n-K}) = \begin{cases} \binom{K}{k} \binom{n-K}{\ell}, & k \geq 2, \\ K \binom{n-K}{\ell} + \binom{n-K}{\ell+1}, & k = 1. \end{cases} \quad (9.3)$$

Theorem 9.2. *Let $\sigma \in \mathfrak{S}_*(132, 231, 312)$. Then the following hold as $n \rightarrow \infty$, with $U \sim \text{U}(0, 1)$.*

(i) If $\sigma = \pi_{k,m-k}$ with $2 \leq k \leq m$, then

$$n^{-m} n_\sigma(\boldsymbol{\pi}_{132,231,312;n}) \xrightarrow{d} W_{k,m-k} := \frac{1}{k!(m-k)!} U^k (1-U)^{m-k}. \quad (9.4)$$

(ii) If $\sigma = \pi_{1,m-1} = \iota_m$, then

$$\begin{aligned} n^{-m} n_\sigma(\boldsymbol{\pi}_{132,231,312;n}) &\xrightarrow{d} W_{1,m-1} := \frac{1}{(m-1)!} U (1-U)^{m-1} + \frac{1}{m!} (1-U)^m \\ &= \frac{1}{m!} (1 + (m-1)U) (1-U)^{m-1}. \end{aligned} \quad (9.5)$$

Moreover, these hold jointly for any set of such σ , and with convergence of all moments. In particular, in case (i),

$$n^{-m} \mathbb{E} n_\sigma(\boldsymbol{\pi}_{132,231,312;n}) \xrightarrow{d} \mathbb{E} W_{k,m-k} = \frac{1}{(m+1)!}, \quad k \geq 2, \quad (9.6)$$

and in case (ii),

$$n^{-m} \mathbb{E} n_\sigma(\boldsymbol{\pi}_{132,231,312;n}) \xrightarrow{d} \mathbb{E} W_{1,m-1} = \frac{2}{(m+1)!}. \quad (9.7)$$

Proof. The limits in distribution (9.4) and (9.5) hold (with joint convergence) by (9.3) and (9.2). Moment convergence holds because the normalized variables in (9.4) and (9.5) are bounded (by 1). Finally, the expectations in (9.6)–(9.7) are computed by standard beta integrals. \square

Corollary 9.3. *The number of inversions has the asymptotic distribution*

$$n^{-2} n_{21}(\boldsymbol{\pi}_{132,231,312;n}) \xrightarrow{d} W := U^2/2 \quad (9.8)$$

with $U \sim \mathcal{U}(0, 1)$. Thus, $2W \sim B(\frac{1}{2}, 1)$, and W has moments

$$\mathbb{E} W^r = \frac{1}{2^r(2r+1)}, \quad r > 0. \quad (9.9)$$

Proof. We have $21 = \pi_{2,0}$ by (9.1), and (9.4) yields (9.8). The remaining statements follow by simple calculations. \square

10. Avoiding $\{132, 231, 321\}$

In this section we avoid $\{132, 231, 321\}$. Equivalent sets are $\{123, 132, 231\}$, $\{123, 213, 312\}$, $\{213, 312, 321\}$, $\{123, 132, 312\}$, $\{123, 213, 231\}$, $\{132, 312, 321\}$, $\{213, 231, 321\}$.

It was shown by Simion and Schmidt [20] that $|\mathfrak{S}_n(132, 231, 321)| = n$, together with the following characterization (in an equivalent form). In this section, let

$$\pi_{k,\ell} := (k, 1, \dots, k-1, k+1, \dots, k+\ell) \in \mathfrak{S}_{k+\ell}, \quad k \geq 1, \ell \geq 0. \quad (10.1)$$

Note that $\pi_{k,\ell}$ equals $\pi_{1,k-1,\ell}$ in the notation (7.1) of Section 7 if $k \geq 2$, and $\iota_{1+\ell}$ if $k = 1$.

Proposition 10.1 ([20, Proposition 16*]).

$$\mathfrak{S}_n(132, 231, 321) = \{\pi_{k, n-k} : 1 \leq k \leq n\}. \quad \square$$

Cf. Proposition 7.1, which characterizes a superset.

Hence, the random $\pi_{132, 231, 321; n} = \pi_{K, n-K}$, where $K \in [n]$ is uniformly random. Obviously, as $n \rightarrow \infty$, (9.2) holds in this case too. Furthermore, if $\sigma = \pi_{k, \ell}$, then it is easy to see, e.g. by (7.4)–(7.5), that

$$n_\sigma(\pi_{K, n-K}) = \begin{cases} \binom{K-1}{k-1} \binom{n-K}{\ell}, & k \geq 2, \\ \binom{n-1}{\ell+1} + \binom{n-K}{\ell}, & k = 1. \end{cases} \quad (10.2)$$

Theorem 10.2. *Let $\sigma \in \mathfrak{S}_*(132, 231, 321)$. Then the following hold as $n \rightarrow \infty$, with $U \sim \mathbf{U}(0, 1)$.*

(i) *If $\sigma = \pi_{k, m-k}$ with $2 \leq k \leq m$, then*

$$n^{-(m-1)} n_\sigma(\pi_{132, 231, 321; n}) \xrightarrow{d} W_{k, m-k} := \frac{1}{(k-1)!(m-k)!} U^{k-1} (1-U)^{m-k}. \quad (10.3)$$

(ii) *If $\sigma = \pi_{1, m-1} = \iota_m$, then*

$$n^{-m} n_\sigma(\pi_{132, 231, 321; n}) = \frac{1}{m!} + O(n^{-1}) \xrightarrow{p} \frac{1}{m!}. \quad (10.4)$$

Moreover, these hold jointly for any set of such σ , and with convergence of all moments. In particular, in case (i),

$$n^{-(m-1)} \mathbb{E} n_\sigma(\pi_{132, 231, 321; n}) \xrightarrow{d} \mathbb{E} W_{k, m-k} = \frac{1}{m!}, \quad k \geq 2. \quad (10.5)$$

Proof. By (10.2) and (9.2), similarly to the proof of Theorem 9.2. \square

Corollary 10.3. *The number of inversions $n_{21}(\pi_{132, 231, 321; n})$ has a uniform distribution on $\{0, \dots, n-1\}$, and thus the asymptotic distribution*

$$n^{-1} n_{21}(\pi_{132, 231, 321; n}) \xrightarrow{d} U \sim \mathbf{U}(0, 1). \quad (10.6)$$

Proof. By (10.1), $12 = \pi_{2, 0}$, and thus (10.2) yields $n_{21}(\pi_{K, n-K}) = K-1$. \square

11. Avoiding $\{132, 213, 321\}$

In this section we avoid $\{132, 213, 321\}$. An equivalent sets is $\{123, 231, 312\}$.

It was shown by Simion and Schmidt [20] that $|\mathfrak{S}_n(132, 213, 321)| = n$, together with the following characterization (in an equivalent form). In this section, let

$$\pi_{k, \ell} := (\ell + 1, \dots, \ell + k, 1, \dots, \ell) \in \mathfrak{S}_{k+\ell}, \quad k \geq 1, \ell \geq 0. \quad (11.1)$$

Note that $\pi_{k, \ell}$ equals $\pi_{k, \ell, 0}$ in the notation (7.1) of Section 7 if $\ell \geq 1$, and ι_k if $\ell = 0$.

Proposition 11.1 ([20, Proposition 16*]).

$$\mathfrak{S}_n(132, 213, 321) = \{\pi_{k, n-k} : 1 \leq k \leq n\}. \quad \square$$

Cf. Proposition 7.1, which again characterizes a superset.

Hence, the random $\pi_{132,213,321;n} = \pi_{K,n-K}$, where $K \in [n]$ is uniformly random, and (9.2) holds again. Furthermore, if $\sigma = \pi_{k,\ell}$, then it is easy to see, e.g. by (7.4)–(7.5), that

$$n_\sigma(\pi_{K,n-K}) = \begin{cases} \binom{K}{k} \binom{n-K}{\ell}, & \ell \geq 1, \\ \binom{K}{k} + \binom{n-K}{k}, & \ell = 0. \end{cases} \quad (11.2)$$

Theorem 11.2. *Let $\sigma \in \mathfrak{S}_*(132, 213, 321)$. Then the following hold as $n \rightarrow \infty$, with $U \sim \mathbf{U}(0, 1)$.*

(i) *If $\sigma = \pi_{k,m-k}$ with $1 \leq k \leq m-1$, then*

$$n^{-m} n_\sigma(\pi_{132,213,321;n}) \xrightarrow{d} W_{k,m-k} := \frac{1}{k!(m-k)!} U^k (1-U)^{m-k}. \quad (11.3)$$

(ii) *If $\sigma = \pi_{m,0} = \iota_m$, then*

$$n^{-m} n_\sigma(\pi_{132,213,321;n}) \xrightarrow{d} W_{m,0} := \frac{1}{m!} (U^m + (1-U)^m). \quad (11.4)$$

Moreover, these hold jointly for any set of such σ , and with convergence of all moments. In particular, in case (i),

$$n^{-m} \mathbb{E} n_\sigma(\pi_{132,213,321;n}) \xrightarrow{d} \mathbb{E} W_{k,m-k} = \frac{1}{(m+1)!}, \quad 1 \leq k < m, \quad (11.5)$$

and in case (ii),

$$n^{-m} \mathbb{E} n_\sigma(\pi_{132,213,321;n}) \xrightarrow{d} \mathbb{E} W_{m,0} = \frac{2}{(m+1)!}. \quad (11.6)$$

Proof. By (11.2) and (9.2), similarly to the proof of Theorem 9.2. \square

Corollary 11.3. *The number of inversions has the asymptotic distribution*

$$n^{-2} n_{21}(\pi_{132,213,321;n}) \xrightarrow{d} W := U(1-U), \quad (11.7)$$

with $U \sim \mathbf{U}(0, 1)$. Thus, $4W \sim B(1, \frac{1}{2})$, and W has moments

$$\mathbb{E} W^r = \frac{\Gamma(r+1)^2}{\Gamma(2r+2)}, \quad r > 0. \quad (11.8)$$

Proof. We have $21 = \pi_{1,1}$ by (11.1), and thus (11.3) yields (11.7). The remaining statements follow by simple calculations, using $4W = 1 - (2U-1)^2$ and a beta integral. \square

References

- [1] Frédérique Bassino, Mathilde Bouvel, Valentin Féray, Lucas Gerin, Mickaël Maazoun & Adeline Pierrot. Universal limits of substitution-closed permutation classes. Preprint, 2017. [arXiv:1706.08333](#).
- [2] Frédérique Bassino, Mathilde Bouvel, Valentin Féray, Lucas Gerin & Adeline Pierrot. The Brownian limit of separable permutations. Preprint, 2016. [arXiv:1602.04960](#).

- [3] Miklós Bóna. *Combinatorics of Permutations*. Chapman & Hall/CRC, Boca Raton, FL, 2004.
- [4] Miklós Bóna. The copies of any permutation pattern are asymptotically normal. Preprint, 2007. [arXiv:0712.2792](https://arxiv.org/abs/0712.2792).
- [5] Miklós Bóna. The absence of a pattern and the occurrences of another. *Discrete Math. Theor. Comput. Sci.* **12** (2010), no. 2, 89–102.
- [6] Miklós Bóna. On three different notions of monotone subsequences. *Permutation Patterns*, 89–114, London Math. Soc. Lecture Note Ser., 376, Cambridge Univ. Press, Cambridge, 2010.
- [7] Miklós Bóna. Surprising symmetries in objects counted by Catalan numbers. *Electron. J. Combin.* **19** (2012), no. 1, Paper 62, 11 pp.
- [8] Louis Comtet. *Advanced Combinatorics*, Reidel, Dordrecht, 1974.
- [9] Wassily Hoeffding. A class of statistics with asymptotically normal distribution. *Ann. Math. Statistics* **19** (1948), 293–325.
- [10] Svante Janson. *Gaussian Hilbert Spaces*. Cambridge Univ. Press, Cambridge, UK, 1997.
- [11] Svante Janson. Brownian excursion area, Wright’s constants in graph enumeration, and other Brownian areas. *Probab. Surv.* **4** (2007), 80–145.
- [12] Svante Janson. Patterns in random permutations avoiding the pattern 132, *Combin. Probab. Comput.*, **26** (2017), 24–51.
- [13] Svante Janson. Patterns in random permutations avoiding the pattern 321. Preprint, 2017. [arXiv:1709.08427](https://arxiv.org/abs/1709.08427).
- [14] Svante Janson. Renewal theory for asymmetric U -statistics. Preprint, 2018. [arXiv:1804.05509](https://arxiv.org/abs/1804.05509).
- [15] Svante Janson, Brian Nakamura and Doron Zeilberger. On the asymptotic statistics of the number of occurrences of multiple permutation patterns. *Journal of Combinatorics* **6** (2015), no. 1-2, 117–143.
- [16] Donald E. Knuth. *The Art of Computer Programming. Vol. 1: Fundamental Algorithms*. 3rd ed., Addison-Wesley, Reading, Mass., 1997.
- [17] *NIST Handbook of Mathematical Functions*. Edited by Frank W. J. Olver, Daniel W. Lozier, Ronald F. Boisvert and Charles W. Clark. Cambridge Univ. Press, 2010.
Also available as *NIST Digital Library of Mathematical Functions*, <http://dlmf.nist.gov/>
- [18] *The On-Line Encyclopedia of Integer Sequences*, N. J. A. Sloane ed., <https://oeis.org>.
- [19] Aaron Robertson, Herbert S. Wilf and Doron Zeilberger. Permutation patterns and continued fractions. *Electron. J. Combin.* **6** (1999), Research Paper 38, 6 pp.
- [20] Rodica Simion and Frank W. Schmidt. Restricted permutations. *European J. Combin.* **6** (1985), no. 4, 383–406.
- [21] Richard P. Stanley. *Enumerative Combinatorics*, Volume 2, Cambridge Univ. Press, Cambridge, 1999.
- [22] Alina F. Y. Zhao. Pattern popularity in multiply restricted permutations. *J. Integer Seq.* **17** (2014), no. 10, Article 14.10.3.

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