

# Predicting Maximal Gaps in Sets of Primes

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## Abstract

Let  $q > r \geq 1$  be coprime integers. Let  $\mathbb{P}_c$  be an increasing sequence of primes  $p$  satisfying two conditions: (i)  $p \equiv r \pmod{q}$  and (ii)  $p$  starts a prime  $k$ -tuple of a particular type. Let  $\pi_c(x)$  be the number of primes in  $\mathbb{P}_c$  not exceeding  $x$ .

We heuristically derive formulas predicting the growth trend of the maximal gap  $G_c(x) = p' - p$  between consecutive primes  $p, p' \in \mathbb{P}_c$  below  $x$ . Computations show that a simple trend formula  $G_c(x) \sim \frac{x}{\pi_c(x)} \cdot (\log \pi_c(x) + O_k(1))$  works well for maximal gaps between initial primes of  $k$ -tuples with  $k \geq 2$  (e. g., twin primes, prime triplets, etc.) in residue class  $r \pmod{q}$ . For  $k = 1$ , however, a more sophisticated formula  $G_c(x) \sim \frac{x}{\pi_c(x)} \cdot (\log \frac{\pi_c^2(x)}{x} + O(\log q))$  gives a better prediction of maximal gap sizes. The latter includes the important special case of maximal gaps between primes ( $k = 1$ ,  $q = 2$ ). In all of the above cases, the distribution of appropriately rescaled maximal gaps  $G_c(x)$  near their respective trend is close to the Gumbel extreme value distribution. Almost all maximal gaps turn out to satisfy the inequality  $G_c(x) \lesssim C_k^{-1} \varphi_k(q) \log^{k+1} x$  (an analog of Cramér's conjecture), where  $C_k$  is the corresponding Hardy–Littlewood constant, and  $\varphi_k(q)$  is an appropriate generalization of Euler's totient function. We conjecture that the number of maximal gaps between primes in  $\mathbb{P}_c$  below  $x$  is  $O_k(\log x)$ .

# 1 Introduction

A *prime gap* is the difference between consecutive prime numbers. The sequence of prime gaps behaves quite erratically (see OEIS [A001223](#) [34]). While the prime number theorem tells us that the *average* gap between primes near  $x$  is about  $\log x$ , the actual gaps near  $x$  can be significantly larger or smaller than  $\log x$ . We call a gap *maximal* if it is strictly greater than all gaps before it. Large gaps between primes have been studied by many authors; see, e.g., [2, 5, 6, 12, 14, 27, 29, 33, 38, 39].

Let  $G(x)$  be the maximal gap between primes not exceeding  $x$ :

$$G(x) = \max_{p_{n+1} \leq x} (p_{n+1} - p_n).$$

Estimating  $G(x)$  is a subtle and delicate problem. Cramér [6] conjectured on probabilistic grounds that  $G(x) = O(\log^2 x)$ , while Shanks [33] heuristically found that  $G(x) \sim \log^2 x$ . Granville [15] heuristically argues that for a certain subsequence of maximal gaps we should expect  $G(x) \sim M \log^2 x$ , with some positive  $M \geq 2e^{-\gamma} > 1$ ; that is,  $\limsup_{x \rightarrow \infty} \frac{G(x)}{\log^2 x} \geq 2e^{-\gamma}$ .

Earlier, we have independently proposed formulas closely related to the Cramér and Shanks conjectures. Wolf [37, 38, 39] expressed the probable size of maximal gaps  $G(x)$  in terms of the prime-counting function  $\pi(x)$ :

$$G(x) \sim \frac{x}{\pi(x)} \cdot \left( \log \frac{\pi^2(x)}{x} + O(1) \right), \quad (1)$$

which suggests an analog of Shanks conjecture  $G(x) \sim \log^2 x - 2 \log x \log \log x + O(\log x)$ ; see also Cadwell [5]. Extending the problem statement to *prime  $k$ -tuples*, Kourbatov [18, 19] empirically tested (for  $x \leq 10^{15}$ ,  $k \leq 7$ ) the following heuristic formula for the probable size of maximal gaps  $G_k(x)$  between prime  $k$ -tuples below  $x$ :

$$G_k(x) \sim a(x) \cdot \left( \log \frac{x}{a(x)} + O(1) \right), \quad (2)$$

where  $a$  is the expected average gap between the particular prime  $k$ -tuples. Similar to (1), formula (2) also suggests an analog of the Shanks conjecture,  $G_k(x) \sim C \log^{k+1} x$ , with a negative correction term of size  $O_k((\log x)^k \log \log x)$ ; see also [11].

In this paper we study a further generalization of the prime gap growth problem, viz.: What happens to maximal gaps if we only look at primes in a specific *residue class mod  $q$* ? The new problem statement subsumes, as special cases, *maximal prime gaps* ( $k = 1$ ,  $q = 2$ ) as well as maximal gaps between *prime  $k$ -tuples* ( $k \geq 2$ ,  $q = 2$ ). One of our present goals is to generalize formulas (1) and (2) to gaps between primes in a residue class — and test them in computational experiments. Another goal is to investigate *how many* maximal gaps should be expected between primes  $p \leq x$  in a residue class, with an additional (optional) condition that  $p$  starts a prime constellation of a certain type.

## 1.1 Notation

$q, r$	coprime integers, $1 \leq r < q$
$p_n$	the $n$ -th prime; $\{p_n\} = \{2, 3, 5, 7, 11, \dots\}$
$\mathbb{P}_c$	increasing sequence of primes $p$ such that (i) $p \equiv r \pmod{q}$ and (ii) $p$ is the least prime in a prime $k$ -tuple of a specific type. <i>Note:</i> $\mathbb{P}_c$ depends on $q, r, k$ , and on the pattern of the $k$ -tuple. When $k = 1$ , $\mathbb{P}_c$ is the sequence of <i>all</i> primes $p \equiv r \pmod{q}$ .
$\gcd(m, n)$	the greatest common divisor of $m$ and $n$
$\varphi(q)$	Euler's totient function (OEIS <a href="#">A000010</a> )
$\varphi_k(q)$	generalization of Euler's totient function (defined in sect. 2.1)
$\text{Gumbel}(x; \alpha, \mu)$	the Gumbel distribution cdf: $\text{Gumbel}(x; \alpha, \mu) = e^{-e^{-\frac{x-\mu}{\alpha}}}$
$\text{Exp}(x; \alpha)$	the exponential distribution cdf: $\text{Exp}(x; \alpha) = 1 - e^{-x/\alpha}$
$\alpha$	the <i>scale parameter</i> of exponential/Gumbel distributions, as applicable
$\mu$	the <i>location parameter (mode)</i> of the Gumbel distribution
$\gamma$	the Euler–Mascheroni constant: $\gamma = 0.57721\dots$
$C_k$	the Hardy–Littlewood constants (see <i>Appendix 5.5</i> )
$\log x$	the natural logarithm of $x$
$\text{li } x$	the logarithmic integral of $x$ : $\text{li } x = \int_0^x \frac{dt}{\log t} = \int_2^x \frac{dt}{\log t} + 1.04516\dots$
$\text{Li}_k(x)$	the integral $\int_2^x \frac{dt}{\log^k t}$ (see <i>Appendix 5.6</i> )
	<i>Gap measure functions:</i>
$G(x)$	the maximal gap between primes $\leq x$
$G_{q,r}(x)$	the maximal gap between primes $p = r + nq \leq x$ (case $k = 1$ )
$G_c(x)$	the maximal gap between primes $p \in \mathbb{P}_c$ not exceeding $x$
$R_c(n)$	the $n$ -th record (maximal) gap between primes $p \in \mathbb{P}_c$
$\bar{a}_c(x)$	the expected average gap between primes in $\mathbb{P}_c$ <i>near</i> $x$ (see sect. 2.2)
$\tilde{a}_c(x)$	the expected average gap between primes in $\mathbb{P}_c$ <i>below</i> $x$ (see sect. 2.2)
$T, \tilde{T}_c, \bar{T}_c$	trend functions predicting the growth of maximal gaps (see sect. 2.3)
	<i>Gap counting functions:</i>
$N_c(x)$	the number of maximal gaps $G_c$ with endpoints $p \leq x$
$N_{q,r}(x)$	the number of maximal gaps $G_{q,r}$ with endpoints $p \leq x$ (case $k = 1$ )
$\tau_{q,r}(d, x)$	the number of gaps of a given even size $d = p' - p$ between successive primes $p, p' \equiv r \pmod{q}$ , with $p' \leq x$ ; $\tau_{q,r}(d, x) = 0$ if $q \nmid d$ or $2 \nmid d$ .
	<i>Prime counting functions:</i>
$\pi(x)$	the total number of primes $p_n \leq x$
$\pi_c(x)$	the total number of primes $p \in \mathbb{P}_c$ not exceeding $x$
$\pi_{q,r}(x)$	the total number of primes $p = r + nq \leq x$ (case $k = 1$ )

Notation  $\pi_{q,r}(x)$  is a compact form of the often-used  $\pi(x; q, r)$ , and  $\pi_c(x)$  is a compact form of  $\pi(x; q, r, k)$ . Quantities with the  $c$  subscript may, in general, depend on  $q, r, k$ , and on the pattern of the prime  $k$ -tuple.

## 1.2 Definitions: prime $k$ -tuples, gaps, sequence $\mathbb{P}_c$

*Prime  $k$ -tuples* are clusters of  $k$  consecutive primes that have an admissible<sup>1</sup> (repeatable) pattern. In what follows, when we speak of a  $k$ -tuple, for certainty we will mean a *densest admissible prime  $k$ -tuple*, with a given  $k \leq 7$ . However, our observations can be extended to other admissible  $k$ -tuples, including those with larger  $k$  and not necessarily densest ones. The densest  $k$ -tuples that exist for a given  $k$  may sometimes be called *prime constellations* or *prime  $k$ -tuplets*. Below are examples of prime  $k$ -tuples with  $k = 2, 4, 6$ .

- *Twin primes* are pairs of consecutive primes that have the form  $(p, p + 2)$ . This is the densest admissible pattern of two primes.
- *Prime quadruplets* are clusters of four consecutive primes of the form  $(p, p + 2, p + 6, p + 8)$ . This is the densest admissible pattern of four primes.
- *Prime sextuplets* are clusters of six consecutive primes of the form  $(p, p + 4, p + 6, p + 10, p + 12, p + 16)$ . This is the densest admissible pattern of six primes.

A *gap* between prime  $k$ -tuples is the distance  $p' - p$  between the initial primes  $p$  and  $p'$  in two consecutive  $k$ -tuples of the same type (i. e., with the same pattern). For example, the gap between twin prime pairs  $(17, 19)$  and  $(29, 31)$  is 12:  $p' - p = 29 - 17 = 12$ .

A *maximal gap* between prime  $k$ -tuples is a gap that is strictly greater than all gaps between preceding  $k$ -tuples of the same type. For example, the gap of size 6 between twin primes  $(5, 7)$  and  $(11, 13)$  is maximal, while the gap (also of size 6) between twin primes  $(11, 13)$  and  $(17, 19)$  is *not* maximal.

Let  $q > r \geq 1$  be coprime integers. Let  $\mathbb{P}_c$  be an increasing sequence of primes  $p$  satisfying two conditions: (i)  $p \equiv r \pmod{q}$  and (ii)  $p$  is the least prime in a prime  $k$ -tuple of a specific type. Importantly,  $\mathbb{P}_c$  depends on  $q, r, k$ , and on the pattern of the  $k$ -tuple. When  $k = 1$ ,  $\mathbb{P}_c$  is the sequence of *all* primes  $p \equiv r \pmod{q}$ . Gaps between primes in  $\mathbb{P}_c$  are defined as differences  $p' - p$  between successive primes  $p, p' \in \mathbb{P}_c$ . As before, a gap is maximal if it is strictly greater than all preceding gaps.

Studying maximal gaps between primes in  $\mathbb{P}_c$  is convenient. Indeed, if the modulo  $q$  used for defining  $\mathbb{P}_c$  is “not too small”, we get *plenty of data* to study maximal gaps; that is, we get many sequences of maximal gaps corresponding to  $\mathbb{P}_c$ 's with different  $r$  for the same  $q$ , which allows us to study common properties of these sequences. (One such property is the *average* number of maximal gaps between primes in  $\mathbb{P}_c$  below  $x$ .) By contrast, data on maximal prime gaps are scarce: at present we know only 80 maximal gaps between primes below  $2^{64}$  [27]. Even fewer maximal gaps are known between  $k$ -tuples of any given type [19].

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<sup>1</sup>A prime  $k$ -tuple with a given pattern is admissible (repeatable) unless it is prohibited by an elementary divisibility argument. For example, the cluster of five numbers  $(p, p + 2, p + 4, p + 6, p + 8)$  is prohibited because one of the numbers is divisible by 5 (and, moreover, *at least one* of the numbers is divisible by 3); hence all these five numbers cannot simultaneously be prime infinitely often. Likewise, the cluster of three numbers  $(p, p + 2, p + 4)$  is prohibited because one of the numbers is divisible by 3; so these three numbers cannot simultaneously be prime infinitely often.

## 2 Heuristics and conjectures

### 2.1 Equidistribution of $k$ -tuples

Everywhere we assume that  $q > r$  are coprime positive integers. Let  $\pi_{q,r}(x)$  be the number of primes  $p \equiv r \pmod{q}$  such that  $p \leq x$ . The prime number theorem for arithmetic progressions [9, p. 190] establishes that

$$\pi_{q,r}(x) \sim \frac{\text{li } x}{\varphi(q)} \quad \text{as } x \rightarrow \infty. \quad (3)$$

Furthermore, the generalized Riemann hypothesis (GRH) is equivalent to the statement

$$\pi_{q,r}(x) = \frac{\text{li } x}{\varphi(q)} + O_q(x^{1/2+\varepsilon}) \quad \text{for any } \varepsilon > 0. \quad (4)$$

That is to say, the primes below  $x$  are approximately equally distributed among the  $\varphi(q)$  “admissible” residue classes modulo  $q$ ; roughly speaking, the GRH implies that, as  $x \rightarrow \infty$ , the numbers  $\pi_{q,r}(x)$  and  $\lfloor \text{li } x / \varphi(q) \rfloor$  almost agree in the left half of their digits.

Based on empirical evidence, below we conjecture that a similar phenomenon also occurs for prime  $k$ -tuples: in every admissible residue class modulo  $q$ , there are infinitely many primes starting an admissible  $k$ -tuple of a particular type. Moreover, such primes are distributed approximately equally among all admissible residue classes modulo  $q$ . Our conjectures are closely related to the Hardy–Littlewood  $k$ -tuple conjecture [17] and the Bateman–Horn conjecture [3].

#### 2.1.1 Counting admissible residue classes

First, consider an example: Which residue classes modulo 4 may contain the lesser prime  $p$  in a pair of twin primes  $(p, p + 2)$ ? Clearly, the residue class 0 mod 4 is prohibited: all numbers in this class are even. The residue class 2 mod 4 is prohibited for the same reason. The remaining residue classes,  $p \equiv 1 \pmod{4}$  and  $p \equiv 3 \pmod{4}$ , are not prohibited. We call these two classes *admissible*. Indeed, each of these two admissible residue classes does contain lesser twin primes — and there are, conjecturally, infinitely many such primes in each admissible class (see OEIS [A071695](#) and [A071698](#)).

In general, given a  $k$ -tuple  $(p, p + \Delta_2, p + \Delta_3, \dots, p + \Delta_k)$ , we say that the residue class  $r \pmod{q}$  is *admissible* if

$$\gcd(r, q) = \gcd(r + \Delta_2, q) = \gcd(r + \Delta_3, q) = \dots = \gcd(r + \Delta_k, q) = 1.$$

That is, the residue class  $r \pmod{q}$  is admissible if it is not prohibited (by divisibility considerations) from containing infinitely many primes  $p$  starting a prime  $k$ -tuple of the given type.

How many residue classes modulo  $q$  are admissible for a given prime  $k$ -tuple? To count admissible residue classes  $\pmod{q}$ , we will need an appropriate generalization of *Euler’s totient function*  $\varphi(q)$ .

**Definition of  $\varphi_k(q)$ .** Let  $\varphi_k(q)$  be the number of admissible residue classes modulo  $q$  that are allowed (by divisibility considerations) to contain infinitely many primes  $p$  starting a densest admissible prime  $k$ -tuple of a particular type,  $(p + \Delta_1, p + \Delta_2, p + \Delta_3, \dots, p + \Delta_k)$ , where  $0 = \Delta_1 < \Delta_2 < \Delta_3 < \dots < \Delta_k$ . More formally, let

$$\begin{aligned}\varphi_k(q) &= \sum_{x=1}^q \nu_{q,k}(x), \\ \nu_{q,k}(x) &= \begin{cases} 1 & \text{if } \gcd(x + \Delta_1, q) = \gcd(x + \Delta_2, q) = \dots = \gcd(x + \Delta_k, q) = 1, \\ 0 & \text{otherwise.} \end{cases}\end{aligned}$$

*Example.* For prime quadruplets  $(p, p + 2, p + 6, p + 8)$  we have

$$k = 4, \quad \Delta_1 = 0, \quad \Delta_2 = 2, \quad \Delta_3 = 6, \quad \Delta_4 = 8,$$

and  $\varphi_4(q) = \text{A319516}(q)$ . For instance, when  $q = 30$ , we have  $\varphi_4(q) = 1$ : indeed, there is only one residue class (namely,  $p \equiv 11 \pmod{30}$ ) where divisibility considerations allow infinitely many primes  $p$  at the beginning of quadruplets  $(p, p + 2, p + 6, p + 8)$ .

Note that  $\varphi_1(q) = \varphi(q)$  is Euler's totient function, [A000010](#);  $\varphi_2(q)$  is Alder's totient function  $\varphi(q, 2)$ , see [A002472](#) [1];  $\varphi_3(q)$  is [A319534](#);  $\varphi_4(q)$  is [A319516](#);  $\varphi_5(q)$  is [A321029](#); and  $\varphi_6(q)$  is [A321030](#).

### 2.1.2 The $k$ -tuple infinitude conjecture

We expect each of the  $\varphi_k(q)$  admissible residue classes to contain infinitely many primes  $p$  starting an admissible prime  $k$ -tuple  $(p, p + \Delta_2, p + \Delta_3, \dots, p + \Delta_k)$ . In other words, the corresponding sequence  $\mathbb{P}_c$  is *infinite*.

*Remarks.* (i) This conjecture generalizes Dirichlet's theorem on arithmetic progressions [7].  
(ii) The conjecture follows from the Bateman–Horn conjecture [3].

### 2.1.3 The $k$ -tuple equidistribution conjecture

The number of primes  $p \in \mathbb{P}_c$ ,  $p \leq x$ , is

$$\pi_c(x) = \frac{C_k}{\varphi_k(q)} \text{Li}_k(x) + O_{q,k}(x^\eta) \quad \text{as } x \rightarrow \infty, \quad (5)$$

where  $\eta < 1$ , the coefficient  $C_k$  is the *Hardy–Littlewood constant* for the particular  $k$ -tuple (see *Appendix 5.5*), and  $\varphi_k(q)$  is an appropriate generalization of Euler's totient function (defined in sect. 2.1.1).

*Remarks.*

- (i) This conjecture is akin to the GRH statement (4); the latter pertains to the case  $k = 1$ .
- (ii) The conjecture is compatible with the Bateman–Horn and Hardy–Littlewood  $k$ -tuple conjectures but does not follow from them.
- (iii) It is plausible that, similar to (4), in (5) we can take  $\eta = \frac{1}{2} + \varepsilon$  for any  $\varepsilon > 0$ .

## 2.2 Average gap sizes

We define the *expected average gaps* between primes in  $\mathbb{P}_c$  as follows.

**Definition of  $\tilde{a}_c(x)$ .** The expected average gap between primes in  $\mathbb{P}_c$  *below*  $x$  is

$$\tilde{a}_c(x) = \frac{\varphi_k(q)}{C_k} \cdot \frac{x}{\text{Li}_k(x)}. \quad (6)$$

**Definition of  $\bar{a}_c(x)$ .** The expected average gap between primes in  $\mathbb{P}_c$  *near*  $x$  is

$$\bar{a}_c(x) = \frac{\varphi_k(q)}{C_k} \cdot \log^k x. \quad (7)$$

In view of the equidistribution conjecture (5), it is easy to see from these definitions that

$$\frac{x}{\pi_c(x)} \approx \tilde{a}_c(x) < \bar{a}_c(x) \quad \text{as } x \rightarrow \infty. \quad (8)$$

We have the limits (with very slow convergence):

$$\lim_{x \rightarrow \infty} \frac{\tilde{a}_c(x)}{\bar{a}_c(x)} = 1, \quad (9)$$

$$\lim_{x \rightarrow \infty} \frac{\bar{a}_c(x) - \tilde{a}_c(x)}{\bar{a}_c(x)} \cdot \log x = k. \quad (10)$$

## 2.3 Maximal gap sizes

Recall that formula (1) is applicable to the special case  $q = 2, k = 1$  [37, 38, 39], while (2) is applicable to the special cases  $q = 2, k \geq 2$  [18]. We are now ready to generalize (1) and (2) for predicting maximal gaps between primes in sequences  $\mathbb{P}_c$ .

### 2.3.1 Case of $k$ -tuples: $k \geq 2$

Consider a probabilistic example. Suppose that intervals between rare random events are exponentially distributed, with cdf  $\text{Exp}(\xi; \alpha) = 1 - e^{-\xi/\alpha}$ , where  $\alpha$  is the mean interval between events. If our observations of the events continue for  $x$  seconds, extreme value theory (EVT) predicts that the most probable maximal interval between events is

$$\text{most probable maximal interval} = \alpha \log \frac{x}{\alpha} + O(\alpha) = \frac{x}{\Pi(x)} \log \Pi(x) + O(\alpha), \quad (11)$$

where  $\Pi(x) \approx x/\alpha$  is the total count of the events we observed in  $x$  seconds. (For details on deriving eq. (11), see e.g. [18, sect. 8].)

By analogy with EVT, we define the expected trend functions for maximal gaps as follows.

**Definition of  $\tilde{T}_c(x)$ .** The *lower trend* of maximal gaps between primes in  $\mathbb{P}_c$  is

$$\tilde{T}_c(x) = \tilde{a}_c(x) \cdot \log \frac{C_k \text{Li}_k(x)}{\varphi_k(q)}. \quad (12)$$

In view of the equidistribution conjecture (5),

$$\tilde{T}_c(x) \approx \tilde{a}_c(x) \cdot \log \pi_c(x) \approx \frac{x}{\pi_c(x)} \cdot \log \pi_c(x) \quad \text{as } x \rightarrow \infty. \quad (13)$$

We also define another trend function,  $\bar{T}_c(x)$ , that is simpler because it does not use  $\text{Li}_k(x)$ .

**Definition of  $\bar{T}_c(x)$ .** The *upper trend* of maximal gaps between primes in  $\mathbb{P}_c$  is

$$\bar{T}_c(x) = \bar{a}_c(x) \cdot \log \frac{x}{\bar{a}_c(x)}. \quad (14)$$

These definitions imply that

$$\tilde{T}_c(x) < \bar{T}_c(x) < C_k^{-1} \varphi_k(q) \cdot \log^{k+1} x \quad \text{as } x \rightarrow \infty; \quad (15)$$

at the same time, we have the asymptotic equivalence:

$$\tilde{T}_c(x) \sim \bar{T}_c(x) \sim C_k^{-1} \varphi_k(q) \cdot \log^{k+1} x \quad \text{as } x \rightarrow \infty. \quad (16)$$

We have the limits (convergence is quite slow):

$$\lim_{x \rightarrow \infty} \frac{\bar{T}_c(x) - \tilde{T}_c(x)}{\bar{a}_c(x)} = k, \quad (17)$$

$$\lim_{x \rightarrow \infty} \frac{C_k^{-1} \varphi_k(q) \log^{k+1} p - \bar{T}_c(x)}{\bar{a}_c(x) \log \log x} = k. \quad (18)$$

Therefore,  $\bar{T}_c(x) - \tilde{T}_c(x) = O_k(\bar{a}_c)$ , while  $C_k^{-1} \varphi_k(q) \log^{k+1} p - \bar{T}_c(x) = O_k(\bar{a}_c \log \log x)$ .

We make the following conjectures regarding the behavior of maximal gaps  $G_c(x)$ .

**Conjecture on the trend of  $G_c(x)$ .** For any sequence  $\mathbb{P}_c$  with  $k \geq 2$ , a positive proportion of maximal gaps  $G_c(x)$  satisfy the double inequality

$$\tilde{T}_c(x) \lesssim G_c(x) \lesssim \bar{T}_c(x) \quad \text{as } x \rightarrow \infty, \quad (19)$$

and the difference  $G_c(x) - \bar{T}_c(x)$  changes its sign infinitely often.

**Generalized Cramér conjecture for  $G_c(p)$ .** Almost all maximal gaps  $G_c(p)$  satisfy

$$G_c(p) < C_k^{-1} \varphi_k(q) \log^{k+1} p. \quad (20)$$

**Generalized Shanks conjecture for  $G_c(p)$ .** Almost all maximal gaps  $G_c(p)$  satisfy

$$G_c(p) \sim C_k^{-1} \varphi_k(q) \log^{k+1} p \quad \text{as } p \rightarrow \infty. \quad (21)$$

Here  $G_c(p)$  denotes the maximal gap that ends at the prime  $p$ .



### 2.3.2 Case of primes: $k = 1$

The EVT-based trend formulas (12), (14) work well for maximal gaps between  $k$ -tuples,  $k \geq 2$ . However, when  $k = 1$ , the observed sizes of maximal gaps  $G_{q,r}(x)$  between primes in residue class  $r \pmod q$  are usually a little less than predicted by the corresponding *lower* trend formula akin to (12). For example, with  $k = 1$  and  $q = 2$ , the most probable values of maximal prime gaps  $G(x)$  turn out to be less than the EVT-predicted value  $\frac{x \log \text{li } x}{\text{li } x}$  — less by approximately  $\log x \log \log x$  (cf. Cadwell [5, p. 912]). In this respect, primes do not behave like “random darts”. Instead, the situation looks as if primes “conspire together” so that each prime  $p_n \leq x$  lowers the typical maximal gap  $G(x)$  by about  $p_n^{-1} \log x$ ; indeed, we have  $\sum_{p_n \leq x} p_n^{-1} \sim \log \log x$ . Below we offer a heuristic explanation of this phenomenon.

Let  $\tau_{q,r}(d, x)$  be the number of gaps of a given even size  $d = p' - p$  between successive primes  $p, p' \equiv r \pmod q$ ,  $p' \leq x$ . Empirically, the function  $\tau_{q,r}$  has the form (cf. [38, p. 5])

$$\tau_{q,r}(d, x) \approx P_q(d) B_q(x) e^{-d \cdot A_q(x)}, \quad (22)$$

where  $P_q(d)$  is an oscillating factor, and

$$\tau_{q,r}(d, x) = P_q(d) = 0 \quad \text{if } q \nmid d \text{ or } 2 \nmid d. \quad (23)$$

The essential point now is that we can find the unknown functions  $A_q(x)$  and  $B_q(x)$  in (22) just by assuming the exponential decay of  $\tau_{q,r}$  as a function of  $d$  and employing the following two conditions (which are true by definition of  $\tau_{q,r}$ ):

$$(a) \quad \text{the total number of gaps is } \sum_{d=2}^{G_{q,r}(x)} \tau_{q,r}(d, x) \approx \pi_{q,r}(x); \quad (24)$$

$$(b) \quad \text{the total length of gaps is } \sum_{d=2}^{G_{q,r}(x)} d \cdot \tau_{q,r}(d, x) \approx x. \quad (25)$$

The erratic behavior of the oscillating factor  $P_q(d)$  presents an obstacle in the calculation of sums (24) and (25). We will assume that, for sufficiently regular functions  $f(d, x)$ ,

$$\sum_d P_q(d) f(d, x) \approx s \sum_d f(d, x), \quad (26)$$

where  $s$  is such that, on average,  $P_q(d) \approx s$ ; and the summation is for  $d$  such that both sides of (26) are non-zero. Extending the summation in (24), (25) to infinity, using (26), and writing<sup>2</sup>  $d = cj$ ,  $j \in \mathbb{N}$ , we obtain two series expressions: (24) gives us a geometric series

$$\sum_{d=2}^{\infty} \tau_{q,r}(d, x) \approx s B_q(x) \sum_{j=1}^{\infty} e^{-cj A_q(x)} = s B_q(x) \cdot \frac{e^{-c A_q(x)}}{1 - e^{-c A_q(x)}} \approx \pi_{q,r}(x), \quad (27)$$

---

<sup>2</sup>In view of (23), in the representation  $d = cj$  we must use  $c = \text{LCM}(2, q) = O(q)$ . Accordingly, the factor  $s$  in (26) can be naturally defined as  $s = \lim_{n \rightarrow \infty} \frac{1}{n} \sum_{j=1}^n P_q(cj)$ .

while (25) yields a differentiated geometric series

$$\sum_{d=2}^{\infty} d \cdot \tau_{q,r}(d, x) \approx csB_q(x) \sum_{j=1}^{\infty} j e^{-cjA_q(x)} = csB_q(x) \cdot \frac{e^{-cA_q(x)}}{(1 - e^{-cA_q(x)})^2} \approx x. \quad (28)$$

Thus we have obtained two equations:

$$sB_q(x) \cdot \frac{e^{-cA_q(x)}}{1 - e^{-cA_q(x)}} \approx \pi_{q,r}(x), \quad csB_q(x) \cdot \frac{e^{-cA_q(x)}}{(1 - e^{-cA_q(x)})^2} \approx x.$$

To solve these equations, we use the approximations  $e^{-cA_q(x)} \approx 1$  and  $1 - e^{-cA_q(x)} \approx cA_q(x)$  (which is justified because we expect  $A_q(x) \rightarrow 0$  for large  $x$ ). In this way we obtain

$$A_q(x) \approx \frac{\pi_{q,r}(x)}{x}, \quad B_q(x) \approx \frac{c\pi_{q,r}^2(x)}{sx}. \quad (29)$$

A posteriori we indeed see that  $A_q(x) \rightarrow 0$  as  $x \rightarrow \infty$ . Substituting (29) into (22) we get

$$\tau_{q,r}(d, x) \approx P_q(d) \frac{c\pi_{q,r}^2(x)}{sx} e^{-d \cdot \pi_{q,r}(x)/x}. \quad (30)$$

From (30) we can obtain an approximate formula for  $G_{q,r}(x)$ . Note that  $\tau_{q,r}(d, x) = 1$  when the gap of size  $d$  is maximal (and/or a first occurrence); in either of these cases we have  $d \approx G_{q,r}(x)$ . So, to get an approximate value of the maximal gap  $G_{q,r}(x)$ , we solve for  $d$  the equation  $\tau_{q,r}(d, x) = 1$ , or

$$\frac{c\pi_{q,r}^2(x)}{x} e^{-d \cdot \pi_{q,r}(x)/x} \approx 1, \quad (31)$$

where we skipped  $P_q(d)/s$  because, on average,  $P_q(d) \approx s$ . Taking the log of both sides of (31) we find the solution  $G_{q,r}(x)$  expressed directly in terms of  $\pi_{q,r}(x)$ :

$$G_{q,r}(x) \approx \frac{x}{\pi_{q,r}(x)} \cdot \left( \log \frac{\pi_{q,r}^2(x)}{x} + \log c \right). \quad (32)$$

Since  $\pi_{q,r}(x) \approx \frac{\text{li } x}{\varphi(q)}$  and  $\log \frac{\pi_{q,r}^2(x)}{x} \approx 2 \log \frac{\text{li } x}{\varphi(q)} - \log x$ , we can state the following

**Conjecture on the trend of  $G_{q,r}(x)$ .** The most probable sizes of maximal gaps  $G_{q,r}(x)$  are near a trend curve  $T(q, x)$ :

$$G_{q,r}(x) \sim T(q, x) = \frac{\varphi(q)x}{\text{li } x} \cdot \left( 2 \log \frac{\text{li } x}{\varphi(q)} - \log x + b \right), \quad (33)$$

where  $b = b(q, x) = O(\log q)$  tends to a constant as  $x \rightarrow \infty$ . The difference  $G_{q,r}(x) - T(q, x)$  changes its sign infinitely often.

Further, we expect that the *width of distribution* of the maximal gaps near  $x$  is  $O_q(\log x)$ , i.e., the width of distribution is on the order of the average gap  $\varphi(q) \log x$ . (This can be heuristically justified by extreme value theory — and agrees with numerical results of sect.3.2.) On the other hand, for large  $x$ , the trend (33) differs from the line  $\varphi(q) \log^2 x$  by  $O_q(\log x \log \log x)$ , that is, by much more than the average gap. This suggests natural generalizations of the Cramér and Shanks conjectures:

**Generalized Cramér conjecture for  $G_{q,r}(p)$ .** Almost all maximal gaps  $G_{q,r}(p)$  satisfy

$$G_{q,r}(p) < \varphi(q) \log^2 p. \quad (34)$$

**Generalized Shanks conjecture for  $G_{q,r}(p)$ .** Almost all maximal gaps  $G_{q,r}(p)$  satisfy

$$G_{q,r}(p) \sim \varphi(q) \log^2 p \quad \text{as } p \rightarrow \infty. \quad (35)$$

Conjectures (34) and (35) can be viewed as particular cases of (20), (21) for  $k = 1$ .

## 2.4 How many maximal gaps are there?

This section generalizes the heuristic reasoning of [24, sect.2.3]. Let  $R_c(n)$  be the size of the  $n$ -th record (maximal) gap between primes in  $\mathbb{P}_c$ . Denote by  $N_c(x)$  the total number of maximal gaps observed between primes in  $\mathbb{P}_c$  not exceeding  $x$ .

Let  $\ell = \ell(x; q, k)$  be a continuous function estimating  $\text{mean}_r(N_c(ex) - N_c(x))$ , the average number of maximal gaps between primes in  $\mathbb{P}_c$ , with the upper endpoints  $p' \in [x, ex]$ . For  $x \rightarrow \infty$ , we will heuristically argue that if the limit of  $\ell$  exists, then the limit is  $k + 1$ . We assume that  $\ell(x; q, k) \rightarrow \ell_*$  as  $x \rightarrow \infty$ , and the limit  $\ell_*$  is independent of  $q$ . Let  $n$  be a “typical” number of maximal gaps up to  $x$ ; our assumption  $\lim_{x \rightarrow \infty} \ell = \ell_*$  means that

$$n \sim \ell_* \log x \quad \text{as } x \rightarrow \infty. \quad (36)$$

For large  $n$ , we can estimate the order of magnitude of the typical  $n$ -th maximal gap  $R_c(n)$  using the generalized Cramér and Shanks conjectures (20), (21):

$$R_c(n) = G_c(x) \lesssim C_k^{-1} \varphi_k(q) \log^{k+1} x \sim C_k^{-1} \varphi_k(q) \frac{n^{k+1}}{\ell_*^{k+1}}. \quad (37)$$

Define  $\Delta R_c(n) = R_c(n+1) - R_c(n)$ . By formula (37), for large  $q$  and large  $n$  we have

$$\begin{aligned} \text{mean}_r R_c(n) &\sim C_k^{-1} \varphi_k(q) \frac{n^{k+1}}{\ell_*^{k+1}}, \\ \text{mean}_r \Delta R_c(n) &= \text{mean}_r (R_c(n+1) - R_c(n)) \\ &\sim \frac{C_k^{-1} \varphi_k(q)}{\ell_*^{k+1}} \cdot ((n+1)^{k+1} - n^{k+1}) \\ &\sim \frac{C_k^{-1} \varphi_k(q)}{\ell_*^{k+1}} \cdot (k+1)n^k, \end{aligned}$$

where the mean is taken over all admissible residue classes; see sect. 2.1.1. Combining this with (36) we find

$$\text{mean}_r \Delta R_c(n) \sim \frac{k+1}{\ell_*} \cdot C_k^{-1} \varphi_k(q) \log^k x. \quad (38)$$

On the other hand, heuristically we expect that, on average, two consecutive record gaps should differ by the “local” average gap (7) between primes in  $\mathbb{P}_c$ :

$$\text{mean}_r \Delta R_c(n) \sim C_k^{-1} \varphi_k(q) \log^k x \quad (\sim \text{average gap near } x). \quad (39)$$

Together, equations (38) and (39) imply that

$$\ell_* = k + 1.$$

Therefore, for large  $x$  we should expect (cf. sect. 3.3, 3.4)

$$N_c(x) \sim (k+1) \log x \quad \text{as } x \rightarrow \infty. \quad (40)$$

In particular, for the number  $N_{q,r}$  of maximal gaps between primes  $p \equiv r \pmod{q}$  we have

$$N_{q,r}(x) \sim 2 \log x \quad \text{as } x \rightarrow \infty. \quad (41)$$

*Remark.* Earlier we gave a semi-empirical formula for the number of maximal prime gaps up to  $x$  (i.e., for the special case  $k = 1$ ,  $q = 2$ ) which is asymptotically equivalent to (41):

$$N_{2,1}(x) \sim 2 \log \text{li } x \quad \text{as } x \rightarrow \infty \quad [23, \text{sect. 3.4, OEIS } \underline{\text{A005669}}]. \quad (42)$$

In essence, formula (42) tells us that maximal prime gaps occur, on average, about *twice* as often as records in an i.i.d. random sequence of  $[\text{li } x]$  terms. Note also the following straightforward generalization of (42) giving a very rough estimate of  $N_{q,r}(x)$  in the general case:

$$N_{q,r}(x) \approx \max \left( 0, 2 \log \frac{\text{li } x}{\varphi(q)} \right) \quad [23, \text{eq. 10}]. \quad (43)$$

Computation shows that, for the special case of maximal prime gaps  $G(x)$ , formula (42) works quite well. However, the more general formula (43) usually *overestimates*  $N_{q,r}(x)$ . At the same time, the right-hand side of (43) is *less than*  $2 \log x$ . Thus the right-hand sides of (41) as well as (43) overestimate the actual gap counts  $N_{q,r}(x)$  in most cases.

In section 3.3 we will see an alternative (a posteriori) approximation based on the average number of maximal gaps observed for primes in the interval  $[x, ex]$ . Namely, the estimated average number  $\ell(x; q, k)$  of maximal gaps with endpoints in  $[x, ex]$  is (see Figs. 6–9)

$$\ell(x; q, k) \approx \text{mean}_r(N_c(ex) - N_c(x)) \approx k + 1 - \frac{\kappa(q, k)}{\log x + \delta(q, k)}. \quad (44)$$

### 3 Numerical results

To test our conjectures of the previous section, we performed extensive computational experiments. We used PARI/GP (see *Appendix* for code examples) to compute maximal gaps  $G_c$  between initial primes  $p = r + nq \in \mathbb{P}_c$  in densest admissible prime  $k$ -tuples,  $k \leq 6$ . We experimented with many different values of  $q \in [4, 10^5]$ . To assemble a complete data set of maximal gaps for a given  $q$ , we used all admissible residue classes  $r \pmod{q}$ . For additional details of our computational experiments with maximal gaps between primes  $p = r + nq$  (i.e., for the case  $k = 1$ ), see also [23, sect. 3].

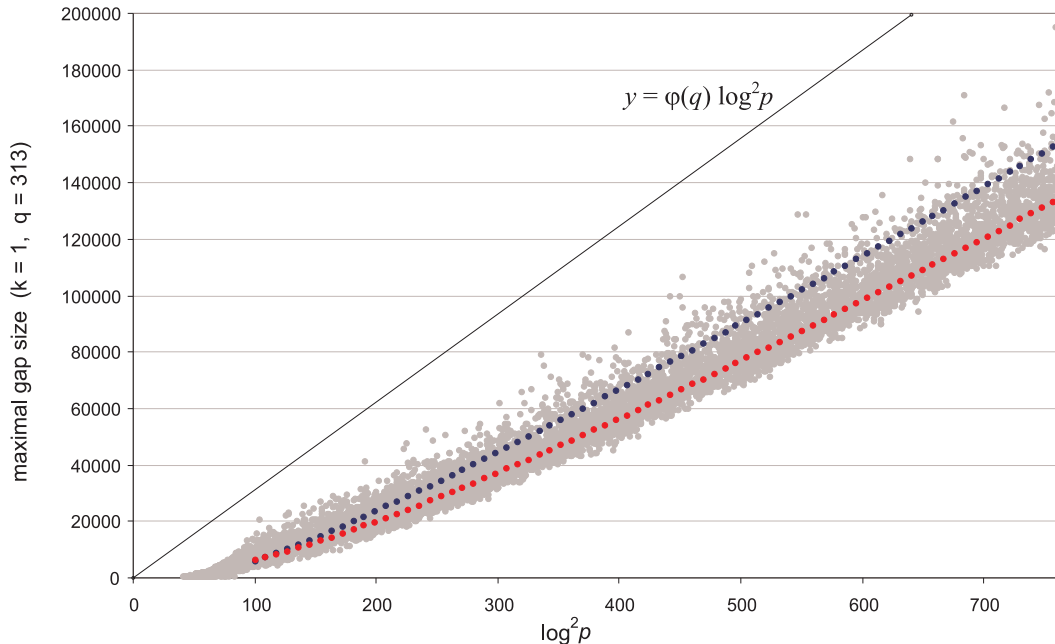


Figure 1: Maximal gaps  $G_{q,r}$  between primes  $p = r + nq \leq x$  for  $q = 313$ ,  $x < 10^{12}$ . Red curve: trend (33), (45); blue curve: EVT-based trend  $\frac{\varphi(q)x}{\text{li } x} \log \frac{\text{li } x}{\varphi(q)}$ ; top line:  $y = \varphi(q) \log^2 p$ .

#### 3.1 The growth trend of maximal gaps

The vast majority of maximal gap sizes  $G_c(x)$  are indeed observed near the trend curves predicted in section 2.3. Specifically, for maximal gaps  $G_c$  between primes  $p = r + nq \in \mathbb{P}_c$  in  $k$ -tuples ( $k \geq 2$ ), the gap sizes are mostly within  $O(\bar{a}_c)$  (that is, within  $O_q(\log^k x)$ ) of the corresponding trend curves of eqs. (12), (14) derived from extreme value theory. However, for  $k = 1$ , the trend eq. (33) gives a better prediction of maximal gaps  $G_{q,r}$ . Figures 1–3 illustrate our numerical results for  $k = 1, 2, 6$ ,  $q = 313$ . The horizontal axis in these figures is  $\log^{k+1} p$  for *end-of-gap* primes  $p$ . Note that all gaps shown in the figures satisfy the generalized

Cramér conjecture, i.e., inequalities (20), (34); for rare exceptions, see *Appendix 5.4*. Results for other values of  $q$  look similar to Figs. 1–3. Numerical evidence suggests that

- For  $k = 1$  (the case of maximal gaps  $G_{q,r}$  between primes  $p = r + nq$ ) the EVT-based trend curve  $\frac{\varphi(q)x}{\text{li } x} \log \frac{\text{li } x}{\varphi(q)}$  goes too high (Fig. 1, blue curve). Meanwhile, the trend (33)

$$T(q, x) = \frac{\varphi(q)x}{\text{li } x} \cdot \left( 2 \log \frac{\text{li } x}{\varphi(q)} - \log x + b \right) \quad (\text{Fig. 1, red curve})$$

satisfactorily predicts gap sizes  $G_{q,r}(x)$ , with the empirical correction term

$$b = b(q, x) \approx \left( b_0 + \frac{b_1}{(\log \log x)^{b_2}} \right) \log \varphi(q) \asymp \log \varphi(q), \quad (45)$$

where the parameter values

$$b_0 = 1, \quad b_1 = 4, \quad b_2 = 2.7 \quad (46)$$

are close to optimal<sup>3</sup> for  $q \in [10^2, 10^5]$  and  $x \in [10^7, 10^{12}]$ .

- For  $k = 2$ , approximately half of maximal gaps  $G_c$  between lesser twin primes  $p \in \mathbb{P}_c$  are below the *lower* trend curve  $\tilde{T}_c(x)$  of eq. (12), while the other half are above that curve; see Fig. 2.
- For  $k \geq 3$ , more than half of maximal gaps  $G_c$  are usually above the lower trend curve  $\tilde{T}_c(x)$  of eq. (12). At the same time, more than half of maximal gaps are usually below the upper trend curve  $\bar{T}_c(x)$  of eq. (14); see Fig. 3. Recall that the two trend curves  $\tilde{T}_c$  and  $\bar{T}_c$  are within  $k\bar{a}_c$  from each other as  $x \rightarrow \infty$ ; see (17).

As noted by Brent [4], twin primes seem to be more random than primes. We can add that, likewise, maximal gaps  $G_{q,r}$  between primes in a residue class seem to be somewhat less random than those for prime  $k$ -tuples; primes  $p \equiv r \pmod{q}$  do not go quite as far from each other as we would expect based on extreme value theory. Pintz [30] discusses various other aspects of the “random” and not-so-random behavior of primes.

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<sup>3</sup>Here the qualifier *optimal* is to be understood in conjunction with the rescaling transformation (47) introduced below in sect. 3.2. A trend  $T(q, x)$  is optimal if after transformation (47) the most probable rescaled values  $w$  turn out to be near zero, and the mode of best-fit Gumbel distribution for  $w$ -values is also close to zero,  $\mu \approx 0$ ; see Fig. 4. In view of (45) it is possible that, for all  $q$ , the “optimal” term  $b$  in (33) has the form  $b(q, x) = (1 + \beta(q, x)) \cdot \log \varphi(q) \sim \log \varphi(q)$ , where  $\beta(q, x)$  very slowly decreases to zero as  $x \rightarrow \infty$ .

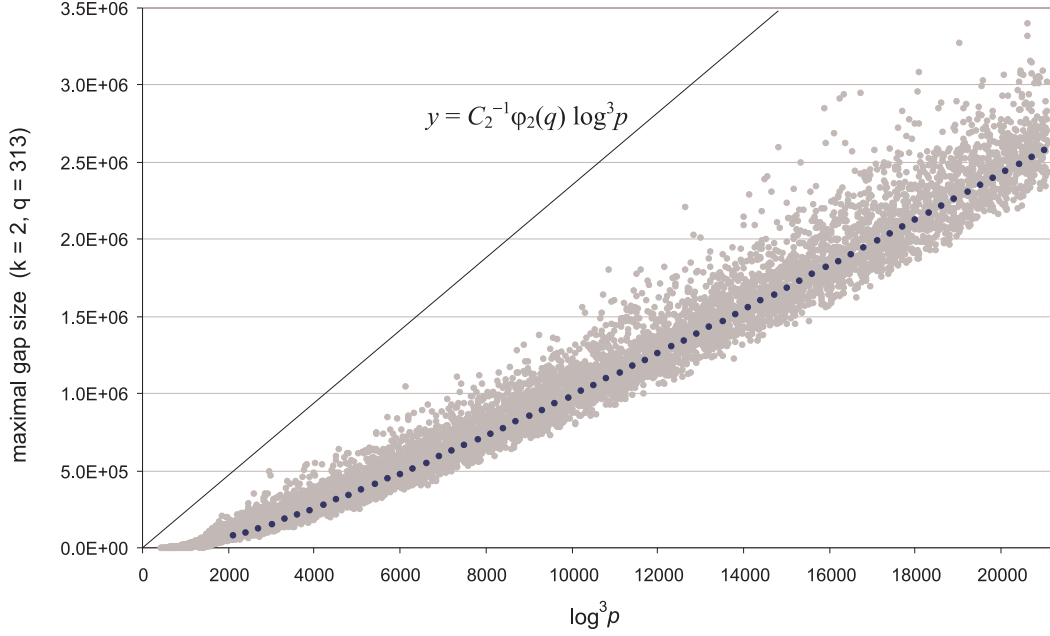


Figure 2: Maximal gaps  $G_c$  between lesser twin primes  $p = r + nq \in \mathbb{P}_c$  below  $x$  for  $q = 313$ ,  $x < 10^{12}$ ,  $k = 2$ . Blue curve: trend  $\tilde{T}_c$  of eq. (12); top line:  $y = C_2^{-1} \varphi_2(q) \log^3 p$ .

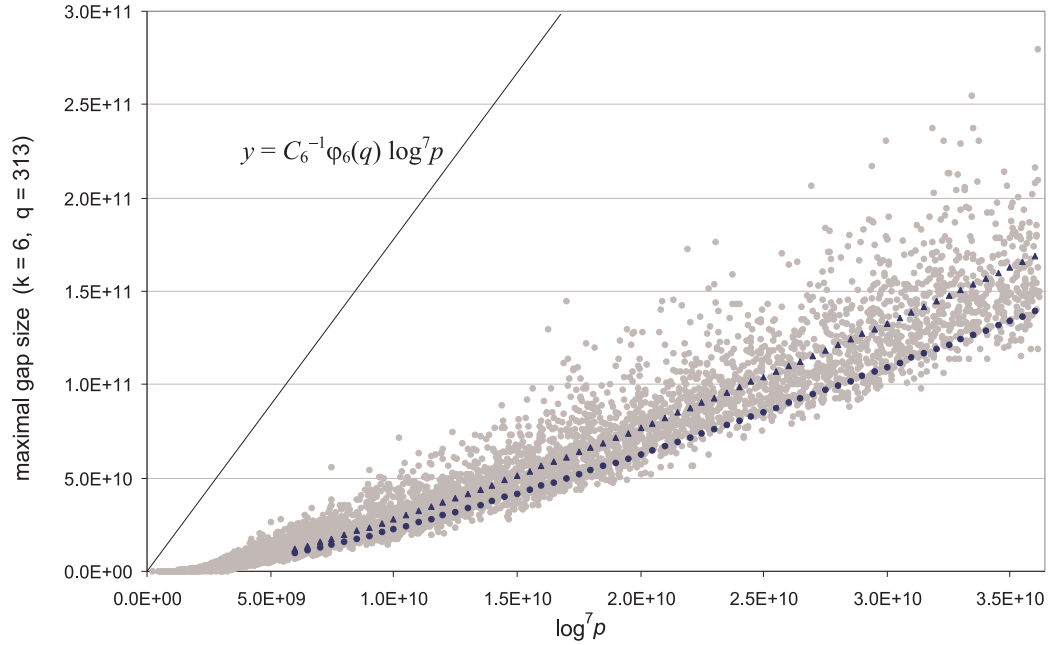


Figure 3: Maximal gaps  $G_c$  between prime sextuplets  $p = r + nq \in \mathbb{P}_c$  below  $x$  for  $q = 313$ ,  $x < 10^{14}$ ,  $k = 6$ . Blue curves: trends  $\tilde{T}_c$  and  $\bar{T}_c$  of (12), (14); top line:  $y = C_6^{-1} \varphi_6(q) \log^7 p$ .

## 3.2 The distribution of maximal gaps

In section 3.1 we have tested equations that determine the growth trend of maximal gaps between primes in sequences  $\mathbb{P}_c$ . How are maximal gap sizes distributed in the neighborhood of their respective trend?

We will perform a rescaling transformation (motivated by extreme value theory): subtract the trend from the actual gap size, and then divide the result by a natural unit, the “local” average gap. This way each maximal gap size is mapped to its rescaled value:

$$\text{maximal gap size } G \mapsto \text{rescaled value} = \frac{G - \text{trend}}{\text{average gap}}.$$

Gaps above the trend curve are mapped to positive rescaled values, while gaps below the trend curve are mapped to negative rescaled values.

**Case  $k = 1$ .** For maximal gaps  $G_{q,r}$  between primes  $p \equiv r \pmod{q}$ , the trend  $T$  is given by eqs. (33), (45), (46). The rescaling operation has the form

$$G_{q,r}(x) \mapsto w = \frac{G_{q,r}(x) - T(q, x)}{a(q, x)}. \quad (47)$$

where  $a(q, x) = \frac{\varphi(q)x}{\text{li } x}$ . Figure 4 shows histograms of rescaled values  $w$  for maximal gaps  $G_{q,r}$  between primes  $p \equiv r \pmod{q}$  for  $q = 16001$ .

**Case  $k \geq 2$ .** For maximal gaps  $G_c$  between prime  $k$ -tuples with  $p = r + nq \in \mathbb{P}_c$ , we can use the trend  $\tilde{T}_c$  of eq. (12). Then the rescaling operation has the form

$$G_c(x) \mapsto \tilde{h} = \frac{G_c(x) - \tilde{T}_c(x)}{\tilde{a}_c(x)}, \quad (48)$$

where  $\tilde{a}_c(x)$  is defined by (6). Figure 5 shows histograms of rescaled values  $\tilde{h}$  for maximal gaps  $G_c$  between lesser twin primes  $p = r + nq \in \mathbb{P}_c$  for  $q = 16001$ ,  $k = 2$ .

In both Figs. 4 and 5, we easily see that the histograms and fitting distributions are skewed to the right, i.e., the right tail is longer and heavier. Among two-parameter distributions, the Gumbel extreme value distribution is a very good fit; cf. [21]. This was true in all our computational experiments. For all histograms shown in Figs. 4 and 5, the Kolmogorov-Smirnov goodness-of-fit statistic is less than 0.01; in fact, for most of the histograms, the goodness-of-fit statistic is about 0.003.

If we look at three-parameter distributions, then an excellent fit is the *Generalized Extreme Value* (GEV) distribution, which includes the Gumbel distribution as a special case. The *shape parameter* in the best-fit GEV distributions is close to zero; note that the Gumbel distribution is a GEV distribution whose shape parameter is exactly zero. So can the Gumbel distribution be the limit law for appropriately rescaled sequences of maximal gaps  $G_{q,r}(p)$  and  $G_c(p)$  as  $p \rightarrow \infty$ ? Does such a limiting distribution exist at all?



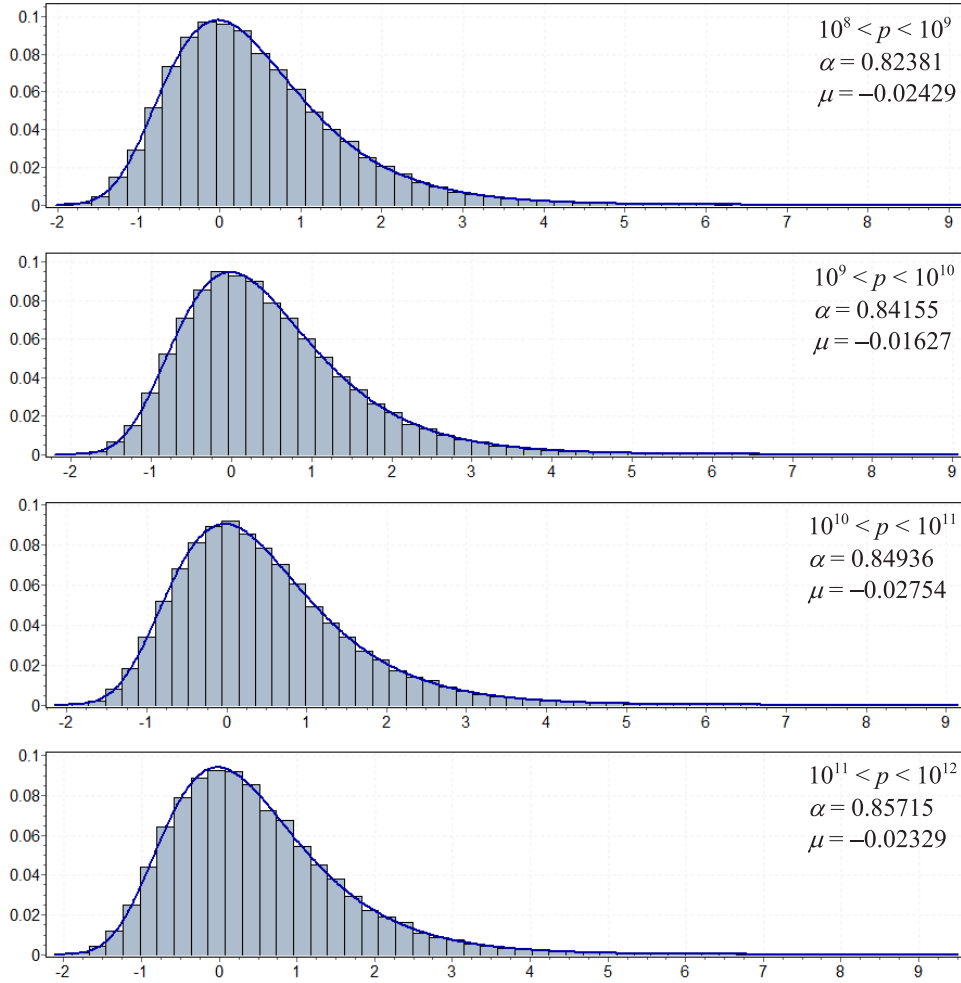


Figure 4: Histograms of  $w$ -values (47) for maximal gaps  $G_{q,r}$  between primes  $p = r + nq$  for  $q = 16001$ ,  $r \in [1, 16000]$ . Curves are best-fit Gumbel distributions (pdfs) with scale  $\alpha$  and mode  $\mu$ .

**The scale parameter  $\alpha$ .** For  $k = 1$ , we observed that the scale parameter of best-fit Gumbel distributions for  $w$ -values (47) was in the range  $\alpha \in [0.7, 1]$ . The parameter  $\alpha$  seems to slowly grow towards 1 as  $p \rightarrow \infty$ ; see Fig. 4. For  $k \geq 2$ , the scale parameter of best-fit Gumbel distributions for  $\tilde{h}$ -values (48) was usually a little over 1; see Fig. 5. However, if instead of (48) we use the (simpler) rescaling transformation

$$G_c(x) \mapsto \bar{h} = \frac{G_c(x) - \bar{T}_c(x)}{\bar{a}_c(x)}, \quad (49)$$

where  $\bar{a}_c$  and  $\bar{T}_c$  are defined, respectively, by (7) and (14), then the resulting Gumbel distributions of  $\bar{h}$ -values will typically have scales  $\alpha$  a little below 1. In a similar experiment with *random* gaps, the scale was also close to 1; see [23, sect. 3.3].

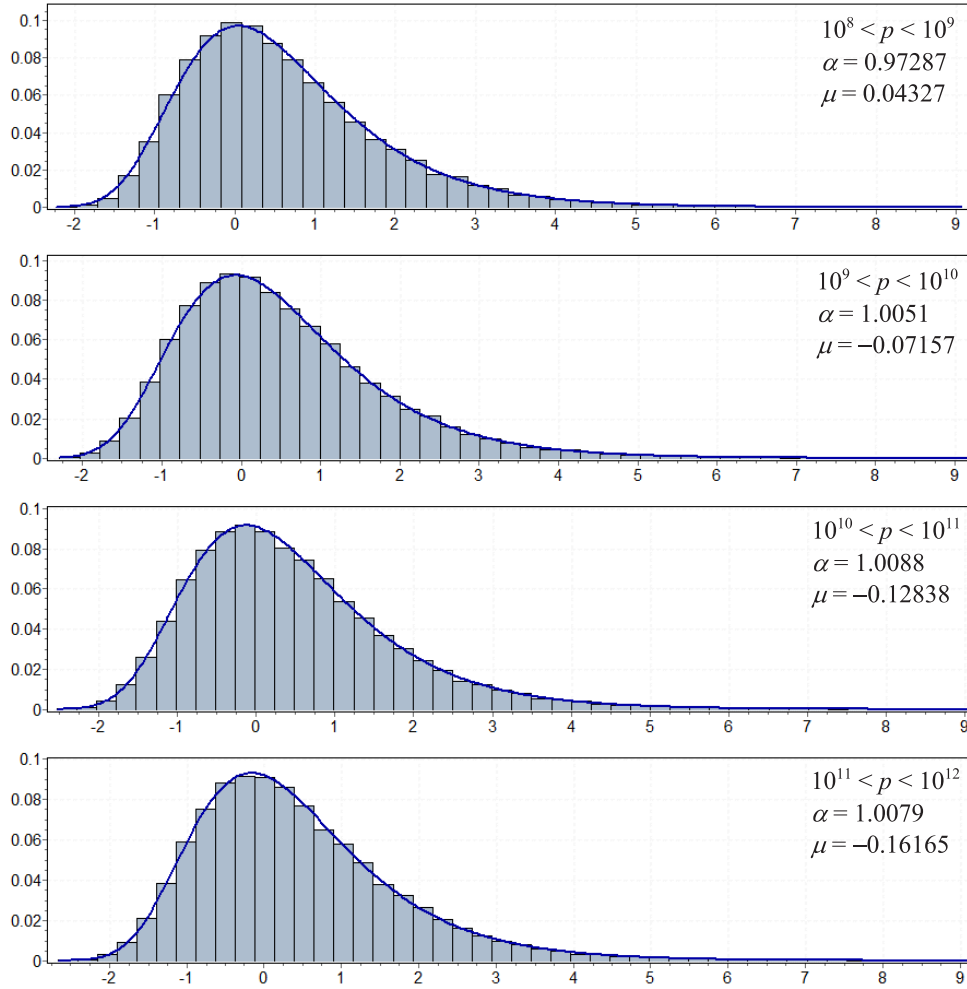


Figure 5: Histograms of  $\tilde{h}$ -values (48) for maximal gaps  $G_c$  between lesser twin primes  $p = r + nq \in \mathbb{P}_c$  for  $q = 16001$  and admissible residue classes  $r \in [1, 16000]$ ,  $r \neq 15999$ . Curves are best-fit Gumbel distributions (pdfs) with scale  $\alpha$  and mode  $\mu$ .

### 3.3 Counting the maximal gaps

We used PARI/GP function `findallgaps` (see source code in *Appendix 5.2*) to determine average numbers of maximal gaps  $G_{q,r}$  between primes  $p = r + nq$ ,  $p \in [x, ex]$ , for  $x = e^j$ ,  $j = 1, 2, \dots, 27$ . Similar statistics were also gathered for gaps  $G_c$ . Figures 6–9 show the results of this computation for  $q = 16001$ ,  $k \leq 4$ . The average number of maximal gaps  $G_c$  for  $p \in [x, ex]$  indeed seems to *very slowly* approach  $k + 1$ , as predicted by (40); see sect.2.4. The graph of  $\text{mean}(N_c(ex) - N_c(x))$  vs.  $\log x$  for gaps  $G_c$  between  $k$ -tuples is closely approximated by a hyperbola with horizontal asymptote  $y = k + 1$ ; see Figs. 6–9.

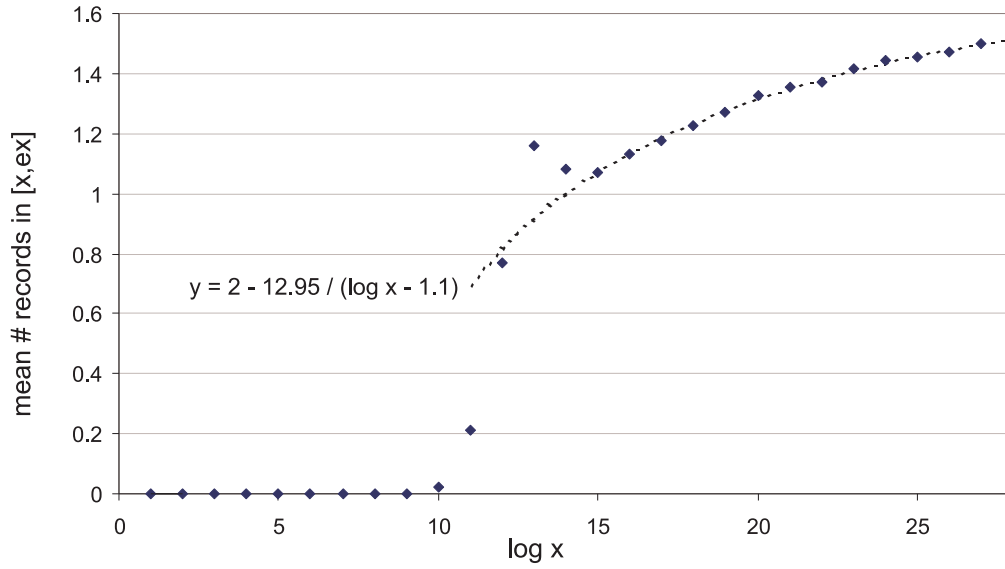


Figure 6: Primes  $p = r + nq$ ,  $k = 1$ ,  $q = 16001$ . Mean number of maximal gaps  $G_{q,r}$  observed for  $p \in [x, ex]$ ,  $x = e^j$ ,  $j \leq 27$ . Averaging for all admissible  $r$ . Dotted curve is a hyperbola with horizontal asymptote  $y = 2$ .

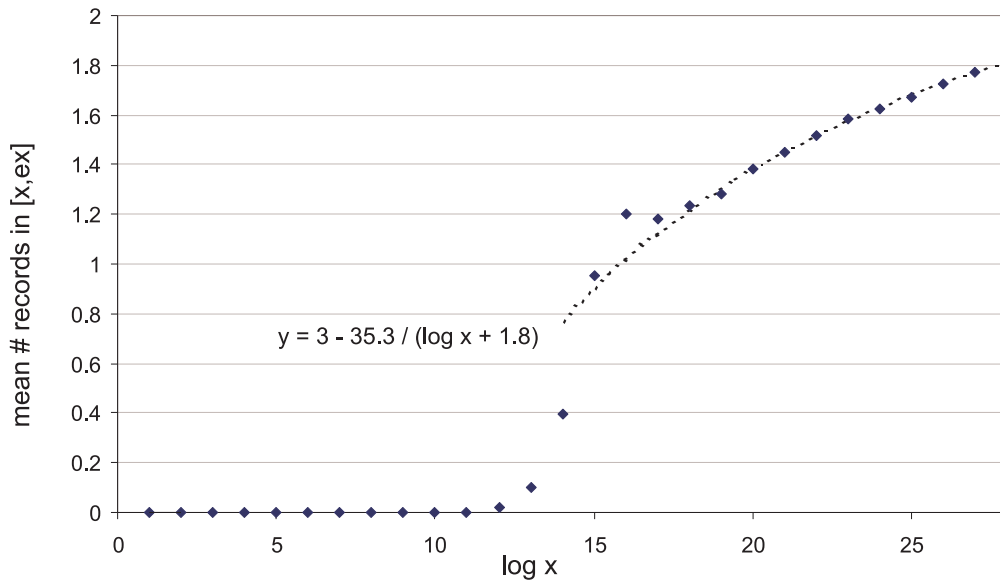


Figure 7: Lesser twin primes  $p = r + nq \in \mathbb{P}_c$ ,  $k = 2$ ,  $q = 16001$ . Mean number of maximal gaps  $G_c$  observed for  $p \in [x, ex]$ ,  $x = e^j$ ,  $j \leq 27$ . Averaging for all admissible  $r$ . Dotted curve is a hyperbola with horizontal asymptote  $y = 3$ .

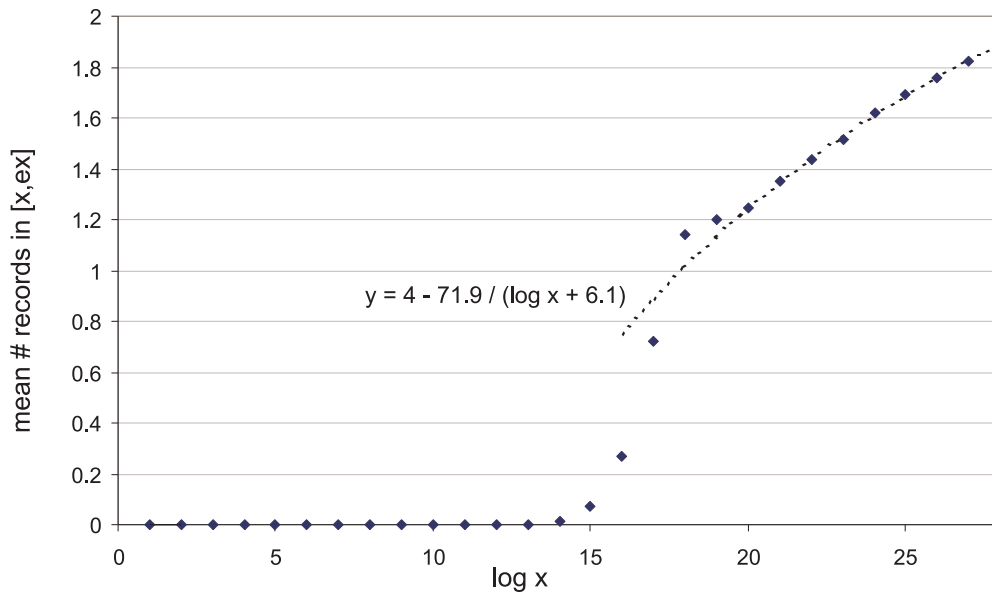


Figure 8: Prime triplets  $(p, p + 2, p + 6)$ ,  $p = r + nq \in \mathbb{P}_c$ ,  $k = 3$ ,  $q = 16001$ . Mean number of maximal gaps  $G_c$  observed for  $p \in [x, ex]$ ,  $x = e^j$ ,  $j \leq 27$ . Averaging for all admissible  $r$ . Dotted curve is a hyperbola with horizontal asymptote  $y = 4$ .

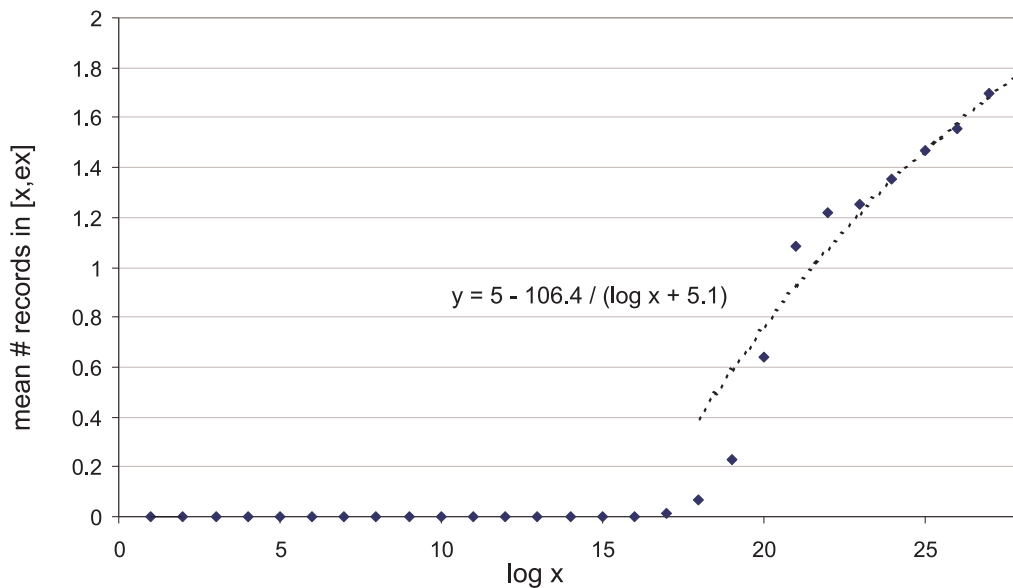


Figure 9: Prime quadruplets  $p = r + nq \in \mathbb{P}_c$ ,  $k = 4$ ,  $q = 16001$ . Mean number of maximal gaps  $G_c$  observed for  $p \in [x, ex]$ ,  $x = e^j$ ,  $j \leq 27$ . Averaging for all admissible  $r$ . Dotted curve is a hyperbola with horizontal asymptote  $y = 5$ .

### 3.4 How long do we wait for the next maximal gap?

Let  $P(n) = \underline{A002386}(n)$  and  $P'(n) = \underline{A000101}(n)$  be the lower and upper endpoints of the  $n$ -th record (maximal) gap  $R(n)$  between primes:  $R(n) = \underline{A005250}(n) = P'(n) - P(n)$ .

Consider the distances  $P(n) - P(n-1)$  from one maximal gap to the next. (In statistics, a similar quantity is sometimes called “inter-record times”.) In Figure 10 we present a plot of these distances; the figure also shows the corresponding plot for *twin primes*. As can be seen from Fig. 10, the quantity  $P(n) - P(n-1)$  grows approximately exponentially with  $n$  (but not monotonically). Indeed, typical inter-record times are expected to satisfy<sup>4</sup>

$$\log(P(n) - P(n-1)) < \log P(n) \sim \frac{n}{2} \quad \text{as } n \rightarrow \infty. \quad (50)$$

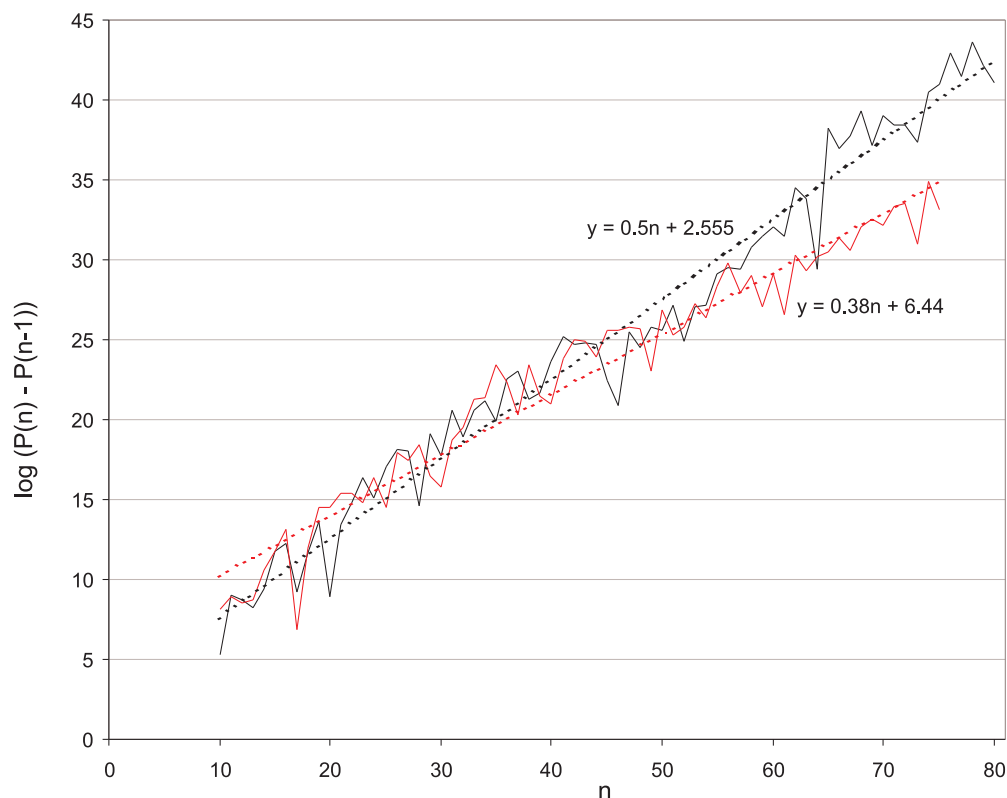


Figure 10: Inter-record times  $P(n) - P(n-1)$  for gaps between primes (black) and a similar quantity  $P_c(n) - P_c(n-1)$  for gaps between twin primes (red). Lines are exponential fits. Values for  $n < 10$  are skipped.

<sup>4</sup>The asymptotic equivalence  $\sim$  in eqs. (50), (51) is a restatement of eqs. (40), (41). It would be logically unsound to suppose that  $\log(P(n) - P(n-1)) \stackrel{?}{\sim} \log P(n)$  because we cannot exclude the possibility that  $\log(P(n) - P(n-1))$  might (very rarely) become as small as  $\log G(x) \approx 2 \log \log x$ , where  $x = P(n)$ .

More generally, let  $P_c(n)$  and  $P'_c(n)$  be the endpoints of the  $n$ -th maximal gap  $R_c(n)$  between primes in sequence  $\mathbb{P}_c$ , where each prime is  $r \pmod{q}$  and starts an admissible prime  $k$ -tuple. Then, in accordance with heuristic reasoning of sect. 2.4, for typical inter-record times  $P_c(n) - P_c(n-1)$  separating the maximal gaps  $R_c(n-1)$  and  $R_c(n)$  we expect to see

$$\log(P_c(n) - P_c(n-1)) < \log P_c(n) \sim \frac{n}{k+1} \quad \text{as } n \rightarrow \infty. \quad (51)$$

In the special case  $k = 2$ , that is, for maximal gaps between *twin primes*, the right-hand side of (51) is expected to be  $\frac{n}{3}$  for large  $n$  (whereas Fig. 10 suggests the right-hand side  $0.38n$  based on a very limited data set for  $10 \leq n \leq 75$ ). As we have seen in sect. 3.3, the average number of maximal gaps between  $k$ -tuples occurring for primes  $p \in [x, ex]$  slowly approaches  $k + 1$  *from below*. For moderate values of  $x$  attainable in computation, this average is typically between 1 and  $k + 1$ . Accordingly, we see that the right-hand side of (51) yields a prediction  $\asymp e^{n/(k+1)}$  that *underestimates* the typical inter-record times and the primes  $P_c(n)$ . Computations may yield estimates  $P_c(n) - P_c(n-1) < P_c(n) \approx Ce^{\beta n}$ , where  $\beta \in [\frac{1}{k+1}, 1]$ , depending on the range of available data.

*Remarks.* (i) Sample graphs of  $\log P_c(n)$  vs.  $n$  can be plotted online at the OEIS website: click *graph* and scroll to the logarithmic plot for sequences [A002386](#) ( $k = 1$ ), [A113275](#) ( $k = 2$ ), [A201597](#) ( $k = 3$ ), [A201599](#) ( $k = 3$ ), [A229907](#) ( $k = 4$ ), [A201063](#) ( $k = 5$ ), [A201074](#) ( $k = 5$ ), [A200504](#) ( $k = 6$ ). In all these graphs, when  $n$  is large enough,  $\log P_c(n)$  seems to grow approximately linearly with  $n$ . We conjecture that the slope of such a linear approximation slowly decreases, approaching the slope value  $1/(k+1)$  as  $n \rightarrow \infty$ .

(ii) Recall that for the maximal prime gaps  $G(x)$  Shanks [33] conjectured the asymptotic equality  $G(x) \sim \log^2 x$ , a strengthened form of Cramér's conjecture. This seems to suggest that (unusually large) maximal gaps  $g$  may in fact occur as early as at  $x \asymp e^{\sqrt{g}}$ . On the other hand, Wolf [36, 38] conjectured that typically a gap of size  $d$  appears for the first time between primes near  $\sqrt{d} \cdot e^{\sqrt{d}}$ . Combining these observations, we may further observe that exceptionally large maximal gaps

$$\text{exceptionally large gaps } g = G(x) > \log^2 x \quad (52)$$

are also those which appear for the first time unusually early. Namely, they occur at  $x$  roughly by a factor of  $\sqrt{d}$  earlier than the typical first occurrence of a gap  $d$  at  $x \asymp \sqrt{d} \cdot e^{\sqrt{d}}$ . Note that Granville [15, p. 24] suggests that gaps of unusually large size (52) occur infinitely often — and we will even see infinitely many of those exceeding  $1.1229 \log^2 x$ . In contrast, Sun [35, Conj. 2.3] made a conjecture implying that exceptions like (52) occur only finitely often, while Firoozbakht's conjecture implies that exceptions (52) never occur for primes  $p \geq 11$ ; see [22]. Here we cautiously predict that exceptional gaps of size (52) are only a zero proportion of maximal gaps. This can be viewed as restatement of the generalized Cramér conjectures (20), (34) for the special case  $k = 1$ ,  $q = 2$  (cf. *Appendix 5.4*).

## 4 Summary

We have extensively studied record (maximal) gaps between prime  $k$ -tuples in residue classes  $(\text{mod } q)$ . Our computational experiments described in section 3 took months of computer time. Numerical evidence allows us to arrive at the following conclusions, which are also supported by heuristic reasoning.

- For  $k = 1$ , the observed growth trend of maximal gaps  $G_{q,r}(x)$  is given by (33), (45). In particular, for maximal prime gaps ( $k = 1, q = 2$ ) the trend equation reduces to

$$G_{2,1}(x) \sim \log^2 x - 2 \log x \log \log x + O(\log x) \quad \text{as } x \rightarrow \infty.$$

- For  $k \geq 2$ , a significant proportion of maximal gaps  $G_c(x)$  are observed between the trend curves of eqs.(12) and (14), which can be heuristically derived from extreme value theory.
- The Gumbel distribution, after proper rescaling, is a possible limit law for  $G_{q,r}(p)$  as well as  $G_c(p)$ . The existence of such a limiting distribution is an open question.
- Almost all maximal gaps  $G_{q,r}(p)$  between primes in residue classes  $\text{mod } q$  seem to satisfy appropriate generalizations of the Cramér and Shanks conjectures (34) and (35):

$$G_{q,r}(p) \lesssim \varphi(q) \log^2 p.$$

- Similar generalizations (20) and (21) of the Cramér and Shanks conjectures are apparently true for almost all maximal gaps  $G_c(p)$  between primes in  $\mathbb{P}_c$ :

$$G_c(p) \lesssim C_k^{-1} \varphi_k(q) \log^{k+1} p.$$

- Exceptionally large gaps  $G_{q,r}(p) > \varphi(q) \log^2 p$  are extremely rare (see *Appendix 5.4*). We conjecture that only a zero proportion of maximal gaps are such exceptions. A similar observation holds for  $G_c(p)$  violating (20).
- We conjecture that the total number  $N_{q,r}(x)$  of maximal gaps  $G_{q,r}$  observed up to  $x$  is below  $C \log x$  for some  $C > 2$ .
- More generally, the number  $N_c(x)$  of maximal gaps between primes in  $\mathbb{P}_c$  up to  $x$  satisfies the inequality  $N_c(x) < C \log x$  for some  $C > k + 1$ , where  $k$  is the number of primes in the  $k$ -tuple pattern defining the sequence  $\mathbb{P}_c$ .

## 5 Appendix: Details of computational experiments

Interested readers can reproduce and extend our results using the programs below.

## 5.1 PARI/GP program maxgap.gp (ver. 2.1)

```
default(realprecision,11)
outpath = "c:\\wgap"

\\ maxgap(q,r,end [,b0,b1,b2]) ver 2.1 computes maximal gaps g
\\ between primes  $p = qn + r$ , as well as rescaled values (w, u, h):
\\ w - as in arXiv:1610.03340 eqs.(1)-(5)
\\ u - same as w, but with constant  $b = \ln \phi(q)$ ;
\\ h - based on extreme value theory (cf. randomgap.gp in arXiv:1610.03340)
\\ Results are written on screen and in the folder specified by outpath string.
\\ Computation ends when primes exceed the end parameter.
maxgap(q,r,end,b0=1,b1=4,b2=2.7) = {
  re = 0;
  p = pmin(q,r);
  t = eulerphi(q);
  inc = q;
  while(p<end,
    m = p + re;
    p = m + inc;
    while(!isprime(p), p+=inc);
    while(!isprime(m), m-=inc);
    g = p - m;
    if(g>re,
      re=g; Lip=li(p); a=t*p/Lip; Logp=log(p);
      h = g/a-log(Lip/t);
      u = g/a-2*log(Lip/t)+Logp-log(t);
      w = g/a-2*log(Lip/t)+Logp-log(t)*(b0+b1/max(2,log(Logp))^b2);

      f = ceil(Logp/log(10));
      write(outpath"\\\"q\"_1e\"f\".txt",
        w" "u" "h" "g" "m" "p" q="q" r="r");
      print(w" "u" "h" "g" "m" "p" q="q" r="r");
      if(g/t>log(p)^2, write(outpath"\\\"q\"_1e\"f\".txt","extra large"));
      if(g%2==0, inc=lcm(2,q));

      \\ optional part: statistics for p in intervals [x/e,ex] for  $x=e^j$ 
      i = ceil(Logp);
      j = floor(Logp);
      if(N!='N,N[j]++); \\ count maxima with p in [x,ex] for  $x=e^j$ 
      write(outpath"\\\"q\"_exp\"i\".txt", w" "u" "h" "g" "m" "p" q="q" r="r");
      write(outpath"\\\"q\"_exp\"j\".txt", w" "u" "h" "g" "m" "p" q="q" r="r");
    )
  )
}
```



## 5.2 PARI/GP: Auxiliary functions for maxgap.gp

```
\\ These functions are intended for use with the program maxgap.gp
\\ It is best to include them in the same file with maxgap.gp

\\ li(x) computes the logarithmic integral of x
li(x) = real(-eint1(-log(x)))

\\ pmin(q,r) computes the least prime p = qn + r, for n=0,1,2,3,...
pmin(q,r) = forstep(p=r,1e99,q, if(isprime(p), return(p)))

\\ findallgaps(q,end): Given q, call maxgap(q,r,end) for all r coprime to q.
\\ Output total and average counts of maximal gaps in intervals [x,ex].
findallgaps(q,end) = {
  t = eulerphi(q);
  N = vector(99,j,0);
  for(r=1,q, if(gcd(q,r)==1,maxgap(q,r,end)));
  nmax = floor(log(end));
  for (n=1,nmax,
    avg = 1.0*N[n]/t;
    write(outpath\\"q"stats.txt", n" "avg" "N[n]);
  )
}
```

## 5.3 Notes on distribution fitting

To study distributions of rescaled maximal gaps, we used the distribution-fitting software EasyFit [26]. Data files created with maxgap.gp are easily imported into EasyFit:

1. From the *File* menu, choose *Open*.
2. Select the data file.
3. Specify *Field Delimiter = space*.
4. Click *Update*, then *OK*.

**Caution:** PARI/GP outputs large and small real numbers in a mantissa-exponent format *with a space* preceding the exponent (e.g. 1.7874829515 E-5), whereas EasyFit expects such numbers *without a space* (e.g. 1.7874829515E-5). Therefore, before importing into EasyFit, search the data files for " E" and replace all occurrences with "E".

## 5.4 Exceptionally large gaps: $G_{q,r}(p) > \varphi(q) \log^2 p$

Table 1. Large maximal gaps:  $G_{q,r}(p) > \varphi(q) \log^2 p$  for  $p < 10^9$ ,  $q \leq 25000$

Gap	$G_{q,r}(p)$	Start of gap	End of gap ( $p$ )	$q$	$r$	$G_{q,r}(p)/(\varphi(q) \log^2 p)$	
(i)	208650	3415781	3624431	1605	341	1.0786589153	
	316790	726611	1043401	2005	801	1.0309808771	
	229350	1409633	1638983	2085	173	1.0145547849	
	532602	355339	887941	4227	271	1.0081862161	
	984170	5357381	6341551	4279	73	1.0339720553	
	1263426	10176791	11440217	4897	825	1.0056800570	
	2306938	82541821	84848759	6907	3171	1.0022590147	
	3415794	376981823	380397617	8497	3921	1.0703375544	
	2266530	198565889	200832419	8785	7319	1.0335372951	
	7326222	222677837	230004059	20017	8729	1.0166221904	
	6336090	10862323	17198413	23467	20569	1.0064940453	
	7230930	130172279	137403209	24595	15539	1.0468373915	
	(ii)	411480	470669167	471080647	3048	55	1.0235488825
		208650	3415781	3624431	3210	341	1.0786589153
316790		726611	1043401	4010	801	1.0309808771	
229350		1409633	1638983	4170	173	1.0145547849	
657504		896016139	896673643	4566	2563	1.0179389550	
1530912		728869417	730400329	6896	3593	1.0684247390	
532602		355339	887941	8454	271	1.0081862161	
984170		5357381	6341551	8558	73	1.0339720553	
1263426		10176791	11440217	9794	825	1.0056800570	
2119706		665152001	667271707	10046	6341	1.0223668231	
1885228		163504573	165389801	10532	5805	1.0000704209	
1594416		145465687	147060103	13512	9007	1.0026889378	
2306938		82541821	84848759	13814	3171	1.0022590147	
3108778		524646211	527754989	15622	12585	1.0098218219	
1896608		164663	2061271	16934	12257	1.0598397341	
3415794		376981823	380397617	16994	3921	1.0703375544	
2266530		198565889	200832419	17570	7319	1.0335372951	
2937868		71725099	74662967	17698	12803	1.0103309882	
2823288		37906669	40729957	18098	9457	1.0162761199	
2453760	11626561	14080321	18176	12097	1.0107626289		
3906628	190071823	193978451	18692	11567	1.1480589845		
(iii)	657504	896016139	896673643	2283	280	1.0179389550	
	2119706	665152001	667271707	5023	1318	1.0223668231	
	3108778	524646211	527754989	7811	4774	1.0098218219	
	1896608	164663	2061271	8467	3790	1.0598397341	
	2937868	71725099	74662967	8849	3954	1.0103309882	
	2823288	37906669	40729957	9049	408	1.0162761199	
	3422630	735473	4158103	14881	6304	1.0368176014	
	3758772	144803717	148562489	15927	11360	1.0000152764	
	3002682	8462609	11465291	16869	11240	1.0107025944	
	8083028	344107541	352190569	19619	9900	1.1134625422	
	4575906	20250677	24826583	22653	21548	1.0463153374	

The above table lists exceptionally large maximal gaps  $G_{q,r}(p) > \varphi(q) \log^2 p$ . No other maximal gaps with this property were found for  $p < 10^9$ ,  $q \leq 25000$ . Three sections of the table correspond to (i) odd  $q, r$ ; (ii) even  $q$ ; (iii) even  $r$ . (Overlap between sections is due to the fact that  $\varphi(q) = \varphi(2q)$  for odd  $q$ .) No such large gaps exist for  $p < 10^{10}$ ,  $q \leq 1000$ .

## 5.5 The Hardy–Littlewood constants $C_k$

The Hardy–Littlewood  $k$ -tuple conjecture [17] allows one to predict the average frequencies of prime  $k$ -tuples near  $p$ , as well as the approximate total counts of prime  $k$ -tuples below  $x$ . Specifically, the Hardy–Littlewood  $k$ -tuple constants  $C_k$ , divided by  $\log^k p$ , give us an estimate of the average frequency of prime  $k$ -tuples near  $p$ :

$$\text{Frequency of } k\text{-tuples} \sim \frac{C_k}{\log^k p}.$$

Accordingly, for  $\pi_k(x)$ , the total count of  $k$ -tuples below  $x$ , we have

$$\pi_k(x) \sim C_k \int_2^x \frac{dt}{\log^k t} = C_k \text{Li}_k(x).$$

The Hardy–Littlewood constants  $C_k$  can be defined in terms of infinite products over primes. In particular, for densest admissible prime  $k$ -tuples with  $k \leq 7$  we have:

$$\begin{aligned} C_1 &= 1 \quad (\text{by convention, in accordance with the prime number theorem}); \\ C_2 &= 2 \prod_{p>2} \frac{p(p-2)}{(p-1)^2} \approx 1.32032363169373914785562422 \quad (\underline{\text{A005597}}, \underline{\text{A114907}}); \\ C_3 &= \frac{9}{2} \prod_{p>3} \frac{p^2(p-3)}{(p-1)^3} \approx 2.85824859571922043243013466 \quad (\underline{\text{A065418}}); \\ C_4 &= \frac{27}{2} \prod_{p>4} \frac{p^3(p-4)}{(p-1)^4} \approx 4.15118086323741575716528556 \quad (\underline{\text{A065419}}); \\ C_5 &= \frac{15^4}{2^{11}} \prod_{p>5} \frac{p^4(p-5)}{(p-1)^5} \approx 10.131794949996079843988427 \quad (\underline{\text{A269843}}); \\ C_6 &= \frac{15^5}{2^{13}} \prod_{p>6} \frac{p^5(p-6)}{(p-1)^6} \approx 17.2986123115848886061221077 \quad (\underline{\text{A269846}}); \\ C_7 &= \frac{35^6}{3 \cdot 2^{22}} \prod_{p>7} \frac{p^6(p-7)}{(p-1)^7} \approx 53.9719483001296523960730291 \quad (\underline{\text{A271742}}). \end{aligned}$$

Forbes [10] gives values of the Hardy–Littlewood constants  $C_k$  up to  $k = 24$ , albeit with fewer significant digits; see also [8, p. 86]. Starting from  $k = 8$ , we may often encounter more than one numerical value of  $C_k$  for a single  $k$ . (If there are  $m$  different patterns of densest admissible prime  $k$ -tuples for the same  $k$ , then we have  $\lceil \frac{m}{2} \rceil$  different numerical values of  $C_k$ , depending on the actual pattern of the  $k$ -tuple; see [10].)

## 5.6 Integrals $\text{Li}_k(x)$

Let  $k \in \mathbb{N}$  and  $x > 1$ , and let

$$\begin{aligned} F_k(x) &= \int \frac{dx}{\log^k x} && \text{(indefinite integral);} \\ \text{Li}_k(x) &= \int_2^x \frac{dt}{\log^k t} && \text{(definite integral).} \end{aligned}$$

Denote by  $\text{li } x$  the conventional logarithmic integral (principal value):

$$\text{li } x = \int_0^x \frac{dt}{\log t} = \int_2^x \frac{dt}{\log t} + 1.04516\dots$$

In PARI/GP, an easy way to compute  $\text{li } x$  is as follows:  $\text{li}(x) = \text{real}(-\text{eint1}(-\log(x)))$ .

The integrals  $F_k(x)$  and  $\text{Li}_k(x) = F_k(x) - F_k(2)$  can also be expressed in terms of  $\text{li } x$ . Integration by parts gives

$$\int \frac{dx}{\log x} = \frac{x}{\log x} + \frac{x}{\log^2 x} + \frac{2x}{\log^3 x} + \frac{6x}{\log^4 x} + \dots + \frac{(k-2)!x}{\log^{k-1} x} + (k-1)! \int \frac{dx}{\log^k x}. \quad (53)$$

Therefore,

$$\begin{aligned} F_2(x) &= \frac{1}{1!} \left( \text{li } x - \frac{x}{\log x} \right) + C, \\ F_3(x) &= \frac{1}{2!} \left( \text{li } x - \frac{x}{\log^2 x} (\log x + 1) \right) + C, \\ F_4(x) &= \frac{1}{3!} \left( \text{li } x - \frac{x}{\log^3 x} (\log^2 x + \log x + 2) \right) + C, \\ F_5(x) &= \frac{1}{4!} \left( \text{li } x - \frac{x}{\log^4 x} (\log^3 x + \log^2 x + 2 \log x + 6) \right) + C, \\ F_6(x) &= \frac{1}{5!} \left( \text{li } x - \frac{x}{\log^5 x} (\log^4 x + \log^3 x + 2 \log^2 x + 6 \log x + 24) \right) + C, \end{aligned}$$

and, in general,

$$F_{k+1}(x) = \frac{1}{k!} \left( \text{li } x - \frac{x}{\log^k x} \sum_{j=1}^k (k-j)! \log^{j-1} x \right) + C. \quad (54)$$

Using these formulas we can compute  $\text{Li}_k(x)$  for approximating  $\pi_c(x)$  (the prime counting function for sequence  $\mathbb{P}_c$ ) in accordance with the  $k$ -tuple equidistribution conjecture (5):

$$\pi_c(x) \approx \frac{C_k}{\varphi_k(q)} \text{Li}_k(x) = \frac{C_k}{\varphi_k(q)} (F_k(x) - F_k(2)).$$

The values of  $\text{li } x$ , and hence  $\text{Li}_k(x)$ , can be calculated without (numerical) integration. For example, one can use the following rapidly converging series for  $\text{li } x$ , with  $n!$  in the denominator and  $\log^n x$  in the numerator (see [31, formulas 1.6.1.8-9]):

$$\text{li } x = \log \log x + \sum_{n=1}^{\infty} \frac{\log^n x}{n \cdot n!} \quad \text{for } x > 1. \quad (55)$$

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