

THE PEAK AND DESCENT STATISTICS OVER BALLOT PERMUTATIONS

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ABSTRACT. A ballot permutation is a permutation π such that in any prefix of π the descent number is not more than the ascent number. By using a reversal-concatenation map, we give a formula for the joint distribution (pk, des) of the peak and descent statistics over ballot permutations, and connect this distribution and the joint distribution (pk, dp, des) of the peak, depth, and descent statistics over ordinary permutations in terms of generating functions. As corollaries, we obtain several formulas for the bivariate generating function for (i) the peak statistic over ballot permutations, (ii) the descent statistic over ballot permutations, and (iii) the depth statistic over ordinary permutations. In particular, we confirm Spiros conjecture which finds the equidistribution of the descent statistic for ballot permutations and an analogue of the descent statistic for odd order permutations.

1. INTRODUCTION

In 1887 Bertrand [5] introduced the *ballot problem*: Consider an election for two candidates A and B with a total of n votes, where A wins a votes and B wins $n - a = b$ votes. What is the probability that at each count A is always ahead? Equivalently, what is the probability of a lattice path from the origin to the point $(b, a - 1)$ is a ballot path? Here a *ballot path* is a lattice path that never goes below the line $y = x$, see [5, 9, 22, 23, 33]. This is one of the beginnings of lattice path enumeration and early problems in probabilistic combinatorics, see Humphreys [18]. Among various ways of solving the ballot problem, there are the well known reflection principle [18] and the cycle lemma [10]. The answer to the ballot problem is the *ballot number*

$$\frac{a - b}{a + b} \binom{a + b}{b},$$

which reduces to the Catalan number $\frac{1}{n+1} \binom{2n}{n}$ when $a = n + 1$ and $b = n$, see [1, 17]. The ballot problem was generalized by Barbier [3] which demands A maintains as more than k times many votes as B , see also Renault [25, 31].

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A *ballot permutation* is a permutation in any prefix of which the number of ascents is at least the number of descents. Every ballot permutation on n distinct integers naturally corresponds to a ballot path $p_1 p_2 \cdots p_n$ such that $p_i p_{i+1}$ is the unit north step $(0, 1)$ if i is an ascent, and the unit east step $(1, 0)$ otherwise.

The problem of enumerating ballot permutations is closely related with that of enumerating ordinary permutations with a given updown signature, see [2, 8, 24, 27]. Bernardi, Duplantier and Nadeau [4] proved that the number of ballot permutations of length n equals the number of odd order permutations of length n , by using compositions of bijections, and thus Theorem 1.1 follows. A short proof is given by the first author and Zhang [32].

Theorem 1.1 (Bernardi, Duplantier and Nadeau). *The number of ballot permutations of length n is*

$$b_n = \begin{cases} [(n-1)!!]^2, & \text{if } n \text{ is even,} \\ n!(n-2)!!, & \text{if } n \text{ is odd,} \end{cases}$$

where $(-1)!! = 1$.

The sequence $\{b_n\}_{n \geq 0}$ can be found in OEIS [28, A000246], of which the exponential generating function is

$$(1.1) \quad \sum_{n \geq 0} b_n \frac{x^n}{n!} = \sqrt{\frac{1+x}{1-x}}.$$

A ballot permutation whose corresponding ballot path ends on the line $y = x$ is said to be a *Dyck permutation*, whose enumeration is the Eulerian-Catalan number, see Bidkhori and Sullivant [6].

Spiro [29] introduced a statistic $M(\pi)$ for odd order permutations, and conjectured that the number of ballot permutations of length n with d descents equals the number of odd order permutations π of length n such that $M(\pi) = d$. In this paper, we confirm the conjecture by computing their bivariate generating functions in terms of the Eulerian numbers, respectively.

Theorem 1.2. *Let $n \geq 1$ and $0 \leq d \leq \lfloor (n-1)/2 \rfloor$. The number of ballot permutations of length n with d descents equals the number of odd order permutations π of length n with $M(\pi) = d$.*

The first author and Zhang [32] refined Spiros conjecture by tracking the neighbors of the largest letter in these permutations, which is still open. They defined a word u as a *factor* of a word w if there exist words x and y such that $w = xuy$, and a word u as a *cyclic factor* of a permutation $\pi \in \mathcal{S}_n$ if u is a factor of some word v such that (v) is a cycle of π . The conjecture is as follows.

Conjecture 1.3 (Wang and Zhang). *For all n , d , and $2 \leq j \leq n-1$, we have $b_{n,d}(1, j) + b_{n,d}(j, 1) = 2p_{n,d}(1, j)$, where $b_{n,d}(i, j)$ is the number of ballot permutations of length n with d descents which have inj as a factor, and $p_{n,d}(i, j)$ is the number of odd order permutations of length n with $M(\pi) = d$ which have inj as a cyclic factor.*

Manes, Sapounakis, Tasoulas and Tsikouras [21] introduced the concept of depth for a lattice path defined to be the difference between the height of the lowest position of the path and that of the starting point. In this paper, we also consider the depth statistic of a ballot permutation π defined to be the depth of the ballot path corresponding to π . Under this notion, a permutation is a ballot one if and only if its depth is zero.

Zhuang [37] studied the generating function $P^{(\text{pk}, \text{des})}$ of the peak number and descent number over ordinary permutations using noncommutative symmetric functions. By using a map which we called *reversal-concatenation*, and an operator tool, we find a relation between $P^{(\text{pk}, \text{des})}$ and the generating function $B^{(\text{pk}, \text{des})}$ of the same statistics over ballot permutations, see Theorem 3.2. The reversal-concatenation enables us to deal with the relations between the joint distribution $(\text{pk}, \text{dp}, \text{des})$ over ordinary permutations and the joint distribution (pk, des) over ballot permutations in a uniform manner which derive several corollaries, see Section 3.

The discovery of the map was inspired by Gessel's combinatorial interpretation of a decomposition of formal Laurent series in terms of lattice paths [16], Bernardi et al.'s path decompositions [4], and the ω -decomposition [32]. Earlier decompositions were described by Feller [13, page 383] and Foata and Schützenberger [15]. The map is also related to the lowest points of the paths, which was used in the studies of the Chung-Feller theorem, see Woan [35], Shapiro [26] and Eu, Fu and Yeh [12]. Basic generating functions calculating is also used throughout the paper, see Wilf [34] and Flajolet and Sedgewick [14] for general techniques of generating functions.

The main results of this paper, besides Theorem 1.2, are Theorems 3.4 and 3.5. In Theorem 3.4 we express the generating function $B^{(\text{pk}, \text{des})}(x, y, t)$ as the image of a rational function under an operator. In Theorem 3.5 we provide a relation between the generating functions $B^{(\text{pk}, \text{des})}(x, y, t)$ and $P^{(\text{pk}, \text{dp}, \text{des})}(x, y, z, t)$. As corollaries, we obtain the bivariate generating functions for peak number over ballot permutations, and that for descent number over ballot permutations, see Theorems 3.6 and 3.7.

The next section consists of necessary notion and notation. In Section 3 we demonstrate the reversal-concatenation map and its consequences. In Section 4, we use Theorem 3.7 to establish Theorem 1.2.

2. PRELIMINARY

Let \mathcal{S}_n be the permutation group on the set $[n] = \{1, 2, \dots, n\}$. A position $1 \leq i \leq n - 1$ in a permutation $\pi = \pi_1\pi_2 \cdots \pi_n \in \mathcal{S}_n$ is a *descent* if $\pi_i > \pi_{i+1}$, and an *ascent* if $\pi_i < \pi_{i+1}$. Denote the number of descents of π by $\text{des}(\pi)$, and the number of ascents by $\text{asc}(\pi)$. We call the number

$$h(\pi) = \text{asc}(\pi) - \text{des}(\pi)$$

the *height* of π . The permutation π is said to be a *ballot permutation* if the height of any prefix of π is nonnegative, namely, $h(\pi_1\pi_2 \cdots \pi_i) \geq 0$ for all $i \in [n]$. Let \mathcal{B}_n denote the set of ballot permutations on $[n]$. Define $\mathcal{B}_0 = \{\epsilon\}$, where ϵ is the empty

permutation. The number of ballot permutations of height 0, which are also called the *Dyck permutations*, is the Eulerian-Catalan number. We define

$$\text{dp}(\pi) := -\min\{h(\pi_1\pi_2\cdots\pi_i) : 1 \leq i \leq n\},$$

and $\text{dp}(\epsilon) = 0$. It is clear that

$$0 \leq \text{dp}(\pi) \leq \text{des}(\pi).$$

A *lowest position* of π is a position $1 \leq i \leq n$ such that $h(\pi_1\pi_2\cdots\pi_i) = -\text{dp}(\pi)$. Denote by $L(\pi)$ the set of lowest positions of π . For example,

$$\text{dp}(5641327) = 1 \quad \text{and} \quad L(5641327) = \{4, 6\}.$$

From the definition, we see that

$$\pi \in \mathcal{B}_n \iff 1 \in L(\pi).$$

Let \mathcal{O}_n be the set of *odd order permutations* of $[n]$, viz., the set of permutations of $[n]$ which are the products of cycles with odd lengths. In order to define an analogue for the descent statistic in the context of odd order permutations, Spiro [29] defines for a permutation π that

$$M(\pi) = \sum_c \min(\text{cdes}(c), \text{casc}(c)),$$

where the sum runs over all cycles $c = (c_1c_2\cdots c_k)$ of π , with the *cyclic descent*

$$\text{cdes}(c) = |\{i \in [k] : c_i > c_{i+1} \text{ where } c_{k+1} = c_1\}|,$$

and the *cyclic ascent*

$$\text{casc}(c) = |\{i \in [k] : c_i < c_{i+1} \text{ where } c_{k+1} = c_1\}| = |c| - \text{cdes}(c),$$

where $|c|$ is the length of c .

For $n \geq 1$ and $0 \leq d \leq n-1$, the *Eulerian number*, denoted as $E(n, d)$ or $\langle n \rangle_d$, is the number of permutations of $[n]$ with d descents, namely

$$E(n, d) = |\{\pi \in \mathcal{S}_n : \text{des}(\pi) = d\}|,$$

see OEIS [28, A008292]. We adopt the convention $E(0, 0) = 1$ and

$$E(n, d) = 0, \quad \text{if } n < 0, \text{ or } d < 0, \text{ or } d = n \geq 1, \text{ or } d > n.$$

As will be seen, this extension helps dealing with summation calculation by simplifying the domain of indices in summations, so that one may focus on the summands. For instance, the notation \sum_i implies that the index i runs over all integers.

The n th *Eulerian polynomial* is

$$A_n(t) = \sum_{\pi \in \mathcal{S}_n} t^{\text{des}(\pi)} = \sum_d E(n, d)t^d \quad \text{for } n \geq 1,$$

and $A_0(t) = 1$, see Kyle Petersen [20, §1.4]. The exponential generating function of the Eulerian polynomials is

$$(2.1) \quad \sum_n A_n(t) \frac{x^n}{n!} = \frac{t-1}{t-e^{(t-1)x}},$$

see [20, Theorem 1.6] and [14, Formula (75)]. (It is often useful to consider the variant $E_n(t)$ of Eulerian polynomials defined by $E_n(t) = tA_n(t)$, which are also called the *Eulerian polynomials* in some literatures, see Stanley [30, §1.3] and Bóna [7, Theorem 1.22].)

The bivariate generating function

$$\begin{aligned} E(x, t) &= \sum_{n \geq 1} \sum_d \frac{E(n, d)t^d x^n}{n!} = \sum_{n \geq 1} A_n(t) \frac{x^n}{n!} \\ &= x + \frac{x^2}{2}(1+t) + \frac{x^3}{3!}(1+4t+t^2) + \frac{x^4}{4!}(1+11t+11t^2+t^3) + \dots \end{aligned}$$

has the closed form

$$(2.2) \quad E(x, t) = \frac{t-1}{t-e^{(t-1)x}} - 1 = \frac{e^{(1-t)x} - 1}{1 - te^{(1-t)x}}.$$

For $m \geq 1$, nonnegative integers n_1, n_2, \dots, n_m , and statistics $\text{st}_1, \text{st}_2, \dots, \text{st}_m$ over \mathcal{S}_n , let

$$\begin{aligned} \mathcal{P}_n^{(\text{st}_1, \text{st}_2, \dots, \text{st}_m)}(n_1, n_2, \dots, n_m) &= \{\pi \in \mathcal{S}_n : \text{st}_i(\pi) = n_i \text{ for all } 1 \leq i \leq m\} \quad \text{and} \\ p_n^{(\text{st}_1, \text{st}_2, \dots, \text{st}_m)}(n_1, n_2, \dots, n_m) &= |\mathcal{P}_n^{(\text{st}_1, \text{st}_2, \dots, \text{st}_m)}(n_1, n_2, \dots, n_m)|. \end{aligned}$$

For convenience, we define $\mathcal{P}_n^{(\text{st}_1, \text{st}_2, \dots, \text{st}_m)}(n_1, n_2, \dots, n_m) = \emptyset$ if any one of n, n_1, \dots, n_m is negative, and

$$\mathcal{P}_0^{(\text{st}_1, \text{st}_2, \dots, \text{st}_m)}(n_1, n_2, \dots, n_m) = \begin{cases} \{\epsilon\}, & \text{if } n_1 = n_2 = \dots = n_m = 0, \\ \emptyset, & \text{otherwise.} \end{cases}$$

We study the joint distribution of the statistics $\text{st}_1, \text{st}_2, \dots, \text{st}_m$ by the generating function

$$\begin{aligned} &P^{(\text{st}_1, \text{st}_2, \dots, \text{st}_m)}(x, x_1, \dots, x_m) \\ &= \sum_{n, n_1, \dots, n_m} p_n^{(\text{st}_1, \text{st}_2, \dots, \text{st}_m)}(n_1, n_2, \dots, n_m) x_1^{n_1} x_2^{n_2} \dots x_m^{n_m} \frac{x^n}{n!} \\ &= 1 + \sum_{n \geq 1} p_n^{(\text{st}_1, \text{st}_2, \dots, \text{st}_m)}(x_1, x_2, \dots, x_m) \frac{x^n}{n!}. \end{aligned}$$

Replacing the set \mathcal{S}_n by \mathcal{B}_n in the above definitions, we obtain analogous definitions of the set $\mathcal{B}_n^{(\text{st}_1, \dots, \text{st}_m)}(n_1, \dots, n_m)$, the number $b_n^{(\text{st}_1, \dots, \text{st}_m)}(n_1, \dots, n_m)$, and the generating function $B^{(\text{st}_1, \dots, \text{st}_m)}(x, x_1, \dots, x_m)$. For convenience, we denote (st_1) by st_1 without the parentheses.

For a permutation $\pi = \pi_1\pi_2\cdots\pi_n$ on distinct integers, denote by π^r the reversal of π , namely, $\pi^r = \pi_n\pi_{n-1}\cdots\pi_1$. The *standardization* of π , denoted $\text{std}(\pi)$, is the permutation $\sigma_1\sigma_2\cdots\sigma_n \in \mathcal{S}_n$ such that $\sigma_i < \sigma_j$ if and only if $\pi_i < \pi_j$. For convenience, we define $\text{std}(\epsilon) = \epsilon$.

3. STATISTICS OVER BALLOT PERMUTATIONS

In this section, we use a map, which we call the *reversal-concatenation*, to establish a series of relations between joint distributions of statistics over \mathcal{S}_n and \mathcal{B}_n . Let

$$\mathcal{P} = \{(\rho, \tau) : \exists 0 \leq l \leq n \text{ such that } \text{std}(\rho) \in \mathcal{B}_l, \text{std}(\tau) \in \mathcal{B}_{n-l}, \text{ and } \rho\tau \in \mathcal{S}_n\}.$$

The reversal-concatenation map ϕ is defined as

$$(3.1) \quad \begin{aligned} \phi: \mathcal{P} &\rightarrow \bigcup_{n \geq 1} \mathcal{S}_n \\ (\rho, \tau) &\mapsto \rho^r \tau. \end{aligned}$$

Suppose that $\rho = \rho_1\rho_2\cdots\rho_l$ and $\tau = \tau_1\tau_2\cdots\tau_{n-l}$. Let $\pi = \phi(\rho, \tau)$. It is easy to check that

$$(3.2) \quad \text{des}(\pi) = l - 1 - \text{des}(\rho) + \text{des}(\tau) + \chi(\rho = \epsilon \text{ or } \rho_1 > \tau_1),$$

where χ is the characteristic function. It is also easy to see that when $(\rho, \tau) \neq (\epsilon, \epsilon)$,

- if $\rho = \epsilon$ or $\rho_1 > \tau_1$, then the position $l + 1$ is the first lowest position of π ;
- if $\tau = \epsilon$ or $\rho_1 < \tau_1$, then the position l is the last lowest position of π .

For example, the first lowest position of the permutation $\phi(341, 265) = 143265$ is 4; the last lowest position of $\phi(134, 256) = 431256$ is 3.

A position $2 \leq i \leq n - 1$ in a permutation $\pi = \pi_1\pi_2\cdots\pi_n \in \mathcal{S}_n$ is a *peak* if $\pi_{i-1} < \pi_i$ and $\pi_i > \pi_{i+1}$. Let $\text{pk}(\pi)$ be the number of peaks in π .

3.1. Connecting the joint distribution (pk, dp, des) over \mathcal{S}_n and the joint distribution (pk, des) over \mathcal{B}_n . The goal of this subsection is to establish Theorems 3.4 and 3.5. For clarification, we show the following specification of Theorem 3.5 first.

Theorem 3.1. *For $n \geq 0, k \geq 0$ and $d \geq 0$ such that $(n, k, d) \neq (0, 0, 1)$, we have*

$$(3.3) \quad p_n^{(\text{pk}, \text{des})}(k, d) + p_n^{(\text{pk}, \text{des})}(k, d - 1) = \sum_{l, i, j} \binom{n}{l} b_l^{(\text{pk}, \text{des})}(i, j) b_{n-l}^{(\text{pk}, \text{des})}(k - i, d - l + j).$$

Proof. It is direct to check Eq. (3.3) for $n \leq d$ and for $d = 0$. Let $n > d \geq 1$ and

$$\begin{aligned} R_n(k, d) &= \{(\rho, \tau) : \exists 0 \leq l \leq n \text{ and } i, j \geq 0 \text{ such that} \\ &\quad \text{std}(\rho) \in \mathcal{B}_l^{(\text{pk}, \text{des})}(i, j), \text{std}(\tau) \in \mathcal{B}_{n-l}^{(\text{pk}, \text{des})}(k - i, d - l + j), \text{ and } \rho\tau \in \mathcal{S}_n\}. \end{aligned}$$

We shall show that both sides of Eq. (3.3) equal $|R_n(k, d)|$.

For the left side, we shall show that $\varphi = \phi|_{R_n(k,d)}: R_n(k,d) \rightarrow U_n(k,d)$ is a bijection, where

$$U_n(k,d) = \mathcal{P}_n^{(\text{pk}, \text{des})}(k,d) \cup \mathcal{P}_n^{(\text{pk}, \text{des})}(k,d-1).$$

First, we verify that $\varphi(R_n(k,d)) \subseteq U_n(k,d)$. Let $(\rho, \tau) \in R_n(k,d)$. Suppose that

$$\rho = \rho_1 \rho_2 \cdots \rho_l \quad \text{and} \quad \tau = \tau_{l+1} \tau_{l+2} \cdots \tau_n,$$

$$\text{std}(\rho) \in \mathcal{B}_l^{(\text{pk}, \text{des})}(i,j) \quad \text{and} \quad \text{std}(\tau) \in \mathcal{B}_{n-l}^{(\text{pk}, \text{des})}(k-i, d-l+j).$$

Let $\pi = \varphi(\rho, \tau)$, i.e., $\pi = \rho^r \tau$.

- If $\rho = \epsilon$ or $\rho_1 > \tau_1$, then the first lowest position of π is $l+1$. By Eq. (3.2),

$$\text{des}(\pi) = (l-1-j) + (d-l+j) + 1 = d.$$

Since every peak in π is either in ρ^r or in τ , we find

$$\text{pk}(\pi) = \text{pk}(\rho^r) + \text{pk}(\tau) = \text{pk}(\rho) + \text{pk}(\tau) = i + (k-i) = k.$$

Thus $\pi \in \mathcal{P}_n^{(\text{pk}, \text{des})}(k,d)$.

- If $\tau = \epsilon$ or $\rho_1 < \tau_1$, then the last lowest position of π is l . Similarly,

$$\text{des}(\pi) = (l-1-j) + (d-l+j) = d-1 \quad \text{and} \quad \text{pk}(\rho^r \tau) = \text{pk}(\rho) + \text{pk}(\tau) = k.$$

Thus $\pi \in \mathcal{P}_n^{(\text{pk}, \text{des})}(k,d-1)$. This completes the verification.

Second, we show that φ is injective. Suppose that $(\rho', \tau') \in R_n(k,d)$ such that $\varphi(\rho', \tau') = \pi$, i.e., $(\rho')^r \tau' = \rho^r \tau$. For $\pi \in U_n(k,d)$, we define

$$l(\pi) = \begin{cases} \max L(\pi), & \text{if } \text{des}(\pi) = d-1, \\ \min L(\pi) - 1, & \text{if } \text{des}(\pi) = d. \end{cases}$$

By the definition of φ , both the permutations ρ and ρ' have length $l(\pi)$. It follows immediately that $\rho' = \rho$ and $\tau' = \tau$. This proves the injectiveness of φ .

Thirdly, we show that φ is surjective. For any $\pi = \pi_1 \pi_2 \cdots \pi_n \in U_n(k,d)$, let

$$\rho = \pi_{l(\pi)} \pi_{l(\pi)-1} \cdots \pi_1 \quad \text{and} \quad \tau = \pi_{l(\pi)+1} \pi_{l(\pi)+2} \cdots \pi_n.$$

It is clear that $\varphi(\rho, \tau) = \pi$. Hence φ is surjective and thus bijective. Therefore,

$$|R_n(k,d)| = |U_n(k,d)| = p_n^{(\text{pk}, \text{des})}(k,d) + p_n^{(\text{pk}, \text{des})}(k,d-1).$$

Now we show the right side of Eq. (3.3) also equals $|R_n(k,d)|$. In fact, since for any pair $(\rho', \tau') \in \mathcal{B}_l^{(\text{pk}, \text{des})}(i,j) \times \mathcal{B}_{n-l}^{(\text{pk}, \text{des})}(k-i, d-l+j)$, there are $\binom{n}{l}$ pairs (ρ, τ) such that $\text{std}(\rho) = \rho'$ and $\text{std}(\tau) = \tau'$, we obtain

$$|R_n(k,d)| = \bigcup_{l,i,j} \binom{n}{l} |\mathcal{B}_l^{(\text{pk}, \text{des})}(i,j) \times \mathcal{B}_{n-l}^{(\text{pk}, \text{des})}(k-i, d-l+j)|,$$

which is simplified to the right side of Eq. (3.3). This completes the proof. \square

Remark 3.1. From the proof of Theorem 3.1, we see that similar statements hold if the statistic pk is replaced by a statistic st such that

- (1) $\text{st}(\epsilon) = 0$;
- (2) $\text{st}(\pi) = \text{st}(\pi^r)$ for any π ;
- (3) $\text{st}(\pi) = \text{st}(\pi_1\pi_2\cdots\pi_{i-1}) + \text{st}(\pi_i\pi_{i+1}\cdots\pi_n)$ for any $\pi = \pi_1\pi_2\cdots\pi_n \in \mathcal{S}_n$, where $i = \min L(\pi)$.

Theorem 3.1 is translated into the language of generating functions as follows.

Theorem 3.2.

$$B^{(\text{pk}, \text{des})}\left(xt, y, \frac{1}{t}\right)B^{(\text{pk}, \text{des})}(x, y, t) = (1+t)P^{(\text{pk}, \text{des})}(x, y, t) - t.$$

Proof. Multiplying each term in Eq. (3.3) by $y^k t^d x^n / n!$ and summing over all integers $n \geq 0$, $d \geq 0$ and $k \geq 0$ such that $(n, k, d) \neq (0, 0, 1)$, we deduce the following equations respectively:

$$\begin{aligned} \sum_{(n,k,d) \neq (0,0,1)} p_n^{(\text{pk}, \text{des})}(k, d) y^k t^d \frac{x^n}{n!} &= \mathcal{P}^{(\text{pk}, \text{des})}(x, y, t), \\ \sum_{(n,k,d) \neq (0,0,1)} p_n^{(\text{pk}, \text{des})}(k, d-1) y^k t^d \frac{x^n}{n!} &= t \sum_{(n,k,d) \neq (0,0,1)} p_n^{(\text{pk}, \text{des})}(k, d-1) y^k t^{d-1} \frac{x^n}{n!} \\ &= t \sum_{(n,k,d) \neq (0,0,0)} p_n^{(\text{pk}, \text{des})}(k, d) y^k t^d \frac{x^n}{n!} \\ &= t \mathcal{P}^{(\text{pk}, \text{des})}(x, y, t) - t, \end{aligned}$$

and

$$\begin{aligned} &\sum_{(n,k,d) \neq (0,0,1)} \sum_{l,i,j} \binom{n}{l} b_l^{(\text{pk}, \text{des})}(i, j) b_{n-l}^{(\text{pk}, \text{des})}(k-i, d-l+j) y^k t^d \frac{x^n}{n!} \\ &= \sum_{n,k,d,l,i,j} b_l^{(\text{pk}, \text{des})}(i, j) \frac{y^i (xt)^l}{t^j l!} \cdot b_{n-l}^{(\text{pk}, \text{des})}(k-i, d-l+j) y^{k-i} t^{d-l+j} \frac{x^{n-l}}{(n-l)!} \\ &= \sum_{l,i,j} b_l^{(\text{pk}, \text{des})}(i, j) \frac{y^i x^l}{t^j l!} \sum_{n,k,d} b_n^{(\text{pk}, \text{des})}(k, d) y^k t^d \frac{x^n}{n!} \\ &= B^{(\text{pk}, \text{des})}\left(xt, y, \frac{1}{t}\right) B^{(\text{pk}, \text{des})}(x, y, t). \end{aligned}$$

Combining them together, we obtain the desired equation. \square

In order to give a formula for the generating function $B^{(\text{pk}, \text{des})}(x, y, t)$, we introduce an operator D^{x_1, x_2} for multivariate formal power series

$$P = P(x_1, x_2, \dots, x_k) = \sum_{n_1, n_2, \dots, n_k} p(n_1, n_2, \dots, n_k) x_1^{n_1} x_2^{n_2} \cdots x_k^{n_k}$$

by defining

$$D^{x_1, x_2} P = \sum_{n_1 \leq (n_2 - 1)/2, n_3, \dots, n_k} p(n_1, n_2, \dots, n_k) x_1^{n_1} x_2^{n_2} \cdots x_k^{n_k}.$$

For example, $D^{t, x}(x + 3x^2yt + 2x^3y^2t) = x + 2x^3y^2t$. From the definition, it is easy to see that

$$(D^{x_1, x_2} P)|_{x_{i_1}=c_1, \dots, x_{i_j}=c_j} = D^{x_1, x_2}(P|_{x_{i_1}=c_1, \dots, x_{i_j}=c_j})$$

for any $3 \leq i_1 < \cdots < i_j \leq k$ and constants c_1, c_2, \dots, c_j .

Besides, we will need the following result of Zhuang [37, Theorem 4.2].

Theorem 3.3. For $n \geq 1$,

$$(3.4) \quad \sum_{\pi \in \mathcal{S}_n} t^{\text{des}(\pi)+1} y^{\text{pk}(\pi)+1} = \left(\frac{1+u}{1+uv} \right)^{n+1} v A_n(v),$$

where

$$(3.5) \quad \begin{cases} u = \frac{1+t^2 - 2yt - (1-t)\sqrt{(1+t)^2 - 4yt}}{2(1-y)t}, \\ v = \frac{(1+t)^2 - 2yt - (1+t)\sqrt{(1+t)^2 - 4yt}}{2yt}. \end{cases}$$

Now we can give a formula for $B^{(\text{pk}, \text{des})}(x, y, t)$.

Theorem 3.4.

$$B^{(\text{pk}, \text{des})}(x, y, t) = \exp \left(D^{t, x} \ln \left(1 + \frac{(1+t)(1+u)v(w-1)}{yt(1+uv)(1-vw)} \right) \right).$$

where u and v are defined by Eq. (3.5), and

$$w = \exp \left(\frac{x(1+u)(1-v)}{1+uv} \right).$$

Proof. By Eq. (3.4) and Eq. (2.1), we can deduce that

$$\begin{aligned} P^{(\text{pk}, \text{des})}(x, y, t) &= 1 + \sum_{n \geq 1} \sum_{\pi \in \mathcal{S}_n} y^{\text{pk}(\pi)} t^{\text{des}(\pi)} \frac{x^n}{n!} \\ &= 1 + \frac{1}{yt} \sum_{n \geq 1} \frac{(1+u)v}{1+uv} A_n(v) \frac{\left(\frac{x(1+u)}{1+uv} \right)^n}{n!} \\ &= 1 + \frac{(1+u)v}{yt(1+uv)} \left(\frac{v-1}{v - \exp\left(\frac{x(1+u)(v-1)}{1+uv}\right)} - 1 \right) \\ &= 1 + \frac{(1+u)v(w-1)}{yt(1+uv)(1-vw)}. \end{aligned}$$

By Theorem 3.2, we have

$$\begin{aligned} \ln\left(B^{(\text{pk}, \text{des})}\left(xt, \frac{1}{t}, y\right)\right) + \ln(B^{(\text{pk}, \text{des})}(x, y, t)) &= \ln((1+t)P^{(\text{pk}, \text{des})}(x, y, t) - t) \\ &= \ln\left(1 + \frac{(1+t)(1+u)v(w-1)}{yt(1+uv)(1-vw)}\right). \end{aligned}$$

Since $\text{des}(\pi) \leq (n-1)/2$ for $\pi \in \mathcal{S}_n$, the expansion of

$$\ln(B^{(\text{pk}, \text{des})}(x, y, t)) = \ln\left(1 + \sum_{n \geq 1} \sum_{\pi \in \mathcal{S}_n} y^{\text{pk}(\pi)} t^{\text{des}(\pi)} \frac{x^n}{n!}\right)$$

is a multivariate formal power series, with terms of the form $x^n y^k t^d$ such that $d \leq (n-1)/2$. Similarly, the terms of the series

$$\ln\left(B^{(\text{pk}, \text{des})}\left(xt, \frac{1}{t}, y\right)\right) = \ln\left(1 + \sum_{n \geq 1} \sum_{\pi \in \mathcal{S}_n} y^{\text{pk}(\pi)} t^{n-\text{des}(\pi)} \frac{x^n}{n!}\right)$$

are of the form $x^n y^k t^d$ such that $d > (n-1)/2$. Therefore,

$$\ln(B^{(\text{pk}, \text{des})}(x, y, t)) = \left(D^{t,x} \ln\left(1 + \frac{(1+t)(1+u)v(w-1)}{yt(1+uv)(1-vw)}\right)\right),$$

which yields the desired equation. \square

Now we give a generalization of Theorem 3.2, by considering the statistic dp over \mathcal{S}_n . It can be shown by a proof that is similar to those of Theorem 3.1 and Theorem 3.2.

Theorem 3.5. *For $n, d, h, k \geq 0$ such that $(n, k, h, d) \neq (0, 0, 1, 1)$,*

$$\begin{aligned} (3.6) \quad p_n^{(\text{pk}, \text{dp}, \text{des})}(k, h, d) + p_n^{(\text{pk}, \text{dp}, \text{des})}(k, h-1, d-1) \\ = \sum_{i, j} \binom{n}{2i+h} b_{2i+h}^{(\text{pk}, \text{des})}(j, i) b_{n-2i-h}^{(\text{pk}, \text{des})}(k-j, d-i-h). \end{aligned}$$

In other words,

$$(3.7) \quad B^{(\text{pk}, \text{des})}\left(xzt, y, \frac{1}{z^2t}\right) B^{(\text{pk}, \text{des})}(x, y, t) = (1+zt)P^{(\text{pk}, \text{dp}, \text{des})}(x, y, z, t) - zt.$$

Proof. Consider the set

$$\begin{aligned} Q(n, k, d, h) &= \{(\rho, \tau) : \rho\tau \in \mathcal{S}_n, \exists i, j \geq 0 \text{ such that} \\ &\quad \text{std}(\rho) \in \mathcal{B}_{2i+h}^{(\text{pk}, \text{des})}(j, i) \text{ and } \text{std}(\tau) \in \mathcal{B}_{n-2i-h}^{(\text{pk}, \text{des})}(k-j, d-i-h)\}. \end{aligned}$$

Similar to the proof of Theorem 3.1, it can be proved that the map $\phi|_{Q(n, k, d, h)}$ is a bijection from $Q(n, k, d, h)$ to the union

$$\mathcal{P}_n^{(\text{pk}, \text{dp}, \text{des})}(k, h, d) \cup \mathcal{P}_n^{(\text{pk}, \text{dp}, \text{des})}(k, h-1, d-1),$$

which implies Eq. (3.6). The desired generating function can be obtained by using standard techniques in generatingfunctionology as that is used in the proof of Theorem 3.2. \square

Remark 3.2. Eq. (3.7) reduces to Theorem 3.2 by specifying $z = 1$.

3.2. The bivariate generating functions for statistics pk, des over \mathcal{B}_n . First, we can deduce the bivariate generating functions for the statistic pk over \mathcal{B}_n .

Theorem 3.6.

$$B^{\text{pk}}(x, y) = \sqrt{\frac{\sqrt{1-y} \cosh(x\sqrt{1-y}) + \sinh(x\sqrt{1-y})}{\sqrt{1-y} \cosh(x\sqrt{1-y}) - \sinh(x\sqrt{1-y})}}.$$

Proof. We set $t = 1$ in Theorem 3.2. Since

$$B^{(\text{pk}, \text{des})}(x, y, 1) = B^{\text{pk}}(x, y) \quad \text{and} \quad P^{(\text{pk}, \text{des})}(x, y, 1) = P^{\text{pk}}(x, y),$$

we find

$$(3.8) \quad [B^{\text{pk}}(x, y)]^2 = 2P^{\text{pk}}(x, y) - 1.$$

It is known that

$$P^{\text{pk}}(x, y) = \frac{\sqrt{1-y} \cosh(x\sqrt{1-y})}{\sqrt{1-y} \cosh(x\sqrt{1-y}) - \sinh(x\sqrt{1-y})},$$

see Entringer [11], Kitaev [19] and Zhuang [36] for instance. Substituting the above equation into Eq. (3.8), we derive the desired equation. \square

The first few terms of $B^{\text{pk}}(x, y)$ are as follows.

$$B^{\text{pk}}(x, y) = 1 + x + \frac{x^2}{2} + \frac{x^3}{3!}(1 + 2t) + \frac{x^4}{4!}(1 + 8t) + \frac{x^5}{5!}(1 + 28t + 16t^2) + \dots$$

The coefficients triangle of the above polynomial is not found in the OEIS [28].

Second, we can deduce the bivariate generating function for the statistic des over \mathcal{B}_n .

Theorem 3.7.

$$B^{\text{des}}(x, t) = \exp\left(x + 2 \sum_{k \geq 1} \sum_{d \leq k-1} E(2k, d) t^{d+1} \frac{x^{2k+1}}{(2k+1)!}\right).$$

Proof. Taking $y = 1$ in Theorem 3.2, since

$$\begin{aligned} B^{(\text{pk}, \text{des})}\left(xt, 1, \frac{1}{t}\right) &= B^{\text{des}}\left(xt, \frac{1}{t}\right), \\ B^{(\text{pk}, \text{des})}(x, 1, t) &= B^{\text{des}}(x, t), \quad \text{and} \\ P^{(\text{pk}, \text{des})}(x, 1, t) &= P^{\text{des}}(x, t) = E(x, t) + 1, \end{aligned}$$

we have

$$B^{\text{des}}\left(xt, \frac{1}{t}\right)B^{\text{des}}(x, t) = (1+t)(E(x, t) + 1) - t = 1 + (1+t)E(x, t).$$

Similar to Theorem 3.4, we have

$$B^{\text{des}}(x, t) = \exp\left(D^{t,x} \ln(1 + (1+t)E(x, t))\right).$$

By Eq. (2.2),

$$2 \sum_{k \geq 1} \sum_d \frac{E(2k, d)t^d x^{2k}}{(2k)!} = E(x, t) + E(t, -x) = \frac{e^{(1-t)x} - 1}{1 - te^{(1-t)x}} + \frac{e^{(1-t)(-x)} - 1}{1 - te^{(1-t)(-x)}}.$$

Therefore,

$$\begin{aligned} 2 \sum_{k \geq 1} \sum_d \frac{E(2k, d)t^{d+1}x^{2k+1}}{(2k+1)!} &= 2t \int_0^x \sum_{k \geq 1} \sum_d \frac{E(2k, d)t^d u^{2k}}{(2k)!} du \\ &= t \int_0^x \left(\frac{e^{(1-t)u} - 1}{1 - te^{(1-t)u}} + \frac{e^{(1-t)(-u)} - 1}{1 - te^{(1-t)(-u)}} \right) du \\ &= \ln \frac{1 - te^{x(t-1)}}{1 - te^{(1-t)x}} - 2xt. \end{aligned}$$

It is not difficult to check that

$$\ln(1 + (1+t)E(x, t)) = x - xt + \ln \frac{1 - te^{x(t-1)}}{1 - te^{(1-t)x}}.$$

Thus

$$\begin{aligned} D^{t,x}(\ln(1 + (1+t)E(x, t))) &= D^{t,x} \left(x + xt + 2 \sum_{k \geq 1} \sum_d \frac{E(2k, d)t^{d+1}x^{2k+1}}{(2k+1)!} \right) \\ &= x + 2 \sum_{k \geq 1} \sum_{d \leq k-1} E(2k, d)t^{d+1} \frac{x^{2k+1}}{(2k+1)!}, \end{aligned}$$

which completes the proof. \square

Expanding the power series $B(x, t)$ in x , we obtain

$$\begin{aligned} B^{\text{des}}(x, t) &= 1 + x + \frac{x^2}{2} + \frac{x^3}{3!}(1 + 2t) + \frac{x^4}{4!}(1 + 8t) + \frac{x^5}{5!}(1 + 22t + 22t^2) \\ &\quad + \frac{x^6}{6!}(1 + 52t + 172t^2) + \frac{x^7}{7!}(1 + 114t + 856t^2 + 604t^3) + \dots \end{aligned}$$

The coefficients triangle of the above polynomial is [28, A321280]. Now, we can establish the bivariate generating function for the statistic dp over \mathcal{S}_n .

Corollary 3.8.

$$P^{\text{dp}}(x, z) = \frac{z}{1+z} + \frac{\sqrt{1-x^2}}{(1-x)(1+z)} \exp\left(xz + 2 \sum_{k \geq 1} \sum_{d \leq k-1} E(2k, k-1-d)z^{2d+1} \frac{x^{2k+1}}{(2k+1)!}\right).$$

Proof. Taking $y = t = 1$ in Eq. (3.7), we obtain

$$(3.9) \quad B^{\text{des}}\left(xz, \frac{1}{z^2}\right) \left(\sum_{n \geq 0} b_n \frac{x^n}{n!} \right) = (1+z)P^{\text{dp}}(x, z) - z,$$

where b_n is the number of ballot permutations of length n . Substituting x by xz in Theorem 3.7, and then replacing t by $1/z^2$, we obtain

$$(3.10) \quad \begin{aligned} B^{\text{des}}\left(xz, \frac{1}{z^2}\right) &= \exp\left(xz + 2 \sum_{k \geq 1} \sum_{d \leq k-1} E(2k, d) z^{2k-2d-1} \frac{x^{2k+1}}{(2k+1)!}\right) \\ &= \exp\left(xz + 2 \sum_{k \geq 1} \sum_{d \leq k-1} E(2k, k-1-d) z^{2d+1} \frac{x^{2k+1}}{(2k+1)!}\right). \end{aligned}$$

Substituting Eqs. (1.1) and (3.10) into Eq. (3.9), one may solve $P^{\text{dp}}(x, z)$ out as desired. \square

The first few terms of $P^{\text{dp}}(x, z)$ are as follows.

$$P^{\text{dp}}(x, z) = 1 + x + \frac{x^2}{2} + \frac{x^3}{3!}(3 + 2z + z^2) + \frac{x^4}{4!}(9 + 11z + 3z^2 + z^3) + \dots$$

The coefficient triangle of the above polynomial is not found in the OEIS [28].

4. A PROOF FOR THEOREM 1.2

For convenience, define $\mathcal{O}_0(0) = \{\epsilon\}$. For any integer pair $(n, d) \neq (0, 0)$, define

$$\mathcal{O}_n(d) = \{\pi \in \mathcal{O}_n : M(\pi) = d\}.$$

Spiro's [29, Proposition 3.2] can be restated as Proposition 4.1.

Proposition 4.1 (Spiro). *For any integer $n \geq 0$ and any integer d ,*

$$(4.1) \quad |\mathcal{O}_{n+1}(d)| = |\mathcal{O}_n(d)| + \sum_i \sum_{k \geq i} 2 \binom{n}{2k} E(2k, i-1) |\mathcal{O}_{n-2k}(d-i)|.$$

Now we are in a position to prove Theorem 1.2.

Proof. In view of Theorem 3.7, it is equivalent to show that the generating function $O(x, t) = \sum_{n,d} |\mathcal{O}_n(d)| t^d x^n / n!$ is

$$(4.2) \quad O(x, t) = \exp\left(x + 2 \sum_{k \geq 1} \sum_{d \leq k-1} E(2k, d) t^{d+1} \frac{x^{2k+1}}{(2k+1)!}\right).$$

In fact, multiplying each term in Eq. (4.1) by $t^d x^n / n!$ and summing over all integers $n \geq 1$ and all integers d , we deduce the following respectively:

$$\sum_{n \geq 1} \sum_d |\mathcal{O}_{n+1}(d)| t^d \frac{x^n}{n!} = \frac{\partial O(x, t)}{\partial x} - |\mathcal{O}_1(0)| - |\mathcal{O}_1(1)| t = \frac{\partial O(x, t)}{\partial x} - 1,$$

$$\sum_{n \geq 1} \sum_d |\mathcal{O}_n(d)| t^d \frac{x^n}{n!} = O(x, t) - 1,$$

and

$$\begin{aligned} & \sum_{n \geq 1} \sum_{d, i} \sum_{k \geq i} 2E(2k, i-1) |\mathcal{O}_{n-2k}(d-i)| t^d \frac{x^n}{(2k)!(n-2k)!} \\ &= 2t \sum_i \sum_{k \geq i} E(2k, i-1) t^{i-1} \frac{x^{2k}}{(2k)!} \sum_{n, d} |\mathcal{O}_{n-2k}(d-i)| t^{d-i} \frac{x^{n-2k}}{(n-2k)!} \\ &= 2t \sum_{k \geq 1} \sum_{i \leq k} E(2k, i-1) t^{i-1} \frac{x^{2k}}{(2k)!} \sum_{n, d} |\mathcal{O}_n(d)| t^d \frac{x^n}{n!} \\ &= 2t O(x, t) \sum_{k \geq 1} \sum_{d \leq k-1} E(2k, d) t^d \frac{x^{2k}}{(2k)!}. \end{aligned}$$

Combining them together, we obtain

$$\frac{\partial O(x, t)}{\partial x} = O(x, t) \left(1 + 2t \sum_{k \geq 1} \sum_{d \leq k-1} E(2k, d) t^d \frac{x^{2k}}{(2k)!} \right).$$

Solving this differential equation out, we obtain Eq. (4.2). \square

As a corollary, we have

Corollary 4.2. *For $n \geq 1$ and $0 \leq d \leq \lfloor \frac{n-1}{2} \rfloor$, we have*

$$b_n^{\text{des}}(d) = |\mathcal{O}_n(d)| = \sum_{m=1}^n \sum_{i=0}^m \sum_{\substack{d_1 + \dots + d_i = d-i, \\ 2k_1 + \dots + 2k_i = n-m}} \frac{2^i}{m!} E(2k_1, d_1) \cdots E(2k_i, d_i).$$

Recall that Bidkhor and Sullivant [6] proved that

$$b_{2n+1}^{\text{des}}(n) = \frac{E(2n+1, n)}{n+1},$$

which leads to the following corollary.

Corollary 4.3. *For $n \geq 1$, we have*

$$\frac{E(2n+1, n)}{n+1} = \sum_{m=0}^n \sum_{i=0}^{2m+1} \sum_{\substack{d_1 + \dots + d_i = n-i \\ k_1 + \dots + k_i = n-m}} \frac{2^i}{(2m+1)!} E(2k_1, d_1) \cdots E(2k_i, d_i).$$

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