

Shapes of Binary Trees

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This is a sequel to our treatment of various attributes of trees [1], expressed in the language of probability. Let $\{Y_t : 0 \leq t \leq 1\}$ be standard Brownian excursion. Define the L_p -norm

$$\|Y\|_p = \begin{cases} \left(\int_0^1 |Y_t|^p dt \right)^{1/p} & \text{if } 0 < p < \infty, \\ \max_{0 \leq t \leq 1} |Y_t| & \text{if } p = \infty \end{cases}$$

and a (new) seminorm

$$\langle Y \rangle_p = \begin{cases} \left(\int_0^1 \int_0^v |Y_u + Y_v - 2 \min_{u \leq t \leq v} Y_t|^p du dv \right)^{1/p} & \text{if } 0 < p < \infty, \\ \max_{0 \leq u < v \leq 1} |Y_u + Y_v - 2 \min_{u \leq t \leq v} Y_t| & \text{if } p = \infty. \end{cases}$$

We examined $\|Y\|_p$ earlier [2]; $\langle Y \rangle_p$ is a less familiar random variable but nevertheless important in the study of trees. Note that $\langle Y \rangle_p$ is not a norm since, for any constant c , $\langle c \rangle_p = 0$ even if $c \neq 0$.

Let T be an ordered (strongly) binary tree with $N = 2n+1$ vertices. The **distance** between two vertices of T is the number of edges in the shortest path connecting them. The **height** of a vertex is the number of edges in the shortest path connecting the vertex and the root.

The **Wiener index** $d_1(T)$ is the sum of all $\binom{N}{2}$ distances between pairs of distinct vertices of T , and the **diameter** $d_\infty(T)$ is the maximum such distance. If $\delta(v, w)$ denotes the distance between vertices v and w , then

$$d_\lambda(T) = \left(\frac{1}{2} \sum_{v,w} \delta(v, w)^\lambda \right)^{1/\lambda}, \quad \lambda > 0,$$

includes both the Wiener index and diameter as special cases.

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The **internal path length** $h_1(T)$ of a tree is the sum of all N heights of vertices of T , and the **height** $h_\infty(T)$ is the maximum such height. Let o denote the root of T . The generalization

$$h_\lambda(T) = \left(\sum_v \delta(v, o)^\lambda \right)^{1/\lambda}, \quad \lambda > 0,$$

includes both the internal path length and height as special cases. If we restrict attention to only those $n + 1$ vertices \hat{v}_k that are leaves (terminal nodes) of T , listed from left to right, then a sequence $\delta(\hat{v}_1, o), \delta(\hat{v}_2, o), \dots, \delta(\hat{v}_{n+1}, o)$ emerges. This is called the **contour** of T .

The **width** $w_\infty(T)$ of a tree is the maximum of $\zeta_l(T)$ over all $l \geq 0$, where $\zeta_l(T)$ is the number of vertices of height l in T . Note that

$$w_\lambda(T) = \left(\sum_{l=0}^{h_\infty(T)} \zeta_l(T)^\lambda \right)^{1/\lambda}, \quad \lambda > 0,$$

includes the trivial case $w_1(T) = N$. The sequence $\zeta_0(T), \zeta_1(T), \dots, \zeta_{h_\infty}(T)$ is known as the **profile** of T .

0.1. Uniform Combinatorial Model. In this model, the $\binom{2n}{n}/(n + 1)$ ordered binary trees are weighted with equal probability, where $N = 2n + 1$ is fixed.

Janson [3] determined the joint distribution of internal path length and Wiener index:

$$\left(\frac{h_1(T)}{2N^{3/2}}, \frac{d_1(T)}{2N^{5/2}} \right) \rightarrow (\|Y\|_1, \langle Y \rangle_1)$$

as $N \rightarrow \infty$. The marginal distribution of $h_1(T)$ was obtained earlier by Takács [4, 5, 6]; the result for $d_1(T)$ is apparently new. No explicit formula for $\mathbb{P}(\langle Y \rangle_1 \leq x)$ is known; see [2] for the corresponding result for $\|Y\|_1$. We have expected values

$$\mathbb{E}(\|Y\|_1) = \frac{1}{2}\sqrt{\frac{\pi}{2}}, \quad \mathbb{E}(\langle Y \rangle_1) = \frac{1}{4}\sqrt{\frac{\pi}{2}}$$

and correlation coefficient

$$\frac{\text{Cov}(\|Y\|_1, \langle Y \rangle_1)}{\sqrt{\text{Var}(\|Y\|_1)}\sqrt{\text{Var}(\langle Y \rangle_1)}} = \sqrt{\frac{48 - 15\pi}{50 - 15\pi}} = 0.5519206030\dots$$

As an aside, we mention that $\|Y\|_1 - \langle Y \rangle_1 \geq 0$ always. Underlying the joint moment [3]

$$\mathbb{E}(\|Y\|_1^k (\|Y\|_1 - \langle Y \rangle_1)^l) = \frac{k!l!\sqrt{\pi}}{2^{(7k+9l-4)/2}\Gamma((3k+5l-1)/2)} a_{k,l}$$

is the following interesting quadratic recursion [7, 8, 9, 10, 11, 12]:

$$a_{k,l} = 2(3k + 5l - 4)a_{k-1,l} + 2(3k + 5l - 6)(3k + 5l - 4)a_{k,l-1} + \sum_{0 < i+j < k+l} a_{i,j}a_{k-i,l-j}$$

with $a_{0,0} = -1/2$, $a_{1,0} = 1 = a_{0,1}$ and $a_{k,l} = 0$ when $k < 0$ or $l < 0$. All $a_{k,l}$ but $a_{0,0}$ are positive integers when $k \geq 0$ and $l \geq 0$. Applications include the enumeration of connected graphs with n vertices and $n + m$ edges. We have asymptotics [3, 13]

$$a_{k,0} \sim \frac{1}{2\pi} 6^k (k-1)!, \quad a_{0,l} \sim C \cdot 50^l ((l-1)!)^2,$$

where the precise identity of the constant

$$C = \frac{1}{50}(0.981038\dots) = 0.01962\dots = \frac{1}{50.9664\dots}$$

remains an unsolved problem.

Chassaing, Marckert & Yor [14] determined the joint distribution of height and width:

$$\left(\frac{h_\infty(T)}{N^{1/2}}, \frac{w_\infty(T)}{N^{1/2}} \right) \rightarrow \left(\int_0^1 \frac{dt}{Y_t}, \|Y\|_\infty \right)$$

as $N \rightarrow \infty$. The marginal distribution of height was obtained earlier by Rényi & Szekeres and Stepanov [15, 16, 17, 18, 19, 20, 21, 22, 23, 24]; earlier works on width include [25, 26, 27, 28, 29, 30]. It turns out that the marginal distributions are identical (up to a factor of 2) and that this is the first of several theta distributions [31] we will see here:

$$\mathbb{P} \left(\frac{1}{2} \int_0^1 \frac{dt}{Y_t} \leq x \right) = \mathbb{P} (\|Y\|_\infty \leq x) = \frac{\sqrt{2}\pi^{5/2}}{x^3} \sum_{k=1}^{\infty} k^2 e^{-\pi^2 k^2 / (2x^2)}.$$

The expected values thus coincide:

$$\mathbb{E} \left(\frac{1}{2} \int_0^1 \frac{dt}{Y_t} \right) = \mathbb{E} (\|Y\|_\infty) = \sqrt{\frac{\pi}{2}}.$$

Rényi & Szekeres also computed the location of the maximum of the probability density [15]:

$$\text{mode} (\|Y\|_\infty) = \frac{1}{2}(2.3151543618\dots) = \frac{1}{2} \sqrt{\frac{2}{0.3731385248\dots}}.$$

Returning to the joint distribution formula, it is clear that $h_\infty(T)$ and $w_\infty(T)$ are negatively correlated. A numerical estimate (let alone an exact expression) for the correlation coefficient evidently remains open [14, 32].

For the generalized height and diameter parameters, we have marginal distributions [3, 14, 33, 34, 35]:

$$\frac{h_\lambda(T)}{2N^{(\lambda+2)/(2\lambda)}} \rightarrow \|Y\|_\lambda, \quad \frac{d_\lambda(T)}{2N^{(\lambda+4)/(2\lambda)}} \rightarrow \langle Y \rangle_\lambda$$

as $N \rightarrow \infty$. The latter includes the special cases of Wiener index ($\lambda = 1$, as mentioned before) and diameter ($\lambda = \infty$):

$$P(\langle Y \rangle_\infty \leq x) = \frac{1024\sqrt{2}\pi^{5/2}}{3x^9} \sum_{k=1}^{\infty} k^2 \left[(3 + \pi^2 k^2) x^4 - 36\pi^2 k^2 x^2 + 64\pi^4 k^4 \right] e^{-8\pi^2 k^2/x^2},$$

which possesses expected value

$$E(\langle Y \rangle_\infty) = \frac{4}{3}\sqrt{2\pi}$$

and maximum location [33]

$$\text{mode}(\langle Y \rangle_\infty) = 3.2015131492\dots = \sqrt{\frac{8}{0.7805116813\dots}}.$$

Nothing is known for other values of λ (even $\lambda = 2$ seems to have been neglected). It would also be good to learn the value of the correlation coefficient of $d_\infty(T)$ and $h_\infty(T)$, or of $d_\infty(T)$ and $w_\infty(T)$.

Consider finally the minimum height $\eta(T)$ of a leaf; that is,

$$\eta(T) = \min_{1 \leq k \leq n+1} \delta(\hat{v}_k, o).$$

It is known that $E(\eta) \rightarrow \sum_{k=1}^{\infty} 2^{k+1-2^k} = 1.5629882961\dots$ as $N \rightarrow \infty$ [36, 37]. Can this result be related to Brownian excursion in some way? We will report more on the properties of leaves of T later.

0.2. Critical Galton-Watson Model. In this model, the size $N = 2n + 1$ is free to vary: All ordered binary trees are included but with weighting 2^{-N} . (We omit subcritical and supercritical cases for reasons of space.)

Let T be a random tree. The probability that T has precisely N vertices is clearly [38]

$$\frac{1}{n+1} \binom{2n}{n} 2^{-N} \sim \sqrt{\frac{2}{\pi}} N^{-3/2};$$

hence the expected number of vertices of T is infinite. We examine this result in another way. If

$$\nu_l = \sum_{k=0}^l \zeta_k$$

where ζ_k is the number of vertices of height k in T , then $\mathbb{E}(\nu_l) = l + 1$ and $\text{Var}(\nu_l) = (2l + 1)(l + 1)l/6$, both which $\rightarrow \infty$ as $l \rightarrow \infty$. More complicated conditional distributions are due to Pakes [39, 40]:

$$\lim_{l \rightarrow \infty} \mathbb{P} \left(\frac{\nu_l}{l^2} \leq x \mid \zeta_l > 0 \right) = \int_0^x f(t) dt,$$

$$\lim_{l \rightarrow \infty} \mathbb{P} \left(\frac{\nu_l}{l^2} \leq x \mid \zeta_m > 0 \text{ for all positive integers } m \right) = \int_0^x g(t) dt,$$

where the first density function is given by

$$f(t) = \frac{2}{\sqrt{2\pi}t^{3/2}} \sum_{k=0}^{\infty} \left(\frac{(2k+1)^2}{t} - 1 \right) \exp \left(-\frac{(2k+1)^2}{2t} \right)$$

with mean $1/3$, variance $2/45$, and Laplace transform

$$\int_0^{\infty} e^{-st} f(t) dt = \sqrt{2s} \operatorname{csch}(\sqrt{2s}).$$

The second density function is not explicitly known, but has mean $1/2$, variance $1/12$ and satisfies

$$\int_0^{\infty} e^{-st} g(t) dt = \operatorname{sech}^2 \left(\sqrt{\frac{s}{2}} \right).$$

Consequently $g(t)$ is the convolution of $\tilde{g}(t)$ with itself, where

$$\tilde{g}(t) = \frac{1}{\sqrt{2\pi}t^{3/2}} \sum_{k=0}^{\infty} (-1)^k (2k+1) \exp \left(-\frac{(2k+1)^2}{8t} \right),$$

but this appears to be as far as we can go.

Define T_l to be the subtree of T consisting of all ν_l vertices up to and including height l . We have the parameters $d_\lambda(T_l)$, $h_\lambda(T_l)$ and $w_\lambda(T_l)$ available for study, but little seems to be known. Of course, $w_1(T_l) = \nu_l$. Athreya [41], building on [42, 43, 44], proved that $\mathbb{E}(w_\infty(T_l)) \sim \ln(l)$ as $l \rightarrow \infty$, which contrasts nicely with the fact that $\mathbb{P}(\zeta_k = 0) \rightarrow 1$ as $k \rightarrow \infty$. See also [45, 46, 47, 48, 49, 50, 51]. Kesten, Ney & Spitzer [52, 53, 54] demonstrated that $\mathbb{P}(h_\infty(T_l) = j) \sim 2/j^2$ as $j \rightarrow \infty$; further references include [55, 56, 57]. Can exact distributional results be found? What about other values of λ ? Is anything known about diameter for Galton-Watson trees?

More material to come...

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